

Genetic control of methane emission, feed efficiency and metagenomics in dairy cattle

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Genetic control of methane emission, feed efficiency and metagenomics in dairy cattle

Gareth Frank Difford

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Abstract

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The dairy industry faces the challenges of increasing production, remaining economically viable whilst simultaneously minimising impacts on the environment. The cost of feed is the highest variable cost of milk production, thus, improved feed efficiency is a strong wish. However, CH₄ is a potent greenhouse gas with an energy value estimated as 2 -12% of the gross feed energy intake and thus represents a loss. There is, therefore, a need to identify the phenotypic and genetic relationships between efficiency of feed utilisation and CH₄ production to ensure optimal breeding methods of increasing profitability and limiting environmental impact of dairy production.

Feed is degraded and CH₄ is produced by rumen microbes and not by the cow. The mechanisms which influence the composition of the rumen microbial community and how they, in turn, influence the feed efficiency and CH₄ production of the host, are not well understood.

Among the possible strategies, selective breeding has the benefit over others by being cumulative and persistent over generations. Genetic improvement through selection requires that phenotypes are recorded on large numbers of animals. Moreover, phenotypes must show variation, a portion of which must be genetic, and must have economic or societal value. Understanding the genetic co-variation behind and between these measures is crucial to simultaneous breeding for a more profitable and climate friendly dairy industry. However, the measurement of CH₄ emissions, feed efficiency and the rumen microbiome under commercial conditions on a large scale is not a trivial task. The aim of this PhD project was to develop and integrate phenotyping measures for CH₄ emission, feed efficiency and the rumen microbiome and to investigate their genetic potential for selective breeding.

Firstly, in **Chapter 2**, improvements were made to the sniffer method of CH₄ breath concentration recording in dairy cattle during automatic milking. An algorithm was developed to efficiently detect and correct for variable and random drift in time series between instruments and to detect when the cow's head is out of the feed bin. Using linear mixed model methodology, repeated measures per cow were used to improve precision and control sources of inaccuracy such as sensor drift, background gas concentrations and diurnal variation, that were subsequently removed. Resultantly, highly repeatable phenotypes were obtained which

demonstrated adequate agreement for the interchangeable use of two instruments. In **Chapter 3**, the ranking of cows under commercial conditions using the sniffer method was compared with the “gold standard” respiration chambers. Individual level correlations estimated as proxies for genetic correlations revealed a high correlation between sniffer-predicted CH₄ production and CH₄ production in the RC. These findings offer a proof of concept that sniffer CH₄ phenotypes recorded over a week of lactation show substantial promise as large scale indicator traits for CH₄ production using RC.

In **Chapter 4**, genetic parameters were estimated between feed intake, milk production and CH₄ breath concentration from sniffers over the course of the first lactation in Holstein cows in Denmark and The Netherlands. Through combining data between countries, genetic residual feed intake and breath gas concentrations were found to be significantly heritable, demonstrating that genetic improvement of feed efficiency and CH₄ breath gas concentration is feasible in dairy cattle. The estimated genetic correlations from the largest dataset indicated that improved feed efficiency will also result in decreased gas emissions. Furthermore, including the breath gas concentrations in a multitrait genetic evaluation increased the accuracy of bull breeding values for gRFI, demonstrating an indirect economic value of CH₄ and CO₂ breath concentration phenotypes.

In **Chapter 5**, we estimated the relative abundance of rumen bacteria and archaea and found a portion of these to be heritable in dairy cattle. The results demonstrate that host additive genetics has an influence on the abundance of some rumen bacteria and archaea. We detected significant associations between certain bacterial genera and differences in CH₄ production of the host cow, further contributing to knowledge of the underlying biological mechanisms driving CH₄ production of the host. We further extended quantitative genetic methods to estimate rumen microbial kinships between cows in place of additive genetic relationships. This enabled the quantification of variation in host CH₄ production explained by the rumen microbial composition, expressed in the new term ‘microbiability’, as the relative proportion of host variation explained by associated microbes. Crucially the microbiability and the heritability of dairy cattle CH₄ production were largely independent. Thus, selective breeding for reduced CH₄ production can be extended by methods perturbing the rumen microbiota towards reduced CH₄ production.

In **Chapter 6** (the general discussion), the value of method comparisons for phenotype development by comparatively quantifying sources of error between cheaper alternative methods and intensive gold standard methods was discussed. The primary constraint to breeding for improved feed efficiency and CH₄ production remains the recording of feed intake on a large scale under commercial conditions

and recording of “true” CH₄ production. It was proposed that the accuracy of bull breeding values for both feed efficiency and CH₄ production can be increased through the use of sniffer phenotypes in robot milking herds, using individual level correlations but a genetic correlation between sniffer phenotypes and RC CH₄ production are still needed. The records required for estimating genetic correlations with meaningful standard errors can only be achieved through substantial financial investments, development of cheaper alternative methods of phenotype recording or international collaborations.

Further to the general discussion, a portion of host phenotypic variation in CH₄ production was found to be associated with the rumen bacterial and archaeal composition. However, research is needed to determine if microbial associations are causative and methods to direct desired changes in the rumen microbial composition are still needed to unlock the potential of this under-exploited resource. The methods developed for quantifying the microbial contribution to host phenotypic variation will be of value to inform research into complex microbial-associated phenotypes, such as diseases and digestion in dairy cattle, other livestock species and humans. This thesis therefore contributes to the understanding of the genetic variation in feed efficiency, methane emissions and rumen metagenome of dairy cows.

Resumé

Mælkeproducenterne står over for udfordringer ved at øge produktionen, hvorvidt den forbliver økonomisk levedygtig samtidig med at miljøpåvirkningerne minimeres. Kostprisen ved foder er den højeste variable omkostning til mælkeproduktion, og derfor er forbedret fodereffektivitet et stærkt ønske. Metan (CH_4) er imidlertid en potent drivhusgas med en energiværdi anslået som 2-12% af bruttoenergiindtaget og udgør således et tab. Der er derfor et behov for at identificere de fænotypiske og genetiske forhold mellem effektiviteten af foderudnyttelse og CH_4 -produktion for at sikre optimale avlsmetoder til øget lønsomhed og begrænse miljøpåvirkning fra mælkeproduktionen.

Foderet nedbrydes og CH_4 produceres af vom-mikrober men ikke af koen selv. De mekanismer, der påvirker sammensætningen af det vom-mikrobielle samfund, og hvordan de igen påvirker fodereffektiviteten og CH_4 -produktionen hos koen som vært, kendes ikke i tilstrækkeligt omfang.

Blandt de mulige strategier til at reducere klimaaftrykket har selektiv avl fordelene overfor andre ved at være kumulative og permanente over fremtidige generationer. Genetisk forbedring gennem udvælgelse kræver, at fænotyper registreres på et stort antal dyr. Desuden skal fænotyper vise variation, hvoraf en del skal være genetisk og skal have økonomisk eller samfundsmæssig værdi. At forstå den genetiske sam-variation bag og mellem disse egenskaber er afgørende for samtidig avlsarbejde mod en mere rentabel og klimavenlig mælkeproduktion. Måling af CH_4 -udledninger, fodereffektivitet og vom-mikrobiomet under kommercielle forhold i stor skala er imidlertid ikke en triviel opgave. Formålet med dette ph.d.-projekt var at udvikle og integrere fænotypebestemmelser for CH_4 -emission, fodereffektivitet og vommikrobiomet og undersøge deres genetiske potentiale for selektiv avl. Resultaterne heraf er beskrevet i 4 af afhandlingens kapitler.

For det første, i **kapitel 2**, blev der foretaget forbedringer i forbindelse med sniffermetoden for CH_4 -åndedrætsmålinger hos køer under automatisk malkning. En algoritme blev udviklet til effektivt at finde og korrigere for drift i tidsserier mellem instrumenter og at udpege, når koens hoved er ude af fodertruget. Ved hjælp af en lineær mixed-model-metode blev gentagne mål pr. ko brugt til at forbedre præcision og reducere unøjagtighed, såsom sensordrift, baggrundskoncentrationer og døgnvariationer, der efterfølgende blev fjernet. Resultatet, i form af højt gentagelige fænotyper blev opnået, hvilket viste en god mulighed for at anvende data indsamlet fra to forskellige instrumenter. I **kapitel 3** blev rangordningen af køer i private besætninger hvor CH_4 er målt med sniffermetoden sammenlignet med "den gyldne standard" respirationskamre. Korrelationer på individuelt niveau blev beregnet som

erstatning for genetiske korrelationer, og viste en stærk sammenhæng mellem sniffer CH_4 produktion og CH_4 produktion i RC. Disse resultater giver et bevis på, at sniffer CH_4 -fænotyper registreret over en uge med laktation viser et stort potentiale som storskala-indikatoregenskaber for CH_4 -produktion målt ved anvendelse af RC.

I **kapitel 4** blev genetiske parametre beregnet mellem foderoptagelse, mælkeproduktion og CH_4 -åndedræts-koncentration fra sniffere i løbet af den første laktation hos Holstein-køer i Danmark og Nederlandene. Ved at kombinere data fra begge lande blev genetisk residual foderoptagelse og åndedræts-koncentrationer fundet i væsentlig grad arvelige, hvilket viser, at genetisk forbedring af fodereffektivitet og CH_4 -koncentration er mulig hos malkekvæg. De beregnede genetiske korrelationer fra det største datasæt viste, at forbedret fodereffektivitet også vil resultere i faldende gasemissioner. Derudover øgede gaskoncentrationerne i en multiegenskabs-genetisk beregning nøjagtigheden af tyres avlsværdier for gRFI, hvilket demonstrerede en indirekte økonomisk værdi af fænotyper af CH_4 og CO_2 -målinger.

I **kapitel 5** vurderede vi den relative forekomst af vom-bakterier og -arkæer og fandt at en del af disse er arvelige hos malkekvæg. Resultaterne viser, at koen genetik, når den ses som vært har indflydelse på forekomsten af nogle vombakterier og archæer. Vi opdagede væsentlige sammenhænge mellem visse bakterielle slægter og forskelle i CH_4 -produktion fra værts-koen, hvilket yderligere bidrager til kendskab til de underliggende biologiske mekanismer, der bestemmer koens CH_4 -produktion. Vi videreudviklede de kendte kvantitative genetiske metoder til også at omfatte vommens mikrobielle slægtskab mellem køer udover de additivt genetiske forhold. Dette muliggjorde kvantificering af variation i værts- CH_4 -produktion forklaret af den vommikrobielle sammensætning udtrykt i den nye betegnelse 'mikrobiabilitet' som den relative andel af værtsvariation forklaret af tilhørende vommikrober. Mikrobiabiliteten og arveligheden af CH_4 produktion var stort set uafhængig. Selektiv avl til reduceret CH_4 -produktion kan således udvides ved fremgangsmåder, der påvirker mikrobiomet mod reduceret CH_4 produktion. I **kapitel 6** (den generelle diskussion) blev værdien af metode-sammenligninger diskuteret i forhold til fænotyper ved kvantificerede fejlkilder hos billigere alternative metoder og intensive "guldstandard" metoder. Den primære begrænsning for avl for forbedret fodereffektivitet og CH_4 produktion forbliver målinger af foderoptagelsen i stor skala under kommercielle forhold og registrering af "sand" CH_4 produktion. Det blev foreslået, at nøjagtigheden af tyre-avlsværdier for både fodereffektivitet og CH_4 -produktion kan øges ved brug af snifferfænotyper i robotmalkebesætninger ved brug af individuelle niveaukorrelationer, men der er stadig brug for en genetisk korrelation mellem snifferfænotyper og RC CH_4 -produktion. De indsatser der kræves

til vurdering af genetiske korrelationer med meningsfulde standardfejl, kan kun opnås gennem betydelige finansielle investeringer, udvikling af billigere alternative metoder til fænotypemålinger eller internationale samarbejder. I forlængelse af den generelle diskussion viste det sig at en del af den fænotypiske variation i CH₄-produktion hænger sammen med vom-bakterielle og arkæiske sammensætning. Imidlertid er der behov for yderligere forskning for at bestemme, om mikrobielle sammenhænge også er årsags-sammenhængende, og der er behov for metoder til at styre ønskede ændringer i den vom-mikrobielle sammenhæng for at udnytte potentialet for denne underudnyttede ressource. Metoderne udviklet til kvantificering af det mikrobielle bidrag til værts-fænotypisk variation vil være af værdi for videre forskning i komplekser mellem mikrobielle fænotyper, såsom sygdomme og fordøjelse i mælkekvæg, også hos andre dyrearter og mennesker. Denne afhandling bidrager derfor til forståelsen af den genetiske variation i fodereffektivitet, metanemissioner og vom-metagenom af malkekøer.

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1

General introduction

1.1 Rational for research

Dairy cattle and other livestock industries face the challenges of increasing production, remaining economically viable whilst simultaneously minimising impacts on the environment. According to a projection by the FAO, global food production needs to increase by approximately 70% to meet the needs of 9.7 billion people by 2050 (UN, 2015). Milk, being a complete source of fat, protein, lactose, vitamins and minerals, must further increase by roughly 80% by 2050 (Steinfeld et al., 2006). The natural resources to sustain expansion to meet these demands are finite as seen by total global cultivated land remaining largely constant since 1991 (O'Mara, 2012). Additionally, increases in milk production are associated with greater greenhouse gas emissions like methane and carbon dioxide, as well as greater feed intake. Accordingly, research into a more sustainable and profitable dairy industry by improving production efficiency whilst limiting environmental impact is becoming increasingly relevant to the producer and consumer. Specifically, research to mitigate the rising variable costs of feed by increasing feed efficiency and to decrease the environmental impact by reducing greenhouse gas (GHG) emissions and nutrient losses to the environment (Connor et al., 2013).

Methane (CH₄) is a potent greenhouse gas (GHG) with a climate change potential ~32 times greater than carbon dioxide (CO₂) (Holmes et al., 2013) and an atmospheric half-life of 12 years, which is substantially shorter than CO₂ (> 100 years) (IPCC, 2014). Therefore, reducing CH₄ emissions from anthropogenic-related sources has been identified as a key area for mitigating climate change with immediate effects (Gerber et al., 2013; IPCC, 2014). Livestock accounts for 14.5% of anthropogenic-related GHG emission of which enteric CH₄ emissions from ruminants directly accounts for 5.8% (3). The CH₄ produced by dairy cattle and other ruminants is a natural by-product of enteric fermentation of high fibre feedstuffs by microbial enzymatic activity (Janssen and Kirs, 2008). Enteric fermentation results in the production of volatile fatty acids (VFA) (predominantly, acetate, propionate and butyrate) which is used as an energy resource by the cow (Moss et al., 2000; Martin et al., 2010). The production of acetate and butyrate releases metabolic hydrogen which is converted to H₂ by hydrogenase-expressing bacterial species, whereas the production of propionate consumes H₂ (Moss et al., 2000). An accumulation of dissolved H₂ reduces the efficiency of carbohydrate degradation and microbial protein synthesis (McAllister and Newbold, 2008). Hydrogenase-expressing bacteria convert metabolic hydrogen from anaerobic fermentation into H₂ which is then converted to CH₄ via methanogenesis and finally leaves the rumen (McAllister and

Newbold, 2008). Furthermore, CH₄ has a caloric value approximately 2 -12% of the gross energy intake (GEI) which represents a net loss of energy to the cow (Johnson and Johnson, 1995). Thus, enteric methanogenesis is a pathway for maintaining optimal conditions for ruminal fermentation but is also a potential indicator of energy loss and consequently production inefficiency.

Feed plays a crucial role in economic and environmental performances of dairy production units, as the cost of feed comprises the largest variable cost of dairy production (~44% on average globally), but can be as high as 84% in some countries (Hemme et al., 2014). Furthermore, the composition of the feed affects not only the price but also the CH₄ emitted. The primary income of dairy production is from milk and milk prices are volatile often failing to increase with feed prices, a situation over which the producer has limited control. A promising strategy to increase profitability is to reduce costs through retaining cows, which require less feed (input) to give contemporary levels of production (output) (i.e. increase feed efficiency). The quantity of feed consumed is directly proportional to the CH₄ produced (Yan et al., 2010). Therefore, cows that require less feed to produce the same level of production as their contemporaries are expected to produce less CH₄ in absolute terms (Waghorn and Hegarty, 2011). However, the organic matter composition of the feed, particularly fibre, also affects the CH₄ emissions at the individual cow level (Hristov et al., 2013) as well as the entire production chain (Møller et al., 2014). Empirical studies on digestibility and digesta passage rate have associated decreased CH₄ yield per kg of milk with reduced diet and poor cell wall digestion (Cabezas-Garcia et al., 2017), indicating that cows which are poor at extracting energy from fibre produce less methane. A greater understanding of the relationships between feed intake, feed efficiency and CH₄ production is therefore needed to ensure concomitant increases in feed efficiency and limiting environmental impact.

The majority of enzymatic processes within the rumen such as cellulases to degrade cellulose and hemicellulose, as well as methyl-coenzyme M reductase to produce CH₄, are encoded for in the genomes of the microorganisms inhabiting the rumen and not the host. The rumen is a highly specialised organ which provides an anaerobic environment hospitable to microorganisms and promotes the acquisition of commensal microbes. The relationship between ruminant host and rumen microbe has evolved over approximately 50 million years, to the extent where host cannot survive without the microbes and some of the rumen microbes cannot survive outside of the rumen. This partnership has been highly successful, leading to ruminants adapting and thriving in some of the harshest and most widely distributed environments, on high fibre diets which are mostly indigestible to many other domains of life. Increasingly, results are showing that the host may directly or

indirectly exert some influence over the rumen microbial composition (Weimer et al., 2010; Roehe et al., 2016). Differences in rumen microbial communities have been associated with differences in CH₄ emissions in sheep (Kamke et al., 2016) and feed efficiency in cattle (Hernandez-Sanabria et al., 2010; Hernandez-Sanabria et al., 2012; Carberry et al., 2012; Jewell et al., 2015). However, how the host and rumen microbiota interact to influence host phenotypes is not currently known, thus the link between causation and association is still lacking.

Although dairy cattle production has shown substantial increases in efficiency in the preceding decades, through improvements to individual milk yield as a result of advances in nutrition, genetics, and animal health, this was achieved in the absence of direct information concerning feed efficiency, CH₄ emission or the rumen microbiome (Capper et al., 2013; Hill et al., 2015). Thus, by considering this information it may be possible to further improve efficiency while reducing the environmental impact of dairy cattle production. Genetic improvement has the benefit of cumulative gains over generations, which are persistent as opposed to short-term transient nutritional, or managerial strategies. However, genetic improvement requires that phenotypes be recorded on a large number of animals, preferably under the commercial conditions in which they are expected to perform. Furthermore, the traits of interest must be variable, a portion of the variability be heritable and have an economic or societal value. The studies above suggest that methane emissions and feed efficiency are linked through feed intake and the influence of the host and its rumen microbiome. Should these phenotypes be recorded on large numbers of related animals under commercial conditions, genetic investigations can determine the potential for breeding for these phenotypes.

1.2 Biological considerations of methane emissions in dairy cattle

An understanding of the underlying biology of CH₄ emissions in dairy cattle precedes that of recording systems and genetic evaluations. The release of CH₄ from the animal is primarily through three routes: first, through eructation directly from the rumen; second, CH₄ is absorbed from the rumen and hindgut into the blood and exhaled through the lungs and third, CH₄ is emitted from the hindgut as flatulence. In cattle and in sheep, it has been estimated using radiolabeled CH₄, that approximately 98% of CH₄ is expired through the breath and eructation and only 2% from flatulence (Murray et al., 1976). This has implications for the methods of measuring CH₄ emitted by a cow as, although 98% is expired in the breath and

eructation, recording only part of the emission (expiration vs flatulence) is not necessarily the same as measuring the entire emission (expiration and flatulence) (Muñoz et al., 2012). This is particularly true if there is animal variation in the proportions emitted through the three different routes, which could be impacted by differences in digestibility.

The rate at which CH₄ is emitted changes throughout the day and from day to day and as such, the cow is in a continually changing biological state in terms of CH₄ emission. The within day diurnal variation is affected by the feeding behaviour, diet and feeding allowance and pattern (Crompton et al., 2010; Bell et al., 2018). The simplest way to avoid diurnal variations in principle is to record CH₄ emissions for every second of the entire 24 hr period to get the true daily CH₄ emission. In practice few methods sample or record continually throughout a 24 hr period and thus rely on an average estimate which is subject to experimental error. The number of measurements and timing of sampling required to obtain a representative sample of the daily CH₄ emission will vary depending on multiple factors such as feeding time, feeding behaviour, feed intake and activity (eating, ruminating etc) (Hegarty, 2013). One approach is to sample throughout the day, over multiple consecutive days and take an average estimate or moving average (Arthur et al., 2017) or model the time of day (Lassen et al., 2012; van Engelen et al., 2018) and in this manner the effects of diurnal variations are reduced or removed.

The rate of CH₄ emission also changes from day to day (Grainger et al., 2007), according to physiological state (growing, lactating and non-lactating) (Ricci et al., 2013), within lactation (early, mid, peak and late) (Rischewski et al., 2017), and from lactation to lactation. It is important to understand the phenotypic and genetic relationships between methane emissions recorded at different points in time during an animal's life to better understand the implications for selection based on CH₄ emission recorded at a particular point in time, in order to optimise recording strategies.

1.3 Methods: Evolution of individual methane recording methods

Numerous methods of recording CH₄ emission on individual dairy cattle *in vivo* are available, each with its own set of advantages, disadvantages and scope of application (Hammond et al., 2016a). The gold standard method is indirect calorimetry in respiration chambers which has been in use for more than 100 years (Krogh, 1916). As the gold standard for CH₄ emission and other respiratory gases,

respiration chambers (RC) are regarded as the most accurate and precise method from which other methods are benchmarked. As will be discussed in detail below, RC are costly, time consuming, labour intensive and not necessarily representative of all environmental conditions. Not surprisingly, numerous technologies are under development which may be cheaper, less invasive, easier to implement, or have a wider scope of application than the gold standard method. This is evidenced by frequent reviews (e.g. Patra, 2012; Storm et al., 2012; Hammond et al., 2016; Hill et al., 2016). However, the relative accuracy, precision and linear association between alternative methods and the gold standard needs to be assessed to determine the relative value and equivalence or lack thereof, between the gold standard and the alternative method (Barnhart et al., 2007).

All methods measure the CH₄ concentration and other gases often using infrared (IR) spectroscopy. The gold standard then converts this to CH₄ emission rate (or flux) in litres or grams per day by recording active air flow during measurement (Gardiner et al., 2015). Three main emission phenotypes frequently used in dairy cattle are: first, methane production as CH₄ emission rate (or flux) in L/d or g/d; second, methane yield per unit of feed intake (e.g. CH₄ production per kilogram dry matter intake (DMI)) and third, methane intensity per unit product (e.g. CH₄ production per kilogram energy corrected milk yield (ECM)). It is clear that the latter two phenotypes are the quotient of CH₄ production and other phenotypes recorded independently of the method recording CH₄. Unless otherwise specified, CH₄ emission within this thesis will refer to the first phenotype, the direct measurement of methane expressed as CH₄ concentration or CH₄ production. Different methods of recording CH₄ emissions have been comprehensively reviewed by Hammond et al., (2016) to which the reader is directed for further details. However, a brief summary of available methods and considerations directly pertaining to genetic evaluations are necessary and further discussed below.

1.3.1 Respiration chambers as the ‘gold standard’

Respiration chambers using indirect calorimetry are extremely diverse in cost, operation, design and complexity (Global Research Alliance, 2012). The basic principle is that animals are confined to airtight or near airtight chambers for a period of 24 hours or more. Air is actively circulated through the chamber, the airflow rate along with the concentration of gases going into and leaving the chamber are recorded. Methane production is determined as the product of airflow rate and the difference in methane concentrations between inflowing and outflowing air. Further corrections are made for the volume of gases using the known gas densities at

standard temperature and pressure as well as humidity. Some respiration chambers make further corrections based on recovery tests.

$$Q_{c(i)} = ([CH_4]_{M(i)} - [CH_4]_{B(i)}) \times C_{p(i)} \times Q_{air} \times 10^6 \quad (1)$$

Where $Q_c(i)$ is the CH₄ emission in Litres per time point (i), the concentration of CH₄ measured is denoted by subscript M , the concentration of background CH₄ is denoted by B , $C_p(i)$ is the fractional capture rate at time i and Q_{air} is the volumetric air flow rate in litre per time point (i). Using the temperature and atmospheric pressure recorded at time point (i) it is possible to convert CH₄ from volume to mass using the ideal gas law. Summing $Q_c(i)$ over a 24 hour period results in CH₄ production in litres per day or grams per day is the ideal gas law conversion is made.

Although, respiration chambers are regarded as the gold standard and thus, by definition are the most accurate and precise measure of gas production, they have two main sources of error which can affect accuracy and precision. These are the airflow rate (Q_{air}) or ducting efficiency (C_p) through the chamber and the mixing of gases within the chamber, which are jointly reflected in the response time (Hammond et al., 2016a). In a joint calibration procedure of respiration chambers in the UK, high within- and between-chamber and facility variation was observed, with airflow rate and chamber mixing at 15.3% and 3.4%, respectively (Gardiner et al., 2015). If the absolute accuracy of CH₄ release rate of the test gas is known with certainty and is constant over time, the recovery rate can be used as a correction factor to adjust measurements. In the UK study, after correction for differential recovery rates was made, the combined uncertainty between chambers and facilities was reduced to 2.1 % (Gardiner et al., 2015). The use of correction factors is discouraged and good practice is to identify the source of error and correct it (McLean and Tobin, 1987). With some authors calling for the publishing of recovery tests pre- and post-experiment as a prerequisite for publication (Gerrits et al., 2018).

A further factor to consider is that confinement within a chamber can stress animals and alter their feeding behaviour resulting in a drop in DMI, the largest driver of CH₄ emissions; this has led many to question the extrapolation of these results to commercial conditions particularly grazing systems (Pinares-Patiño et al., 2013). Some developments in respiration chamber methods have led to animal friendly chambers constructed from cheaper transparent materials, lowering the cost and invasiveness of the method with minimal disruption to the accuracy and precision of the measurement and showing no drop in the DMI of cows under confinement (Hellwing et al., 2012).

The throughput and capital cost of respiration chambers is the biggest challenge towards usage in genetic evaluations. Assuming a single day of acclimation and 2 consecutive days of recording, a single chamber can record CH₄ production on 120 cows a year. In practice, this is likely to be far less (30-50) a year, as was reported in the only large scale genetic evaluation of CH₄ emission in 1042 growing angus steers and heifers (Donoghue et al., 2016a). This cohort of cattle demonstrates that CH₄ production is repeatable ($t = 0.97$) over consecutive days (Donoghue et al., 2016b), heritable ($h^2 = 0.27 \pm 0.07$) (Donoghue et al., 2016a) and achieves moderate genomic prediction accuracies of 0.32 ± 0.04 (Hayes et al., 2016). However, no such large cohort studies have been reported for dairy cows in RC.

1.3.2 The Sulphur hexafluoride tracer gas technique

The Sulphur hexafluoride (SF₆) technique has been in use for two decades and was originally developed to record CH₄ emission of animals whilst grazing in paddocks or on pasture (Johnson et al., 1994). The principle behind this method is that CH₄ production can be measured if SF₆ gas emission rate from the rumen is known and employed as a tracer gas. Small permeation tubes containing SF₆ with a predetermined release rate are placed into the rumen of test animals. The SF₆ gas is mixed with the rumen-produced gases and exhaled and eructed through the mouth and nostrils. Test animals are fitted with gas sampling apparatus consisting of a halter, to support capillary tubing inlets in close proximity to the mouth and nostrils, and connected to an evacuated canister under vacuum. Exhaled air is sampled continuously through capillary tubing with in-line flow restrictors for 24 hours or usually longer. The CH₄ and SF₆ contained within the canister is determined as a batch sampled over single or multiple days using gas chromatography. Daily CH₄ emission is calculated using the ratio of background corrected CH₄:SF₆ the previously determined rate of release from the permeation tubes to give an estimate of CH₄ production (Williams et al., 2011)(eq 2):

$$RCH_4 = RSF_6 \times \frac{[CH_4]_M - [CH_4]_B}{[SF_6]_M - [SF_6]_B} \times \frac{MW_{CH_4}}{MW_{SF_6}} \times 1000 \quad (2)$$

Where RCH₄ is the release rate of CH₄ in grams per day, RSF₆ is the predetermined release rate of SF₆ tracer gas from the permeation tubes (mg/day). The measured concentrations of CH₄ and SF₆ are given with subscript *M*, the recorded background gas concentrations are given by subscript *B* and the molecular weights are denoted by *MW* and multiplied by the constant 1000 to obtain units in grams per day.

The SF₆ method has been compared to RC in a number of studies and has been found to differ in accuracy in the range of (3 - 8%) which can, in part, be due to true method differences and also CH₄ emitted as flatulence which is not captured by the SF₆ method (Grainger et al., 2007; Muñoz et al., 2012). Only a single study has compared the precision of the two methods and found the coefficient of variation (CV) within animal to be marginally higher with the SF₆ (CV = 6.1%) versus RC (CV = 4.3%), when this is expressed as repeatability (t), the SF₆ was slightly less repeatable (t = 0.27) than the RC (t = 0.30) (Muñoz et al., 2012). The primary sources of inaccuracy and imprecision are variations in the release rate of SF₆ from permeation tubes which decays over time. One method of correction is to model the rate of release using Michaelis-Menten kinetics (Moate et al., 2015). Another source of error is the rate of sampling from capillary tubes as flow restrictors. Deighton et al. (2014) detected that 16% inaccuracy was due to variation in sampling rate. This led to the development of orifice plate flow controllers to reduce error in sampling flow rates (Deighton et al., 2014). The SF₆ method is also reliant on accurate estimation of background gas concentrations which need to be relatively low in comparison to recording concentrations (Berndt et al., 2014). Williams et al. (2011) estimated that the method of background estimation can influence the accuracy of the SF₆ methods by approximately 6% and recommends the use of 'sentinel' canisters for sampling purely background gas concentrations. Further recommendations are made for outdoor, indoor and mixed use of the SF₆ methods (Berndt et al., 2014).

Large-scale genetic evaluation of CH₄ emissions in dairy cattle is feasible using the SF₆ method and is well suited to testing animals under grazing conditions. Limitations to the throughput of the method are the numbers of apparatus available and the labour and capacity of laboratory facilities recording captured gases in the canisters. The minimum recommended recording time is 5-7 consecutive days in groups of animals 15 or larger (Berndt et al., 2014) which would culminate in approximately 750 cattle per year. Studies have shown the SF₆ method to be highly repeatable over 5-10 days (t = 0.86) (Arbre et al., 2016). To date, only a single study has reported heritability estimates for CH₄ production with SF₆ ($h^2 = 0.33 \pm 0.15$) which realised phenotyping on approximately 100 cows per year (Breider et al., 2018).

1.3.3 The GreenFeed system

The GreenFeed® (C-Lock Inc., Rapid City, South Dakota, USA) is a patented standalone automatic head chamber system where a concentrate feeder is fitted with a ventilation hood. During visits to the unit to consume concentrate, the exhaled and eructed gases are spot sampled by recording CH₄ concentration, head

position and airflow rate (Zimmerman, 2011; Huhtanen et al., 2015). A radio frequency identification system (RFID) allows for the investigator to control the frequency and timing of visits to the unit - for a detailed description see Hristov et al. (2015b). Measurements are short (3 – 7 min per visit) and made multiple times a day over a period of multiple, consecutive days. Visitation to the unit is voluntary and requires training and adaption to the unit. Daily CH₄ emissions (g/d) are estimated using volumetric airflow rate adjusted to standard temperature and pressure in a manner analogous to respiration chambers (Huhtanen et al., 2015; Jonker et al., 2016):

$$Q_{c(i)} = \sum_i^n V_i \times \left(([CH_4]_{M(i)} - [CH_4]_{B(i)}) \times \frac{1}{C_{p(i)}} \times Q_{air} \times 10^6 \right) / \sum_i^n V_i \quad (3)$$

Where $Q_{c(i)}$ is the CH₄ emission in litres per day, the concentration of CH₄ measured is denoted by subscript M , the concentration of background CH₄ is denoted by B , V_i is the visit time, $C_{p(i)}$ is the fractional capture rate at time i and Q_{air} is the volumetric air flow rate in L/(i). The weighted average is found by summing the product in parenthesis and visit time (V_i) and dividing by the summation of V_i for all visits of each cow (n). Using the temperature and atmospheric pressure recorded at time point (i) or over the period ($i - i_n$), it is possible to convert CH₄ from volume to mass using the ideal gas law.

The GreenFeed system has been compared to RC and has shown mixed results for agreement, with poor concordance (CCC = 0.1) in dairy cattle (Hammond et al., 2015) but high concordance in beef cattle (CCC = 0.84) (Velazco et al., 2015). In most studies, the accuracy was not significantly different from that of the RC, but the GreenFeed had higher variation and was less precise (Velazco et al., 2015; Jonker et al., 2016; Rischewski et al., 2017). Typical sources of imprecision for the GreenFeed include the duration and number of visits. In a study optimizing recording procedures using the GreenFeed, it was found that setting a minimum visit time threshold of 3 minutes and 30 visits per animal showed a significant 72% reduction in variance as compared to visits 2 minutes in length (Arthur et al., 2017). Since the diurnal variations in CH₄ emission vary throughout the day and animals visit the GreenFeed voluntarily, the timing and number of visits affect the accuracy and precision of the method. Some studies report as few as 1.3 – 2.6 visits per animal per day and failed to detect diurnal variation or treatment effects with the GF which were detected with other methods (Hammond et al., 2015; Arthur et al., 2017). Conversely, other studies report 3.4 – 4.4 visits per animal per day and detected

significant treatment and diurnal variations (Jonker et al., 2016; Waghorn et al., 2016). The number and duration of visits need to be considered when estimating daily CH₄ production in order to account for diurnal variations. Some studies have used weighted average as in equation 3 above (Jonker et al., 2016), whilst others have used a moving average (Arbre et al., 2016; Arthur et al., 2017) or modelled repeated measures using statistical methods (Arthur et al., 2017). Thus, careful control is required of the operator to ensure adequate visitation to the unit and timing of the visits to ensure adequate statistical power of the method.

The capture rate, Cp, is another potential source of error and presumed to be 1 in equation 3 above. However, the influence of wind speed is shown to decrease the capture efficiency to 85% of static conditions and increase the variability three fold (Huhtanen et al., 2015). Hence, recommendations have been made for corrections to the capture efficiency for outdoor and potentially indoor use of the GreenFeed system (Huhtanen et al., 2015). One of the primary drawbacks of the GreenFeed system is that animals need to be trained to visit the unit. Whilst animals that have been trained will readily use the unit in subsequent experiments (Velazco et al., 2017), some animals will never use the unit (Hammond et al., 2016). This animal-instrument interaction has implications for the design and balance of treatments and the screening of animals in genetic evaluations.

Manufacturer specifications recommend 15 to 25 animals per GreenFeed unit and 30 or more visits are recommended per animal, which can be achieved in 7-20 days depending on the frequency of visits (Arthur et al., 2017). If all animals visit the unit adequately, throughput per unit is likely to be 750 to 1,300 animals per year. Repeatability estimates for CH₄ production using the GF system are high ($t = 0.77-0.78$) over 7-20 days (Arbre et al., 2016). Despite adequate numbers for genetic analysis in beef cattle ($n=616$) (Arthur et al., 2017) and dairy cattle ($n=713$) (Hristov et al., 2018), the only heritability estimate reported is from a combined data analysis using GF and sniffers with the eructation method on 769 Holstein cows ($h^2 = 0.26 \pm 0.07$) (Biscarini, 2016).

1.3.4 Breath sampling methods: ‘sniffers’

‘Sniffers’ or the sniffer method was originally developed for sampling hydrocarbon gases such as CH₄ from marine liquid samples and using nondispersive infrared (NDIR) analysers to detect increasing concentrations near marine petroliferous and gas seeps (Dunlap et al., 1960). The term was somewhat erroneously applied to the recording of CH₄ concentration using NDIR in the breath of dairy cattle, to distinguish the method from those discussed above (Pinares-Patiño et al., 2013) but is now commonly used and understood. The principle behind sniffers

is to sample the breath of individual dairy cows either during automated milking or visits to concentrate feeders when the cow's head is placed in the feed bin in close proximity to a gas sensor inlet (Madsen et al., 2010; Garnsworthy et al., 2012; Negussie et al., 2016). The appeal of this method is that it can record CH₄ concentrations (and other gases) on many animals, without disrupting animal behavior or farm management practices. The method is in formative stages of development and as yet, has not standardized protocols or best management practices, with a wide variety of experimental setups and sensors in use (Garnsworthy et al., 2012; Lassen et al., 2012; Negussie et al., 2016; Pszczola et al., 2017). Two methods are predominantly used for the conversion or estimation of CH₄ production using sniffer methods, which differ in their sources of error and will be discussed below. The first method is the eructation peaks method described in Garnsworthy et al. (2012) and the second is the CO₂ tracer gas method (Madsen et al., 2010). Both methods are short-term sampling methods and are thus subject to potential sources of error, such as diurnal variations discussed in the GF system above and method specific errors, which will be discussed below.

The Eructation Peaks method: This method follows a similar approach to RC and GF but with very low passive air sampling rates, namely 1 L/min described in equation 4 below:

$$Q_c = \left(P_f \times [CH_4]_{BC} \times \frac{1}{\delta} \times Q_{ex} \right) \times 57.2 + 252 \quad (4)$$

Where Q_c is the CH₄ production in grams per day and P_f is the frequency of eructation events. $[CH_4]_{BC}$ is the mean background corrected CH₄ concentration, calculated by taking the integral of CH₄ concentration during a peak minus the CH₄ concentration at the start of a peak. The symbol δ is the mean recovery efficiency and Q_{ex} is the efficiency of breath sampling which is assumed to be 1. The term within parenthesis is referred to as the methane emission rate, MER_m in Garnsworthy et al. (2012). The entire term in parenthesis is multiplied by a regression coefficient 57.2 and adjusted for the intercept 252 from an orthogonal regression calibration function described in Garnsworthy et al. (2012). This method was found to correlate well ($r = 0.89$) with subsequent measurements in the RC for 12 Holstein-Friesian cows (Garnsworthy et al., 2012). The concordance between the sniffer and RC using the same 12 cows used to generate the calibration equation was high (CCC = 0.88) (Garnsworthy et al 2016, personal communication). Although these results suggest good agreement between the two methods, in practice, the variability of sniffers using the eructation peaks method is far larger (CV = 27 – 67%) (Garnsworthy et al.,

2012; Bell et al., 2014) than those reported for RC (CV = 3 -34%) (Grainger et al., 2007; Yan et al., 2010).

Potential sources of error which can increase variation (imprecision) include variable wind conditions and sampling distance due to cow head movement, which will influence the recovery efficiency (δ) (Huhtanen et al., 2015; Wu et al., 2018). Systematic errors which could affect accuracy include variation in cow exhalation rate, which will influence the efficiency of breath sampling (Q_{ex}) (Wu et al., 2018) and inter-barn or inter-sampling point variation in air velocity and ventilation, which systematically changes accuracy between locations due to changes in recovery efficiency (δ). Furthermore, the calibration function generated from 12 cows explained 89% of the variation in CH₄ production between cows in the RC, with the certainty around the intercept 252 ± 21.4 S.E. and the regression coefficient 57.2 ± 9.5 which can affect both accuracy and precision of the method (Garnsworthy et al., 2012).

Sniffers using CO₂ tracer gas method: The approximation of CH₄ production using breath concentrations recorded by sniffers follows a similar approach to that of the SF₆ method above. The ratio of CH₄ to CO₂ breath concentrations is recorded during milking and the predicted CO₂ production, based on other animal characteristics, is used as a tracer gas. The equation is shown in 5 below:

$$Q_c = PCO_2 \times \frac{[CH_4]_M - [CH_4]_B}{[CO_2]_M - [CO_2]_B} \quad (5)$$

Where Q_c is the estimated CH₄ production L/day and PCO_2 is the predicted CO₂ production using a prediction equation (Madsen et al., 2010). The measured concentrations of CH₄ and CO₂ are given with subscript M and the recorded background gas concentrations are given by subscript B . Using the temperature and atmospheric pressure recorded at measurement, it is possible to convert CH₄ from volume to mass using the ideal gas law.

The CO₂ ratio method using sniffers installed in concentrate feeders together with a prediction equation based on metabolisable energy intake, is shown to be highly correlated ($r = 0.80$) and concordant (CCC = 0.70) with RC in 22 Finnish Ayrshire cows (Negussie et al., 2016). Potential sources of error are the same for this method as for sniffers using the eructation peak method and short-term spot sampling methods. A method specific source of error is the predicted CO₂ used as a tracer gas, any cow specific bias, inaccuracy or imprecision in this prediction equation will affect the final

estimate for CH₄ production. In a study using 157 records of Danish Holstein cows from multiple experiments and diets it was reported that the predicted CO₂ using the ECM, BW and days carried calf (DCC) was largely underestimating the relationships between CO₂ production and ECM (Hellwing et al., 2013). As a result, the final estimation of CH₄ production showed a high correlation ($r = 0.74$) but a significant systematic drop in accuracy of 147 L CH₄/day. A principle shortcoming of the current prediction equations for CO₂ production as a tracer gas is the unaccounted effects of mobilisation of fat particularly in early lactation (Madsen et al., 2010). There is a risk that a cow mobilising a substantial amount of body fat will have a low measured CH₄ to CO₂ ratio but an over predicted CO₂ production and the resultant effects on the estimated CH₄ production are not known.

The greatest advantages of sniffer methods is the high throughput, portable and non-invasive nature of the sampling, which is limited only by the number of animals visiting the automatic milking station or concentrate feeder. This allows for: 1) recording of CH₄ under commercial farm conditions without altering natural behaviour. 2) Provides the possibility for continuous recording over longer periods of time, resulting in full within- and over lactation recording. 3) The cost effectiveness of some sniffers offers the possibility for rapid upscaling by permanently installing sensors in robotic milking herds and concentrate feeders. A single sniffer gas analyser can phenotype 2000 – 3000 cows per year. Lassen and Løvendahl (2016) recorded CH₄ emission on 3121 cows over two summer periods in milking robot herds in Denmark. Methane production using the eructation peak method in AMS is repeatable over 7-10 days ($t = 0.89$) (Bell et al., 2014) and the CO₂ ratio method in AMS milking visit ($t = 0.34$) (Lassen et al., 2012) and the CO₂ ratio method using concentrate feeders ($t = 0.40-0.46$) over lactation (Negussie et al., 2016). Methane production using the CO₂ ratio technique in AMS is shown to be heritable for measurements over 7 days ($h^2 = 0.21 - 0.25$) (Lassen and Løvendahl, 2016; Lassen et al., 2016; Zetouni et al., 2018) and over the entire lactation ($h^2 = 0.27 \pm 0.09$) (Pszczola et al., 2017). Furthermore, the direct measurement of CH₄ concentration in AMS per milking is shown to be heritable ($h^2 = 0.11 \pm 0.02$) (van Engelen et al., 2018). The heritability of CH₄ production using the eructation peak method and the GF combined was reported as ($h^2 = 0.26 \pm 0.07$) (Biscarini, 2016).

It is clear that all methods have inherent problems, but some methods like SF₆ and sniffers are more amenable to genetic evaluations in certain production systems as is evidence by heritability estimates in literature. Depending on the intended use of the method, a balance can be reached between accuracy, precision and correlation of a method and the number and duration of visits per cow and the number of cows.

1.4 Efficiency of feed and energy utilization in dairy cattle

In contrast to CH₄ emissions, research into genetic variation in feed efficiency of dairy cattle has a long history of research (Mason et al., 1957; Veerkamp et al., 1995) but implementation in current breeding indices is limited (Pryce et al., 2015). This was preceded by research using direct calorimetry to determine energy contents of feed, faeces and urine (Kellner, 1909), as well as indirect calorimetry (i.e. respiration chambers) which identified the relationships between heat production and gaseous exchange (Brouwer, 1965). This allowed for the quantification of the difference between gross energy (GE) (energy indigested in feed) and faecal energy into digestible energy (DE). Furthermore, the recording of gas emission (including CH₄ emission) and urine energy allowed for further partitioning of DE into metabolisable energy (ME), which further partitions into net energy for maintenance, production and heat increment of feeding. This enables the determination of net energy of feeds and the devising of energy requirements of animals upon which many of the net energy systems are based (Van Es, 1978; Volden, 2011).

A key finding is that the digestibility of feeds (ratio of ME to DE) is, to a very small extent, influenced by between cow differences but, to a large extent, the composition of feeds (Cabezas-Garcia et al., 2017). However, variation in the partitioning of ME into different biological functions (energy sinks) such as milk production, growth, accretion of body fat reserves, maintenance of body tissue, activity and reproduction requirements have been observed between cows and diets (Bauman et al., 1985; Veerkamp and Emmans, 1995). Detecting a genetic component to these partial energy efficiencies is limited due to low throughput of calorimetric methods (Korver, 1988; Veerkamp and Emmans, 1995). Directing changes in the partial energy efficiencies towards promotion of energy sinks of economic value (i.e. milk production) relative to other energy sinks is the very definition of feed efficiency. Nonetheless, great improvements in energy partitioning to milk production have been realized in dairy cattle through selective breeding for milk yield without knowledge of genetic variation in partial efficiencies or feed intake phenotypes (Capper et al., 2009). This is due to the 'dilution of maintenance' where high producing dairy cows partition greater portions of ME to milk production than maintenance relative to low yielding cows of comparable size (Bauman et al., 1985; Connor et al., 2012). As feed intake and milk yield are highly genetically correlated, a biological limitation to improving feed efficiency through selection is expected, as digestibility of feed decreases in cows that consume higher levels of feed (in

multiples of maintenance requirements) (Vandehaar et al., 2016). Resultantly, the effectiveness of improving feed efficiency by selecting for increased milk yield only will be diminished. Therefore, further genetic improvement of feed efficiency requires direct measurement of feed intake on larger numbers of related individuals under commercial conditions in order to identify cows with high milk yield relative to feed intake and other energy sinks.

1.4.1 Recording of individual feed intake

In practice, feed efficiency is not a directly observable metric; rather, it is feed intake relative or conditional on production traits (Jensen et al., 1992). Not surprisingly, with such a broad definition, there is a wide variety of phenotype definitions (Berry and Pryce, 2014) but common to all of which is feed intake records. Likewise, feed intake phenotypes can vary greatly in dairy cattle. The simplest measure is the mass of feed intake (kg/day) converted to DMI (kg/day) to remove variation in moisture content of feeds. As the energy contents of feed varies greatly, this can be converted to GE (MJ/day) and to DE (MJ/day) based on predicted digestibility of feedstuffs but the use of these phenotypes are more limited in genetic evaluations (Van Arendonk et al., 1991).

Recording of DMI was initially conducted by offering individual cattle in tie stalls feed at a 5-10% refusal rate and manually weighing refusals periodically (Nielsen et al., 2003; Zom et al., 2012). This approach is time- and labour-consuming, with limited phenotyping throughput, as well as altered animal behaviour by restricting mobility. The gold standard method for recording individual cow DMI is with RFID access gates and automatic feed bins which records the mass of feed consumed at each visit which is summed for a 24 hour period (Tolkamp and Kyriazakis, 1997). This method facilitates recording of group housed animals without restrictions to mobility and limited changes to behaviour. Validation experiments have found deviations in specificity (cow is not present in feed bin but is detected) (87 – 99.8%) and sensitivity (cow is present in feed bin and detected) (99.2 – 99.9%) of the RFID transponders (DeVries et al., 2003; Bach et al., 2004; Chizzotti et al., 2015). Furthermore, differences in accuracies between automated bins and manual weighing ranging from 120 – 600 grams a day have been reported (DeVries et al., 2003; Chapinal et al., 2007; Chizzotti et al., 2015). Additionally, negative values were recorded after a cow has visited the feed bin and not consumed any feed as well as cows observed pushing feed out of the feed bin (Chapinal et al., 2007) and stealing behaviour between cows (Bossen and Weisbjerg, 2009). None of the validation studies reported correlations or concordance between methods but on the basis of sources of error listed above, deviations from 1 in the ranking of animals are

expected. Regardless, automated feed bins have increased throughput of DMI recording in research and specialised herds to the point of facilitating multinational genetic investigations into DMI (Berry et al., 2014).

Despite these sources of error, records of DMI on individual cows are highly repeatable within lactation ($t = 0.66$ with a range 0.46 – 0.84) (Sondergaard et al., 2002; Berry et al., 2014; Tempelman et al., 2015) and over lactation (Veerkamp et al., 2013). Furthermore, DMI is heritable ($h^2 = 0.34.5$ with a range 0.11 - 0.58) (Koenen and Veerkamp, 1998; Berry et al., 2014) and the heritability varies over lactation (Manzanilla Pech et al., 2014; Li et al., 2017). Lastly, DMI is highly genetically correlated to milk production (r_g ranging from 0.20 – 0.73) but none of the estimated correlations are high enough to suggest indirect selection using milk production alone would be sufficient for improving feed efficiency (Spurlock et al., 2012; Manzanilla Pech et al., 2014).

1.4.2 Metabolic efficiency of lactating dairy cattle

Calorimetric methods have identified two sources of efficiency: one, maintenance efficiency i.e. improving milk production relative to energy required to maintain body tissue and two, metabolic efficiency i.e. more efficient partitioning of energy toward milk production and other traits of economic and energetic importance. As smaller body size has lower maintenance requirements, improving maintenance efficiency is readily achieved by the addition of selection against body size to a breeding goal, which already selects for improved milk production (Pryce et al., 2015). On the other hand, improving metabolic efficiency is more challenging as it requires expressing feed intake independent of maintenance requirements and other energy sinks.

The definitions of feed efficiency phenotypes currently under investigation can be broadly categorized into ratio- or residual-based traits (Kennedy et al., 1993). Inherent deficiencies in ratio traits have been thoroughly discussed elsewhere (Veerkamp and Emmans, 1995; Connor et al., 2012) and will only be briefly addressed here. Ratio based traits, like gross feed efficiency (milk production/feed intake), fail to take into account variation in other energy sinks that are not explicitly included in the calculation. Critically, ratio traits are not genetically independent of the component traits and are rarely linearly associated with component traits (Sutherland, 1965). Ratio traits consequently cannot distinguish between maintenance efficiency and metabolic efficiency. Not surprisingly, including biased correlations based on non-linear relationships into a selection index results in suboptimal or unpredictable response to selection (Gunsett, 1984; Zetouni et al., 2017; Shirali et al., 2018).

Residual based traits are a form of a restricted selection index where DMI is phenotypically or genetically independent on production traits or energy sinks (Kennedy et al., 1993). The most widely used method of estimating residual feed intake is using a phenotypic multiple regression of production traits on DMI and extracting the random residuals called phenotypic residual feed intake (RFI_p). This method, in essence, removes the phenotypic variation between DMI and production traits and thus, the phenotypic correlations between DMI and production traits are expected to be zero but genetic correlations may deviate from zero (Veerkamp et al., 1995; Huttman et al., 2009; Shirali et al., 2018). Extension to this two-step method are the single step RFI_p models, which can account for heterogeneity between research stations and different dietary rations (Tempelman et al., 2015). This phenotype is shown to be repeatable within lactation ($t = 0.46 - 0.90$) (Connor et al., 2013; Tempelman et al., 2015; Byskov et al., 2017) as well as over lactations ($t = 0.20$) (Connor et al., 2013). Furthermore, it is heritable ($h^2 = 0.02 - 0.25$) (Pryce et al., 2014; Tempelman et al., 2015) and also changes over lactation (Li et al., 2017). In many instances, RFI_p remains genetically correlated to energy sinks such as milk production ($r_g = -0.16 - -0.53$) (Veerkamp et al., 1995; Huttman et al., 2009; de Haas et al., 2011; Spurlock et al., 2012) and so careful weighting of RFI_p is still required in the selection index and makes defining correlations with other traits challenging.

A method closer to the selection indices used in practice is genetic RFI (RFI_g) (Kennedy et al., 1993; Lu et al., 2015). In this method, DMI rendered genetically independent of component traits, by including DMI and energy sinks such as body weight, growth, body composition and milk production in a multi trait model the estimated genetic (co)variances can be used to derive a restricted selection index (i.e. RFI_g). This allows for the partitioning of genetic components of metabolic efficiency and maintenance efficiency. This has led to increased focus on DMI directly as a means to improve feed efficiency, as selection for decreased DMI is a selection index is essentially equivalent to selection for RFI_g (Veerkamp et al., 2013). However, it is still of interest to researchers to assess the genetic parameters of RFI_g to ascertain the heritability and genetic correlations with other traits of interest. RFI_g can be equivalent to RFI_p in the special case that the genetic relationships between DMI and energy sink traits are identical to that of the phenotypic relationships (Kennedy et al., 1993). RFI_g is found to repeatable within lactation ($t = 0.54$) and over lactations ($t = 0.21$), heritable ($0.14 - 0.41$) and genetically correlated to RFI_p (0.93) (Lu et al., 2015, 2016). A prerequisite of both RFI_p and RFI_g is that suitably large numbers of animals with records are required to accurately determine the partial efficiencies between DMI and energy sink traits.

1.5 Metagenomics of the rumen microbiome

The microbes inhabiting the rumen, henceforth referred to as the rumen microbiome, have gained considerable research interest in recent years as is evidenced by reviews to which the reader is directed (Denman and McSweeney, 2014; Bickhart and Weimer, 2017; Tapio et al., 2017; Wallace et al., 2017). This is because the rumen microbiota form the 'missing' link between the host and its diet, their being responsible for the degradation of feed into end products like VFAs and microbial protein which is utilised by the host and end products like CH₄ released into the environment (Sutton, 1985; Denman and McSweeney, 2014). Initial characterisation of the rumen microbiota was conducted using microscopy and culture based techniques which was further stimulated by the development of a culture medium of sterile rumen fluid (Hungate, 1947). Culture based techniques allow for the classification, functional annotation and genomic sampling of a particular organism. Currently, the Hungate1000 project has collected 410 rumen bacterial and archaeal pure cultures and have sequenced and catalogued the genomes (Seshadri et al., 2018). It is estimated that 10-15% of rumen bacteria, for example, are culturable (Morgavi et al., 2012; Denman and McSweeney, 2014).

The standard for culture independent microbial ecology and diversity is sequencing of single genes encoding ribosomal subunits, such as the 16S rRNA gene, circumventing the need for laboratory pure culture (Head et al., 1998). Sequencing and alignment of variable portions of highly conserved genes such as the 16S and 18S rRNA can be clustered, based on sequence identity, into operational taxonomic units (OTUs) which are indicative of species or strains. This is the most cost effective and a high throughput means of genotyping microbial populations (Bickhart and Weimer, 2017). The reference sequences of OTUs are then assigned taxonomy against open online databases generated from name sequences such as those from pure cultures. A limitation on single gene sequence techniques in the rumen is the paucity of cultured and functionally annotated microbiota, with as many as 74% of OTUs remain unassigned at genus level in some studies (Kong et al., 2010). In addition, full genome sequence studies have found that mesophilic anaerobic bacteria (such as those in the rumen) horizontally exchange genes between different phyla, exchanging as much as 35% of metabolic genes (Caro-Quintero and Konstantinidis, 2014), which cannot be captured by simply sequencing the ribosomal

genes. Costlier and thus, lower throughput limiting methods such as metagenomics (full genomic sequencing) or metatranscriptomics (full sequencing of transcribe rRNA) can overcome some of the limitations to single gene sequencing such as redundancies due to horizontal gene transfer (Bickhart and Weimer, 2017; Wallace et al., 2017). With decreasing costs of sequencing, these methods may become the standard for high throughput genotyping of microbial populations.

There are potential sources of error in ribotyping studies, for instance, up to 100 fold differences in microbial abundances have been detected between different DNA extraction methods (Henderson et al., 2013). Differences between 16S rRNA abundances have been detected between the fluid, solid and epimural phases or rumen contents in some studies (Kong et al., 2010; Henderson et al., 2013) but not in others (Ji et al., 2017). Subsequent metagenomic approaches quantified the abundance variations at 0.07 – 0.48% between sample positions top, middle and bottom of the rumen (Ross et al., 2012). Few differences have been reported between communities sampled using rumen oral-esophageal tubing and rumen fistula (Henderson et al., 2013), whereas significant differences have been detected between oral-oesophageal tubing and non-invasive buccal swabs, although these differences explained variation is minimal 0.26% (Kittelmann et al., 2015). A careful design of sampling experiments are required to control or mitigate these sources of variation and standardised protocols are needed if genetic investigations are the primary focus.

Crucially, the spatial and temporal changes in the OTUs in an animal can be associated with biological changes in the animal (Morgavi et al., 2012). This has led to some key developments in ruminant microbial ecology: first, that rumen microbial composition changes depending on the sampling site, solid, liquid or epimural phase (Kong et al., 2010). Second, a core microbiome is present across cows across a broad geographical and dietary range (Henderson et al., 2015). Third, that microbial communities are host specific, reverting almost completely back to the original composition after near complete rumen transfaunation (Weimer et al., 2010). Fourth, communities can change in response to host related changes such as age or lactating vs non-lactating (Zhu et al., 2017). Finally, that communities can change in response to changes in feeding and diet (Tajima et al., 2001).

1.5.1 Associating rumen microbes with the host phenotype

A fundamental challenge in rumen microbiology is associating differences in rumen microbial composition and differences in phenotypes of the host and linking this with the biological and functional roles of individual microbes (Bickhart and Weimer, 2017). In the case of feed efficiency, early studies have detected

significant associations between numerous OTUs, species or higher taxonomic levels in dairy cattle (Jami et al., 2014; Jewell et al., 2015; Shabat et al., 2016) and beef cattle (Guan et al., 2008; Hernandez-Sanabria et al., 2010; Myer et al., 2015). In the case of feed efficiency, direct knowledge of specific biochemical pathways in the rumen contributing to differences in host feed efficiency phenotypes are not known. Furthermore, taxonomical classifications and functional annotations for many rumen microbiota are severely limited, thus linking the functional role of specific microbiota with differences in the host phenotype is extremely challenging. In the case of CH₄ production, the methanogenic pathways in archaea and some bacteria are well characterized. Initial results suggested that the abundance of archaea would be directly proportional to CH₄ production as methanogenesis is the primary pathway for ATP synthesis in most archaea (Wallace et al., 2014). Subsequently, many studies have not detected associations between archaeal abundance and CH₄ production, but have detected associations with non-methanogenic bacteria (Zhou et al., 2011; Danielsson et al., 2017). This could partly be due to the discovery of rumen archaea which metabolise methyl containing compounds and alcohol (Leahy et al., 2010; Poulsen et al., 2013), or studies demonstrating the same archaeal and *McrA* gene abundance, but differential gene expression in high and low emitting sheep (Shi et al., 2014). The mean archaea to bacteria ratio was from sire progeny groups was found to rank the same as CH₄ emission in 50 crossbred beef steers resulting from a two breed rotational cross for Limousin and Angus (steers were 2/3 and 1/3 of either breed) in respiration chambers (Roehe et al., 2016). Suggestive of host additive genetic influence on the microbiome and that bacteria and archaea affect host CH₄ emission (Roehe et al., 2016). These results indicate that there is some host phenotype (P) by rumen microbes interactions (M).

1.5.2 Rumen microbes and diet

Dietary strategies for CH₄ mitigation have been an intense area of research and are the subject of frequent review (Hristov et al., 2013; Knapp et al., 2014; Patra, 2016). These include plant, fungal and macro-algae metabolites, nitrates, nitrooxy and halogenated compounds as well as ionophores. Some of these strategies target the rumen microbes directly for example by inhibiting protozoa and associated archaea serve as hydrogen sinks or are toxic to archaea (Patra, 2016). In the case of ionophores reducing methane emissions also resulted in increased feed efficiency in dairy cattle (Duffield et al., 2008). The majority of these strategies have temporary effects and are not cumulative, but this could be overcome by rotating different additives (Klop et al., 2017). The temporary effects of these strategies suggest rapid microbial community adaption to substrate perturbation (Makkar and Becker, 1997;

Klop et al., 2017). Conversely, 3-nitrooxypropanol is an additive with a narrow spectrum of specificity inhibiting the enzyme methyl co-enzyme-M reductase, the last step in the methanogenic pathway (Duval and Kindermann, 2012). This compound has been shown to persistently reduce CH₄ production in dairy cows by 30% (Hristov et al., 2015a). These results indicate that there is a diet (D) by microbes (M) interaction.

1.5.2 Rumen microbes and the host

The near complete reversion of rumen bacteria to pre-switch composition after 95% rumen contents switch between cows was the first evidence of host specificity (Weimer et al., 2010). Subsequently, Roehe et al., (2016) sequenced the full rumen metagenome of eight crossbred beef steers divergent for CH₄ emission and detected associations between microbial-host cross talk genes *TSTA3* and *Fucl* with host feed conversion ratio. Gonzalez-Recio et al. (2018) found significant associations between relative abundance of rumen microbiota and the first two principal coordinates of a genomic relationships matrix in ten Holstein and 8 Brown Swiss cows. A study involving 48 Israeli Holsteins detected potentially heritable rumen bacteria with 22 rumen bacterial OTUs with heritability estimates in excess of 0.7 (Sasson et al., 2017). Although the validity of heritability estimates in such a small number of animals can be called into question, these results do indicate some host (H) by M interactions.

This evidence would suggest that differences between cows in complex phenotypes (i.e. methane emission and feed efficiency) can be explained by host (H) x microbes (M) x diet (D) interactions (Holobiont concept) (Bickhart and Weimer, 2017). What is needed is substantially larger datasets to determine and quantify the relative contribution of host genetics to rumen microbes and their joint contribution to host phenotypes in order to identify phenotypes influenced by the collective host and microbiome as well as new opportunities for phenotype improvement. Under holobiont theory, the collective unit of host and associated microbes are considered to contribute to differences in host phenotypes upon which natural selection and random genetic drift can act (Zilber-Rosenberg and Rosenberg, 2008; Rosenberg and Zilber-Rosenberg, 2011; Bordenstein and Theis, 2015). There is literature to suggest co-evolutionary genomics of binary (host with specific microbe) host-microbes interactions, however evidence from complex microbial communities (like the rumen microbiome) and the host are lacking. (Yeoman et al., 2011; Bordenstein and Theis, 2015). The consequences of ignoring host-microbe interactions when breeding for CH₄ emissions or feed efficiency could result in selective breeding against the every successful symbiosis between ruminant and rumen microbes,

leading to suboptimal or unpredictable responses to selection, conversely heritable microbes have the potential to further improve breeding from complex traits which are under their influence.

1.6 Research Aims:

The main objective of this PhD project is to investigate and contribute to resolving the lacking understanding of genetic variation behind methane emission, feed efficiency and the rumen microbiome in dairy cattle, by applying integrative and multidisciplinary approaches to phenotype recording. To solve this major task, the PhD project was divided into two main research areas:

- (i) Development and validation of a high throughput non-invasive method for recording CH₄ emissions under commercial conditions (Chapters 2-3); and
- (ii) Investigate the genetic relationships between CH₄ emissions and the interrelations with feed efficiency (Chapter 4) and the rumen microbiome (Chapter 5).

A total of four papers are presented as chapters in this PhD dissertation (full publication list is available in Curriculum Vitae) and the aim of each study is presented below.

Chapter 2 *'Interchangeability between methane measurements in dairy cows assessed by comparing precision and agreement of two non-invasive infrared methods'* The aim was to develop and validate a fast and efficient method for detecting and correcting fixed and variable shifts in times series. To compare two sniffer instruments for CH₄ and CO₂ breath concentrations in order to identify and quantify sources of disagreement and thereafter standardised instruments for joint recording.

[Published]

Chapter 3 *'Rapid Communication: Ranking cows' methane emissions under commercial conditions with sniffers versus respiration chambers'* The aim was to assess the consistency in ranking of dairy cows for CH₄ emission phenotypes obtained using sniffers and respiration chambers.

[Manuscript draft]

Chapter 4 *'Can greenhouse gases in breath be used to genetically improve feed efficiency of dairy cows?'* The aim was to estimate genetic parameters for feed intake

related traits and greenhouse gas traits in primiparous Holsteins in Denmark and The Netherlands. Secondly, assess the use of breath concentration GHG traits as large-scale indicator traits for improving the accuracy of breeding values for DMI and genetic residual feed intake.

[Manuscript draft]

Chapter 5 ‘Host genetics and the rumen microbiome jointly associate with methane emissions in dairy cows’ The aim was to detect and quantify associations between rumen microbes and CH₄ production. To assess the heritability of rumen bacterial and archaeal relative abundance

[Published]

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Chapter 2

Interchangeability between methane measurements in dairy cows assessed by comparing precision and agreement of two non-invasive infrared methods

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Abstract

In this study we assess the interchangeability and statistical agreement of two prevalent instruments from the non-invasive “sniffer” method and compare their precision. Furthermore, we develop and validate an effective algorithm for aligning time series data from multiple instruments to remove the effects of variable and fixed time shifts from the instrument comparison. The CH₄ and CO₂ gas concentrations for both instruments were found to differ for population means ($P < 0.05$) and intra-cow variation (precision) ($P < 0.05$) and for inter-cow variation ($P < 0.05$). The CH₄ and CO₂ gas concentrations from both instruments can be used interchangeably to increase statistical power for example, in genetic evaluations, provided sources of disagreement are corrected through calibration and standardisation. Additionally, averaging readings of cows over a longer period of time (one week) is an effective noise reduction technique which provides phenotypes with considerable inter-cow variation.

Key words: Equivalence, Instrument comparison, Interchangeability, Methane, Sniffer, Statistical agreement

2.1 Introduction

Methane (CH₄) is an abundant and potent greenhouse gas with a global warming potential substantially larger than that of carbon dioxide (CO₂) (IPCC, 2014). Dairy cattle through enteric methanogenesis contribute up to 20% of global livestock greenhouse gas emissions (Gerber et al., 2013). Research into mitigation strategies such as nutritional additives, housing, vaccination and genetic improvement has gained impetus in recent years. The assessment of strategies requires accurate and repeatable individual measurements under commercial conditions. Multiple instruments and or techniques have been developed to measure enteric CH₄ intensity and emissions from cattle and other ruminants, each with their own scope of applications, merits and demerits (Hill et al., 2016). No single method is perfect in all aspects and thus in many instances a reference method from which to make comparisons is lacking. An emerging method for the measurement of CH₄ and CO₂ concentrations in the breath of dairy cattle, which is high throughput, non-invasive and viable in commercial conditions, is the “sniffer” method (Lassen et al., 2012; Garnsworthy et al., 2012a). Air is continuously sampled from the concentrate bin of automated milking systems (AMS) during individual milking and sample gas concentrations recorded. Two prevalent instruments are the Gasmeter DX-4000 (Gasmeter; Gasmeter Technologies Oy, Helsinki, Finland) (Lassen et al., 2012; Haque et al., 2014) and the Guardian NG/Gascard (Guardian Plus; Edinburgh Instruments Ltd., Livingston, UK) (Garnsworthy et al., 2012a,b; Bell et al., 2014a,b). While the techniques and calculations differ, with the former employing a prediction equation based on the ratio of the two gas concentrations and production traits (Madsen et al., 2010) and the latter utilising a scaling factor and methane emission rate (Garnsworthy et al., 2012a), both methods rely on gas concentration readings. The cost of non-invasiveness is restricting the animal to instrument interface and introducing sources of error and imprecision between readings due to air turbulence within the AMS and movement of the cows head in the AMS concentrate bin (Huhtanen et al., 2015). Repeating spot samples over a number of days to obtain a phenotype e.g. average gas concentrations over a week, reduces sources of error by a function of $1 + t(n - 1)/n$ where t is the intra-class correlation (repeatability) and n the number of records; thus obtaining a representative value capable of ranking animals (Hegarty, 2013; Hill et al., 2016). No comparative studies have

been conducted on the two instruments to determine their equivalence or lack thereof.

Assessing the statistical agreement between instruments is crucial to informing the manner in which information from multiple instruments can be combined, for example, towards genetic evaluations. When measurements from both instruments on large-scale numbers of individuals are available, a genetic correlation between methods exceeding 0.8 is suitable to ascertain equivalence for genetic evaluations (Robertson, 1959). However, one may wish to establish agreement or lack thereof prior to measuring large numbers of individuals. According to Barnhart et al. (2007a), methods may disagree due to different population means, differing between-subject variances and differing within-subject variances. Population means can be corrected through calibrations, but different variances either require the more cumbersome instrument variance reduction or standardisation techniques (Barnhart et al., 2007a). As Bland and Altman (1999) pointed out, the partitioning of the random error variance into within-subject variances (imprecision) cannot be done without replicate measurements per subject per instrument. Analysing replicate measures on cows from AMS is challenging as the number of visits per cow to the AMS (replicates) is variable. Furthermore, time has elapsed between measures and thus the underlying biology has changed between measures due to factors such as diurnal variation patterns of CH₄ and CO₂ concentrations (Lassen et al., 2012). Thus replicate measures per cow must be taken simultaneously with each instrument and treated as paired observations i.e. “linked” replicates (Carstensen, 2011).

Choosing the correct indices to assess agreement must be done with care, for instance, despite having been discouraged for decades as being irrelevant and misleading, some authors still compute Pearson’s correlation coefficient in method comparison studies (Altman and Bland, 1983; Bland and Altman, 1986; Carstensen, 2011). Even if one of the methods is perfect, it will correlate poorly to a second less precise method (Barnhart et al., 2007a). Likewise, unscaled agreement indices such as the coefficients of variation within- and between-animals, as well as scaled agreement indices such as Lin’s three forms of concordance correlation coefficients (CCC) and intra-class correlations coefficients (ICC), are reliant on between-animal variance. Therefore, imprecise methods recorded on heterogeneous populations will still appear to agree favourably (Barnhart et al., 2007a). An agreement index suited to repeated measures with large errors and less reliance on population heterogeneity is the coefficient of individual agreement (CIA) (Barnhart et al., 2007b). Methods are regarded as interchangeable only if individual measurements

between instruments are similar to replicated measures within instrument (Barnhart et al., 2007a).

An additional challenge when comparing instruments with time-stamped measurements, is the clock synchronisation problem, where clocks can have fixed and variable shifts in time (Ridoux and Veitch, 2007). In the absence of synchronised time stamping, as is often the case when comparing readings from multiple instruments, it is possible to obtain a misleading result. Even the most precise instrument will compare poorly when timestamped by an inaccurate clock.

The objectives of this paper were: (1) Demonstrate a fast method for detecting fixed and variable shifts in time series. (2) Conduct a method comparison analysis in the presence of linked and variable number of replicates from each instrument. (3) Standardise instrument recordings to achieve satisfactory agreement for joint analysis.

2.2 Materials and methods

2.2.1 Design animals and feeding

Data was recorded over a three week period from end of April to mid May 2015 at the Danish Cattle Research Centre (DCRC, Foulum, Denmark). A total of 56 Holstein cows, average body weight 686.6 ± 86.5 kg (mean \pm sd), milk production 38.4 ± 0.34 kg/day roughage dry matter 20.47 ± 4.43 kg/day and concentrates 2.5 ± 0.28 kg/day were recorded during the experimental period. Cows were of mixed parity 44% 1st parity, 35% 2nd parity, 21% 3rd parity at mixed stages of lactation 36% early, 27% mid and 38% late (14–100 DIM early, 100–200 DIM mid, 200–305 late). The DCRC barn is a free stall housing system with cubicles. Cows had access to an AMS (DeLaval International AB, Tumba, Sweden) where they were provided up to 3 kg of concentrate a day within the concentrate bin. Cows were offered a TMR consisting of corn silage, rapeseed meal and soybean meal ad libitum in individualised feeding troughs (RIC-system, Insentec, Marknesse, The Netherlands). Data on feed intake (concentrate and roughage), weight and milk production are recorded continuously at the DCRC. The study was conducted without altering management protocols or feeding schemes conducted at the research centre. Cows had free access to AMS with a minimum visit cycle limitation of 4 h, except during the two daily automated cleaning cycles. Cows presented for milking on average 2.4 ± 0.86 visits/day (mean \pm sd) during the trial period. The data in this study is generated on cows performing under typical commercial conditions which are representative of a general dairy cattle population in Denmark.

2.2.2 Breath sampling analysis

CH₄ and CO₂ gas concentrations are routinely analysed at DCRC by using infrared gas analysers installed within each AMS (Guardian NG/Gascard, Edinburgh Instruments Ltd, Livingston, UK) with a range of 0–1% CH₄ and 0–5% CO₂ and logged with NOVUS FIELD LOGGER software (NOVUS Automation, www.fieldlogger.net). The air inlet was custom installed in the upper left rear side of the AMS feed bin so as to be aligned with the nostrils of a feeding cow as per the second experiment described by Garnsworthy et al. (2012a). Air is sampled continuously at a rate of 1 L/min through a 4 mm polyurethane tube approximately 3 m in length with an inline particulate filter to remove dust and a permeable tube with pressurised dehumidified air to remove water vapour before reaching the sensors. The exhaust port of the analyser is vented a minimum of 3 m clear of any sampling point. Data is logged at 1 s intervals and stored perpetually through the use of remote access. Sensors were calibrated prior to the experiment by flushing the sensor inlet with a calibration gas containing 0.0% CH₄ and 0.0% CO₂ to set the lower range and then flushed with a calibration gas containing 1.0% CH₄ and 3.0% CO₂ to set the upper range (both gases in synthetic air HiQ 4.0; AGA, Fredericia, Denmark). Sensors were installed for the recording of entry and exit times within the same time series as the continuous gas concentrations.

The portable Fourier transformed infrared analyser FTIR (Gasetm DX 4000, Gasetm Technologies Oy, Helsinki, Finland) was installed at DCRC as per methods used for sampling through-out Denmark (Lassen and Løvendahl, 2016). The inlet was installed within the feed bin of the AMS in a manner analogous to the Guardian with the exception that the location was in the upper right rear of the bin to mirror and the inlet of the Guardian, in order to prevent the differential pumping rates from creating turbulence at the inlets. Air was sampled continuously through the integral pump at a rate of 4 L/min, starting with an inline particulate filter at the inlet via a 5 m long hose heated to 180 °C before entering the sensor unit. The exhaust gases were vented more than 3 m away from any sampling points. Data was logged continuously at 5 s intervals using Calcmeter Software and stored on an integral laptop, thus the Guardian and Gasetm data was timestamped by different data logger software on different servers. The analyser provides reading for the multiple gases as well as water vapour, external and internal temperatures, CH₄ at a range of 0.0–0.5% and CO₂ at a range of 0.0–5.0% (Teye et al., 2009). Prior to the start of the experiment, the inlet pipe heating elements are heated to 180 °C and the sensor chamber is flushed out with 100% nitrogen gas for 0 point calibration. The identification number (ID) of the cow, start and exit time of the first visit to the AMS

after installation and calibration, is recorded to facilitate alignment with AMS data and determine initial fixed differences between time series.

The AMS records cow ID as well as total yield, status of visit, entry and exit times. Thus the AMS data is a third time series, time stamped by a separate server. Twice daily the AMS enters into an automated cleaning cycle of approximately 30 min in duration, during which no cow can enter the AMS. The cleaning cycles occur between 3:30 and 5:00 as well as 15:50 and 17:00 daily. The values recorded during the morning cleaning cycle are used to determine the ambient concentrations (Baseline) of CO₂ and CH₄ for each day.

2.2.3. Algorithm steps for prediction of cow entry times and head lifting

In order to align all three time series datasets to a common time, it was first necessary to account and correct for the fixed or variable time shifts between the servers used to timestamp the records. Since a cow exhaling into the concentrate bin of the AMS is accompanied by sudden and large changes in the concentrations of gases recorded by both instruments, it is possible to identify cow entry times in the datasets of the two instruments and align with the entry times of the AMS dataset. The algorithm used to identify cow entry was undertaken in progressive steps in the analysis. Firstly, the algorithm used Holt's double exponential smoothing to establish a time step to time step smoothed value, slope and forecast of CO₂ gas concentrations for both instruments. Then the detection of cow entry was based upon the smoothed value and slope. An added benefit of the algorithm is it enables the filtering of data when the cow is deemed (indirectly) to have lifted its head in a manner similar to that implemented by Garnsworthy et al. (2012a) and Huhtanen et al. (2015). See Appendix A for a detailed description of the algorithm.

2.2.4. Gas concentration variables

Both the Guardian and the Gasmeter series were brought into time alignment with the AMS dataset and expanded. The data where the cow's head is predicted to be out of the feed bin was filtered and means of CO₂ and CH₄ concentrations for each AMS visit were calculated for further analysis. Furthermore, the AMS time for the start of each visit was converted to 24 h angular radians and modelled in a Fourier series approach and ambient gas concentrations from the AMS morning cleaning cycles stored (Løvendahl and Bjerring, 2006; Lassen et al., 2012).

2.2.5. Statistical analysis

Gas concentration data from both instruments was transformed by means of a natural logarithmic transformation to meet the requirements of normality of residuals. Data was analysed using linear mixed effects (LME) models with Kronecker product covariance structure in a doubly multivariate set-up (Hamlett et al., 2004; Roy, 2009), by means of the MIXED procedure of SAS with repeated measures linked through time (ver. 9,3; SAS Institute Inc., Cary, NC). A Kenward–Roger correction was utilised for computing the correct denominator degrees of freedom of fixed effects in the presence of repeated measures (Kenward and Roger, 1997). As the number of replicate measures per cow per day can vary from 2 to 5 and may not be equal between cows, the number of replications by each method on each cow is equal to p_i . Therefore, the number of observations on the i th subject n_i is $2 p_i$. The LME for the i th cow can be written as:

$$y_i = \beta_0 + \beta_1 M_{i1} + \beta_2 M_{i2} + \beta_3 Date + \sum_{g=1}^3 [f_{1g} \sin \phi + f_{2g} \cos \phi] + \beta_6 Base + b_{1i} Z_{i1} + b_{2i} Z_{i2} + \epsilon_i$$

where every method replication response is denoted by subscript i , terms M_{i1} and M_{i2} indicate the instrument each response belongs to. The regression coefficients b_0 is the intercept, b_1 and b_2 are the respective fixed effects of instruments, b_3 is the fixed effect of date of measurement, f_{1g} and f_{2g} are fixed regression coefficients of Fourier series linear covariates of the time of day of measurement. Symbols ϕ denote the time of day as angular radians. b_6 is a fixed regression coefficient on the linear covariate of the daily background gas concentrations during the morning AMS cleaning cycle. Terms b_{1i} and b_{2i} are random effect parameters for each of the two instruments and Z_{i1} and Z_{i2} relate responses y_i to methods. The above equation can be re-written in matrix notation equations for further clarification:

$$y_i = X_i \beta + Z_i b_i + \epsilon_i$$

with $b_i \sim N_m(0, D)$

and $\epsilon_i \sim N_{n_i}(0, R_i)$

where (y_i) is all the responses of the i th cow, modelled as dependent on systematic effects b given design matrix of fixed effects X_i , described for the equation above. Random effects of each instrument across cows is given by b_i with incidence matrix Z_i . Variance–covariance matrix D gives between cow variation and variance–covariance matrix R_i gives within cow sources of variation.

The method described in detail by Roy (2009) was employed to formally test the significance of differences for the between- subject, within-subject and total variances of the two instruments by means of log likelihood ratio tests between models differing for various combinations of structured or unstructured variance–co-variance matrices D and R_i . The CIA was calculated as per Barnhart et al. (2007a) for comparisons without a reference method.

Once sources of disagreement were identified, the (y_i) responses of the Gasmeter were calibrated and standardised to that of the Guardian mean and total variance for both gases by means of STANDARD procedure in SAS. The aforementioned models were re-run to evaluate the effects of calibration and standardisation on measurement agreement, the results of which are summarised in Table 2.2. Furthermore, the sum of the intercepts, random solutions and residuals for each instrument and gas variable was taken as the per visit gas concentration corrected for daily background, time of day and date. The corrected per visit gas concentrations were then used to calculate the mean corrected gas concentration per day and the mean corrected daily concentrations per week. A reduced version of the model equation above only including fixed effects of day or week of measurement was utilised to estimate means and variances of the averaged gas phenotypes, the results of which can be found in Table 2.3 below.

2.3. Results and discussion

2.3.1. Time series alignment

The alignment algorithm correctly identified a fixed time difference of 1 h between the Guardian time series and AMS time series due to inconsistent Daylight Saving Time (DST) implementation. Furthermore, a 320 s fixed difference was correctly identified between the Gasmeter time series and AMS time series as per the differences between the start times recorded for the first cow after installation. The variable component of time drift between the time series of both instruments, the door sensors in the Guardian time series and the AMS time series can be seen in Fig. 1 below.

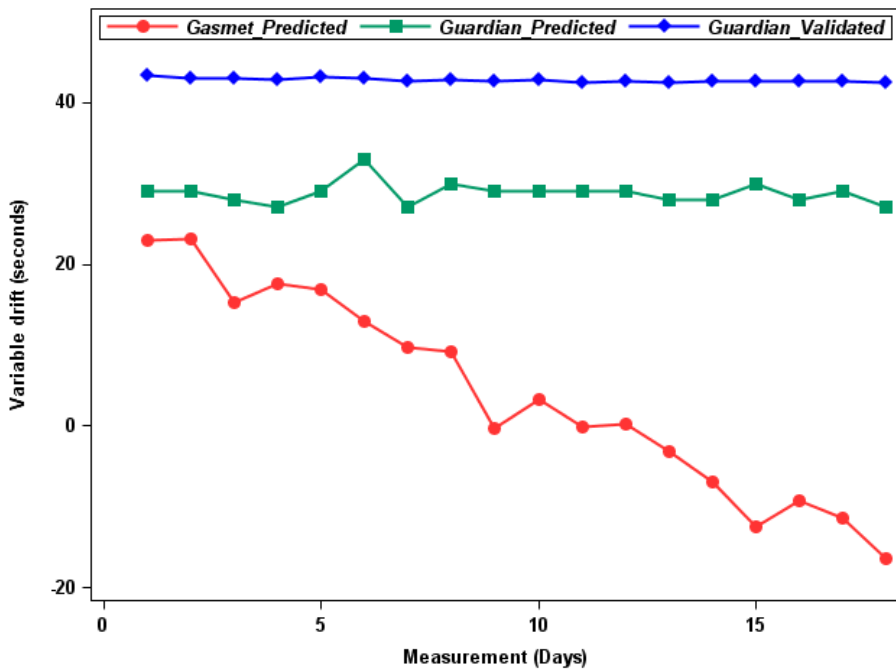


Figure 2.1 Variable clock drift between instrument time series (Gasmot_Predicted & Guardian_Predicted), Guardian door sensors (Guardian Validation) and AMS time in seconds.

For the Guardian, the prediction algorithm correctly identified the variable drift with a constant difference of 8 ± 4 s as compared with the door sensors in the AMS. The constant difference is attributed to the lag in time between the door opening and the first air sample containing the cow's breath reaching the gas sensor. The Gasmot time series drifted approximately 1 s per day whereas the Guardian time series drifted approximately 0.2 s per day. For the Guardian time series the detection rate of the alignment algorithm was verified at 97% for fixed and constant drift on a daily basis as compared to the door sensors. For the Gasmot only the observed (fixed) difference between the time series for the first day was possible. The alignment algorithm greatly improves the alignment of gas concentration signal with the associated cow accounting for fixed time shifts and variable time shifts. Thus the algorithm removes a potential source of bias and error from the subsequent instrument comparisons. A plot of sample data for CH_4 concentrations (PPM) for four cow visits from both instruments after alignment can be found in Fig. 2 below. A visual inspection of Fig. 2 reveals that whilst visible differences exist between

instrument readings, variable and fixed time drift are no longer a contributing factor to differences between instruments.

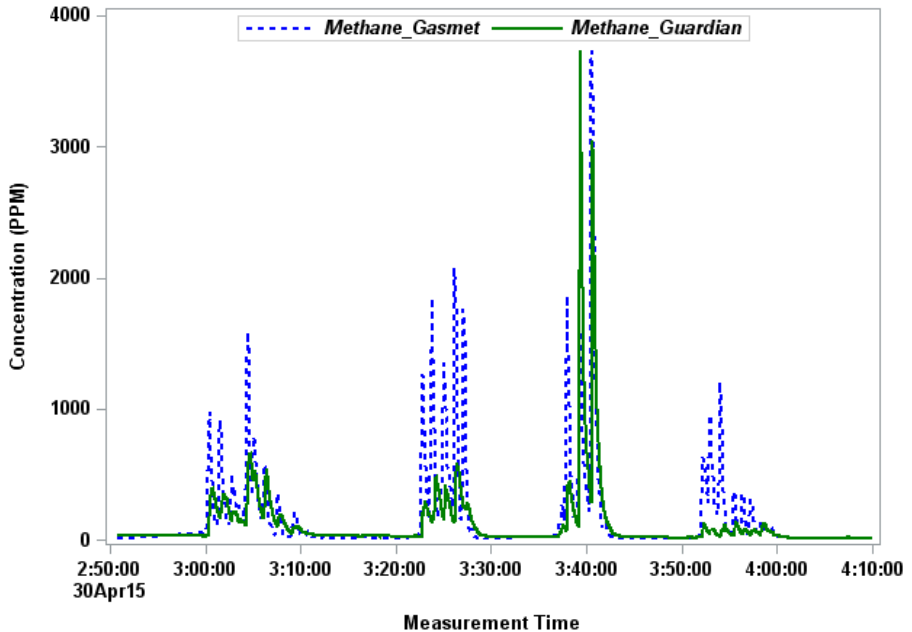


Figure 2.2 CH₄ gas signal for multiple cow visits after time alignment Guardian NG (Green solid line) Gasmet (Blue dashed line)

2.3.2. Instrument comparison, calibration and standardisation

The population means and all variance components differed ($P < 0.05$) for both gas concentrations between instruments (Table 2.1). As neither method is the reference method, we cannot report which instrument has the most accurate population means for either of the gas concentrations reported. The ICC estimates are moderate and similar in magnitude for each of the instruments and gas concentrations. Furthermore, they are directly comparable with those reported by Lassen et al. (2012) for the mean Gasmet gas concentrations on a per visit basis. Though discouraged in comparisons studies, we provide the Pearson's correlation coefficients, which were moderate and positive between instruments for both gases. Considering that all sources of (dis)agreement differed significantly, it is not surprising that the CIA estimates are moderate and low for CH₄ and CO₂ concentrations respectively, with the estimate for CO₂ being particularly poor and worse than for CH₄. Based on the criteria for statistical agreement of equivalence of

Table 2.1 Sources of (dis)agreement for CH₄ and CO₂ concentrations per AMS visit after common time alignment

| | Methane concentration ln(PPM) | | Carbon Dioxide concentration ln(PPM) | |
|--------------------------|-------------------------------|-----------------------------|--------------------------------------|-----------------------------|
| | Guardian | Gasmet | Guardian | Gasmet |
| Mean ± S.E. | 6.808 ^a ± 0.0176 | 6.564 ^b ± 0.0292 | 8.478 ^a ± 0.0371 | 9.128 ^b ± 0.0239 |
| Inter-Cow Variation | 0.017 ^a | 0.045 ^b | 0.075 ^a | 0.031 ^b |
| Intra-Cow Variation | 0.026 ^a | 0.091 ^b | 0.080 ^a | 0.035 ^b |
| Total Variation | 0.042 ^a | 0.136 ^b | 0.155 ^a | 0.066 ^b |
| Intra-Class Correlations | 0.39 ^a | 0.33 ^a | 0.48 ^a | 0.47 ^a |
| Pearson's Correlation | | 0.55 | | 0.44 |
| CIA | | 0.54 | | 0.18 |

^{a,b} estimates with subscripts differ (P<0.05)

means and total variability suggested by Roy (2009), and the threshold of “good” interchangeability (0.455) defined by Barnhart et al. (2007b), we can see that CH₄ measures of both instruments can be used interchangeably but the CO₂ cannot, without further remedy to the sources of disagreement.

After calibration and standardisation of the Gasmeter measurements to those of the Guardian it can be seen that the population means, inter-cow variation and total variation no longer differ ($P > 0.05$) (Table 2.2). While lack of a reference method precludes inference about the accuracies, inference about precision is possible using the intra-cow variation. The intra-cow variation differed ($P < 0.05$) for both CH₄ and CO₂ between instruments, with the Gasmeter having marginally higher estimates for both gases. The ICCs and Pearson’s correlations are very similar to those computed before calibration and standardisation as they are largely unaffected by calibrations and variance scaling. The CIA estimates for CH₄ and CO₂ increased substantially with 145% and 356% increases respectively, and were thus more similar for CH₄ than for CO₂. Using the criteria of statistical agreement and interchangeability we see that for CH₄ and CO₂ the instruments can be used interchangeably. It will thus be beneficial to combine datasets from both instruments in the interests of increasing statistical power.

The descriptive statistics, ICC and Pearson’s correlations for mean visit concentration per day and per week, after standardisation and calibration, are contrasted for CH₄ and CO₂ in Table 2.3 below. The population means do not differ for all pairs of gas concentration variables. The ICC estimates increase with each increase in time and thus, number of visits used to estimate mean values. The intra-cow variation decreased as a function of $1 + r(n - 1)/n$ where r is the intra-class correlation and n is the number of repeated measures thus, the effect of noise is reduced while the animal variance remains constant i.e. averaging out imprecision.

The ICC estimates for CH₄ and CO₂ of the Guardian are in the order of magnitude of those reported by Bell et al. (2014b) for a similar phenotype estimated on per week of lactation basis. Differences between instruments were expected as instruments differed on multiple technical aspects e.g. gas concentration sensitivity ranges, differential pumping rates, differential sensitivities to temperature and humidity, etc. Moreover the Guardian instrument in this study is dedicated solely to measuring CH₄ and CO₂ intensities, whereas the Gasmeter records and transforms IR spectra into ten different gases including CH₄ and CO₂. Auto-correlations between flanking spectral regions could be the cause of the marginal increase in imprecision observed with the Gasmeter instrument.

Table 2.2 Sources of (dis)agreement for CH₄ and CO₂ concentrations per AMS visit after common time alignment, calibration and standardization

| | Methane concentration ln(PPM) | | Carbon Dioxide concentration ln(PPM) | |
|--------------------------|-------------------------------|-----------------------------|--------------------------------------|-----------------------------|
| | Guardian | Gasmet | Guardian | Gasmet |
| Mean ± S.E. | 6.806 ^a ± 0.0178 | 6.791 ^a ± 0.0169 | 8.479 ^a ± 0.0169 | 8.439 ^a ± 0.0371 |
| Inter-Cow Variation | 0.017 ^a | 0.015 ^a | 0.074 ^a | 0.075 ^a |
| Intra-Cow Variation | 0.028 ^a | 0.031 ^b | 0.078 ^a | 0.086 ^b |
| Total Variation | 0.045 ^a | 0.046 ^a | 0.151 ^a | 0.161 ^a |
| Intra-Class Correlations | 0.38 | 0.33 | 0.49 | 0.46 |
| Pearson's Correlation | | 0.49 | | 0.42 |
| CIA | | 0.78 | | 0.65 |

^{a,b} estimates with subscripts differ (P<0.05)

Table 2.2 Average daily and weekly gas concentrations for standardised measurements

| | Methane Concentration | | Carbon Dioxide Concentration | |
|------------------------------------|--------------------------------|--------------------------------|--------------------------------|---------------------------------|
| | Guardian | Gasmet | Guardian | Gasmet |
| <i>Mean Visit per Day Ln(PPM)</i> | | | | |
| Mean \pm S.E. | 6.658 ^a \pm 0.018 | 6.641 ^a \pm 0.017 | 8.313 ^a \pm 0.037 | 8.272 ^a \pm 0.037 |
| Inter-Cow Variation | 0.017 | 0.016 | 0.074 | 0.072 |
| Intra-Cow Variation | 0.012 | 0.015 | 0.036 | 0.050 |
| Intra-Class correlations | 0.58 | 0.51 | 0.67 | 0.58 |
| Pearson's correlation | 0.45 | | 0.41 | |
| <i>Mean Visit per week Ln(PPM)</i> | | | | |
| Mean \pm S.E. | 6.655 ^a \pm 0.018 | 6.642 ^a \pm 0.017 | 8.307 ^a \pm 0.037 | 8.273 ^a \pm 0.0371 |
| Inter-Cow Variation | 0.016 | 0.016 | 0.074 | 0.071 |
| Intra-Cow Variation | 0.003 | 0.005 | 0.012 | 0.018 |
| Intra-Class correlations | 0.84 | 0.77 | 0.87 | 0.80 |
| Pearson's correlation | 0.49 | | 0.42 | |

^{a,b} estimates with subscripts differ ($P < 0.05$)

2.4. Conclusion

We detected fixed and variable shifts in time series between instruments and our algorithm was effective in aligning data series to a common clock time on a per day basis. Instead of genetic correlations between measurements exceeding 0.8 or simultaneous measurements with a reference method, we make use of statistical agreement and interchangeability to assess equivalence of methods. Measurements of CH₄ and CO₂ from both instruments can be used interchangeably provided sources of disagreement are identified and corrected for by means of calibration and standardisation. However, simultaneous repeated measures on multiple cows by both instruments are required to identify sources of disagreement. The ICC estimates show the inter-cow variation is considerable, particularly with measurements recorded over a one week period and is encouraging for genetic improvement of the phenotype. It is pertinent to note, the phenotypes estimated here are gas concentrations and not absolute emissions. Further comparisons between techniques which estimate emissions from concentration are required.

2.4.1. Conflicts of Interest

The authors have no affiliation with either company producing instruments and have no conflicts of interest to declare.

2.4.1. Acknowledgements

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2.5. Appendix

2.5.1. Algorithm steps for prediction of cow entry times and head lifting

Step 1: Holt's double exponential smoothing:

The smoothed value (\bar{L}_t), slope (\bar{b}_t), and forecast (\bar{F}_{t+1}) of CO₂ gas concentrations for both instruments for each time point were established:

$$\bar{L}_t = \alpha y_t + (1 - \alpha)(\bar{L}_{t-1} + \bar{b}_{t-1})$$

$$\bar{b}_t = \beta(\bar{L}_t - \bar{L}_{t-1}) + (1 - \beta)\bar{b}_{t-1}$$

$$\bar{F}_{t+1} = \bar{L}_t + \bar{b}_t$$

The smoothing factors α and β gives relative weights to new information and was determined by means of a grid search though parameter space 0 to 1 to increments of 0.1 to minimize the Mean Square Forecast Error (MSFE):

$$MSFE = \sum_{i=1}^t (\bar{F}_t - y_t)^2 / t$$

Step 2: Deviations from Smoothed Value:

For each time step, the deviation between the observed CO₂ concentration and the smoothed value was calculated as the square of the simple difference:

$$\bar{x}_t^2 = (y_t - \bar{L}_t)^2$$

Step 3: Detection of Cow entry:

The square deviations of smoothed value were expected to be nearly free of diurnal variations and thus reach 0 between cow visits. Therefore square deviations exceeding a threshold (δ) indicated a sudden change. A rule was applied whereby a reading exceeding the threshold with three past square deviations below the threshold, along with three consecutively increasing slope values, was defined as the start of a cow entry.

Step 4: Detection of cow head lifting:

A second threshold (ϕ) was defined as the upper 95% confidence interval from the ambient CO₂ concentrations found during the morning AMS cleaning cycles. If the slope increased for three consecutive time steps followed by three decreasing time steps, it was defined as a peak. Thus readings must contain peaks and the smoothed value exceeding the threshold (ϕ) to be defined as cow head inside the bin.

Step 5: Calculation of time displacement:

The time elapsed between each predicted cow entry point and the successive predicted entry point was extracted and stored in an $n \times 2$ array, where n is the total number of predicted visits. Similarly, the time elapsed between each AMS entry time and the successive entry time was extracted and stored in a separate $n \times 2$ array where n is the number of visits. For both arrays each row was summed and

match merged by use of SQL Procedure in SAS with a tolerance of 20 seconds and added limitation that matches must have the same date. Three consecutive correctly predicted cow entries are required to achieve a match with three consecutive AMS cow entry times. The difference between the times of the predicted cow entry and the AMS cow entry gives the time displacement between each time series. Time displacement estimates are further filtered by narrowing the range of displacement to +/- (01:01:00 hours) allowing for clock changes at Daylight Saving Time onset and reset. The mean of all displacement estimates for each day are then subtracted from the time variable of one of the datasets to bring the time series into 'common time' and allow merging by time.

Step 6: Optimization and validation:

A subset of the data from each instrument was used as a training set to optimize thresholds for cow entry prediction. This consisted of the 3:00 to 6:00 period during which a cleaning cycle occurred and readings are assumed to have fallen to ambient barn concentrations, and the 10:00 until 13:00 when many milkings follow in quick succession and thus readings between milkings are assumed to not always have fallen to ambient levels. If the threshold (δ) is set too high the detection rate in the later subset is low. Conversely, if the threshold is too low multiple false positives are predicted within each visit. Since the two scenarios are antagonistic, an optimization was performed. The $\alpha = 0.9$, $\beta = 0.1$, $\delta = 500$ and ϕ ranged between 600 and 1000 for both instruments.

The Guardian had integral door sensor readings from the gates of the AMS. A 2 x n array was constructed as per those above and match merged to the array from the AMS dataset. The displacement between the door sensors and AMS was then utilized as a known time displacement to validate the algorithm for the Guardian dataset only.

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Chapter 3

Rapid Communication: Ranking cows' methane emissions under commercial conditions with sniffers versus respiration chambers

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Abstract

The objective of the study was to assess the ranking of dairy cows using individual level correlations for methane (CH₄) emission phenotypes on-farm using sniffers and in respiration chambers. In total 20 lactating dairy cows, ten Danish Holstein and ten Danish Jerseys were recorded for three weeks of lactation prior to relocation and acclimation at the respiration chamber (RC) facility where they were each recorded on three occasions within the RC. Pairwise bivariate linear mixed models were used to determine the individual level correlations (r_i) between sniffer phenotypes and RC phenotypes as proxies for genetic correlations. Despite differences in feeding and management, the predicted CH₄ production on farm from sniffers correlated well with CH₄ production in the RC (CH₄_RC) $r_i = 0.77 \pm 0.18$ and the direct CH₄ breath concentration (CH₄_C) correlated nearly as well with CH₄_RC $r_i = 0.75 \pm 0.20$. The correlations between CH₄ emission phenotypes on-farm from sniffers and CH₄_RC exceeded that of energy corrected milk yield, live weight and dry matter intake demonstrating the potential of sniffers measurements as large-scale indicator traits for CH₄ emissions in dairy cattle.

Key words: Methane, sniffers, breath concentration, respiration chambers

3.1 Introduction

Methane (CH₄) is a potent greenhouse gas produced by dairy cattle and other ruminants as a natural by-product of fermentation. Research on numerous mitigation strategies such as nutritional additives, vaccines, and genetic improvement, has gained impetus in recent years (Hill et al., 2016). Whilst a reliable and accurate measure of CH₄ emission forms the basis of evaluating all of the aforementioned strategies, they differ in requirements of accuracy and precision as well as the number, frequency and duration of measurement (Hammond et al., 2016). For instance, genetic evaluations require measurements on large numbers of animals under the environmental conditions they are expected to perform (Falconer and Mackay, 1996).

Indirect calorimetry respiration chambers (RC) are the 'gold standard' for CH₄ emission meaning RC are regarded as the most accurate and precise measurement from which different instruments or techniques are benchmarked (Grainger et al., 2007; Hill et al., 2016). The high accuracy and precision of RC have the benefit of detecting relatively small effects of diets and treatments on CH₄ production with small numbers of animals (Patra, 2016). However, the cost for using RC are high and labor intensive, and are prohibitive to obtaining measurements on large numbers of animals (Grainger et al., 2007; Madsen et al., 2010). Furthermore, the effects of confinement within the chamber may alter animal behaviour and are not necessarily representative of all production systems like extensive grazing systems (Storm et al., 2012), although promising developments have led to reductions in costs and reduced stress due to confinement (Hellwing et al., 2012). Thus the merit of genetic selection under RC conditions and the expectation of reduced CH₄ emissions under environmental conditions in which animals are expected to perform, have been called into question (Lassen and Løvendahl, 2016).

A method that has proven well suited to obtaining large numbers of records on individual animals under commercial conditions is the high-throughput, cost effective, and non-invasive 'sniffers' installed in the feed bin of automated milking stations (AMS) or concentrate feeders (Garnsworthy et al., 2012; Lassen et al., 2012; Negussie et al., 2016). The non-invasiveness is achieved by limiting the animal-to-instrument interface, so that the animal is not aware it is being measured and measurement does not disrupt farm activities. The disadvantages of this are a loss of precision due to variable barn gas dynamics and added noise if head movement of the cow and background barn gas levels are not accounted for (Huhtanen et al., 2015; Difford et al., 2016; Wu et al., 2018). Furthermore, sniffer methods record the concentration of gases in the captured breath of the cow during

milking and are thus a spot measure of gas concentrations and not a full 24 hr mass flux measure (Huhtanen et al., 2015). Rather they utilise the ratio of measures CH₄ to carbon dioxide (CO₂) and predicted CO₂ as a tracer gas to approximate CH₄ production in litres per day (Madsen et al., 2010).

New methods which are cheaper, faster or less invasive are continually under development and are of value in situations where the gold standard proves expensive or prohibitive to recording on large numbers of animals. From the genetic evaluation perspective a new method can replace a gold standard method when their genetic correlation exceeds 0.80 (Robertson, 1959). Furthermore, if the genetic correlation is moderate to high and the new method is heritable, it can serve as an indicator trait (Negussie et al., 2017). However, the number of related animals with simultaneous measurements by both methods required to accurately estimate genetic correlations with meaningful standard errors, numbers in the order of thousands of animals (Visscher, 1998). Should measurements with both methods be made on different animals or animals at different points in time the number of animals required to accurately estimate genetic correlations will be far higher (Bijma and Bastiaansen, 2014). In the case of CH₄ emission, the RC constitute a separate environment as most other methods cannot be recorded simultaneously within the RC and remain representative. Genetic correlations remain elusive, despite calls for international collaborations and the need for genetic correlations between different methods of recording methane emission in dairy cattle (Pickering et al., 2013). It is of interest to researchers and stakeholders to determine the potential difference between methods in ranking of animals prior to the investment in thousands of records using both methods, particularly when the gold standard is expensive and cost limiting. To this end, estimating the repeatability of each method and the individual level correlations between methods is of value. The repeatability of a method serves as the upper threshold for heritability estimates and individual level correlations serve as proxies for genetic correlations, respectively (Falconer and Mackay, 1996; Wolak et al., 2012).

The objective of the present study was to assess the consistency in ranking of dairy cows for CH₄ emission phenotypes obtained using sniffers and respiration chambers.

3.2 Materials and methods

3.2.1 Design and animals

This experiment was designed to compare CH₄ measurements obtained by the non-invasive sniffer method during AMS milking of cows with the open circuit respiration chamber method, which is the intensive and traditional method. The link between the two methods was obtained by having 20 cows (ten Holstein and ten Jersey) measured first by the sniffer method in a semi-commercial setting and then transferred to the respiration chamber facility for repeated measurements during a feeding experiment.

3.2.2 Sniffer AMS measurement

Data on CH₄ and CO₂ gas concentrations from the breath of individual Holstein and Jersey cows recorded during milking at Danish Cattle Research Centre (DCRC, Foulum, Denmark), where sniffer sensors were installed in each of the three AMS (DeLaval International AB, Tumba, Sweden). The sniffer instrumentation is comprised of two sensors, the CH₄ sensor (Guardian NG, Edinburgh Instruments Ltd, Livingston, UK) and the CO₂ sensor (Gascard, Edinburgh Instruments Ltd, Livingston, UK). The equipment installation, technical specifications, and calibration procedures for the sensors are described elsewhere (Difford et al., 2016). The DCRC barn is a free-stall housing system with individual cubicles and two Holstein and one Jersey management groups, each with an AMS. Cows were offered TMR with an approximate forage to concentrate ratio of (70:30) ad libitum in individualized feeding troughs (RIC-system, Insentec, Marknesse, The Netherlands). Cows were provided up to 3 kg of concentrate per day (Table 3.1) within the feed bin of the AMS, based on levels of production and thus differences in the forage to concentrate ratio between cows is to be expected. Cows had free access to AMS and presented on average 2.4 ± 0.86 visits/d (mean \pm SD) during the period of measurement. Data on live weight is recorded 10 times every second during AMS milking and the processed as described in (Bossen et al., 2009). Milk production from the AMS and fat, protein and lactose percentage estimated from 48 hr periods each week (Løvendahl and Bjerring, 2006). The estimated milk components were used to correct milk production for fat, protein, and lactose content (ECM) (Sjaunja et al., 1991).

Gas concentrations measured in the on-farm AMS was aligned and merged with the entrance and exit time for each cow visit to the AMS. Data was omitted when the cow's head was predicted to be outside the feed bin using the algorithm described in Difford et al. (2016). The CH₄ and CO₂ gas concentrations for the

morning cleaning cycle, when the AMS is empty of any cow, were taken as the ambient barn concentrations for each day, and deducted from the means of each cow visit to obtain background corrected gas concentrations. The starting time for each visit in the AMS was converted to 24 h angular radians for further modelling of diurnal variation in a Fourier series approach (Lassen et al., 2012). The following model was used on all available DCRC data, within AMS to obtain daily visits corrected for sensor drift, daily variation, and time of day:

$$y_{ijklm} = \mu + d_i + b_j + \sum_{k=1}^3 (f_{1k} \sin \theta + f_{2k} \cos \theta) + C_i + e_{ijklm} \quad (1)$$

where y_{ijklm} is the natural logarithm of background corrected AMS visit means of CH₄ and CO₂; d_i is the fixed effect of test day i ($i = 23$ d); b_j is the fixed effect of the first full day after each calibration j ($j = 3$); f_{1k} and f_{2k} are the fixed regression coefficients of Fourier series linear covariates of the time of day of measurement, modelled as harmonic pairs. The time of day of visit expressed as 24 h angular radians is denoted by ϑ . Term C_i is the random effect parameter for each cow $C_i \sim ND(0, \sigma^2_c)$, and e_{ijklm} is the random residuals $\sim ND(0, \sigma^2_e)$. In order to correct daily AMS visit means, the random residuals for each visit are combined with random cow solutions, intercept, calibration day, and the regression coefficients f_{1k} and f_{2k} multiplied by the angular radian corresponding to 12:00:00 a.m. Further, the CH₄ concentration in parts per million and CO₂ in ppm as the average per week of lactation, was obtained by taking the corrected means of visits weighted by the number of seconds for each visit. The average weekly breath concentrations on CH₄ and CO₂ were natural log transformed, here after defined as CH₄_C and CO₂_C and combined with the weekly performance data from DCRC for ECM (ECM_C), LW (LW_C), and gestation length (GL). The average daily CH₄ production per week of lactation (CH₄_P; L CH₄/d) was calculated using the ratio of CH₄_C to CO₂_C and the equation for CO₂ production from heat production units utilizes ECM_C, LW_C, and GL (CIGR, 2002) and the conversion from heat production to CO₂ to obtain predicted CO₂ production (CO₂_P; L CO₂/d) (Pedersen et al., 2008) as suggested by Madsen et al. (2010). Data for the 10 Holstein and 10 Jersey cows from the last three weeks prior to relocation to the facilities with the RC was retained for further analysis together with RC records.

3.2.3 Respiration chamber measurements

The 10 lactating Holstein cows and 10 lactating Jersey cows were entered into a trial which contained two dietary treatments, a control and a high concentrate diet with respective forage to concentrate ratios of (68:32) and (39:61) in a cross-

over design with back-cross (Olijhoek et al., 2018). The trial was divided into three periods with blocking consisting of four cows per block (five blocks in total per period) with the same cows in each block over periods. Each cow was recorded for duration of 3 d in period 1 and 2 d in periods 2 and 3. Methane and CO₂ production in the RC (CH₄_RC and CO₂_RC) was calculated from the product of the total flow of outgoing air at standard temperature and pressure, and the difference between the gas concentrations in the outgoing air and the gas concentrations in the incoming air from the barn (background). Methane and CO₂ concentration readings from when the chambers were opened twice daily for milking were omitted before calculating the average CH₄ and CO₂ production for each cow, during each period. The average LW and ECM from the RC visits (LW_RC and ECM_RC) was retained for further comparisons. The records from these cows were retained for comparison with sniffer measures recorded at DCRC for the same 20 cows as in the RC.

3.2.4 Statistical Analysis

Pairwise bivariate animal repeatability models were performed to estimate the variance components and control for fixed effects for the RC traits and the on-farm traits. All analyses were performed using DMU version 6 (Madsen and Jensen, 2014). The model for the RC traits was as follows:

$$y_{ijklmn} = \mu + B_i + P_j + L_k + D_l + C_m + e_{ijklmn} \quad (2)$$

Where y_{ijklmn} is the trait of interest (CH₄_RC, CO₂_RC, LW_RC and ECM_RC), μ is the intercept, B is the i'th breed (I = 2 levels), P is the j'th period effect (j = 3 levels), L is the k'th lactation number (k = 3 levels), D is the l'th effect of diet (l = 2 levels), C_m is the random effect of the m'th cow $N(0, \sigma^2_c)$ and e is the random residual $N(0, \sigma^2_e)$.

The model for the on-farm traits was as follows:

$$y_{ijklm} = \mu + W_i + BR_j + L_k + C_l + e_{ijklm} \quad (3)$$

Where y_{ijklm} is the trait of interest (CH₄_P, CO₂_P, CH₄_C, CO₂_C, LW_C and ECM_C), μ is the intercept, W is the i'th week of lactation (I = 2 levels), BR is the j'th breed nested within AMS (j = 3 levels), L is the k'th lactation number effect (k = 3 levels), C_l is the random effect of the l'th cow $N(0, \sigma^2_c)$ and e is the random residual $N(0, \sigma^2_e)$.

For all pairwise comparisons between RC and on-farm traits it was necessary to restrict residual covariance to zero as cows were recorded in different environments. Repeatability estimates (t) were obtained from the variance components by using the equation (4):

$$t = \frac{\sigma_c^2}{(\sigma_c^2 + \sigma_e^2)} \quad (4)$$

Individual level correlations (r_l), were computed as the correlation between random cow effects using variance components as shown in equation 4:

$$r_l = \frac{\sigma^2_{c1,c2}}{\sqrt{\sigma^2_{c1}}\sqrt{\sigma^2_{c2}}} \quad (5)$$

The standard errors of the individual level correlations and repeatability estimates were derived using Taylor series approximations.

3.3 Results and discussion

The descriptive statistics for on-farm phenotypes and RC phenotypes can be found in Table 3.1. The CH₄ and CO₂ phenotypes followed a similar trend where the predicted mass flux phenotypes CH₄_P and CO₂_P were closer to that of RC mass flux phenotypes CH₄_RC and CO₂_RC but with lower means, higher variability and consequently higher coefficients of variation (CV). The breath concentration phenotypes were then further from the means of the RC and predicted mass flux phenotypes with lower means and higher variability and CV. All CH₄ and CO₂ phenotypes were moderately to highly repeatable ranging from $t = 0.53$ for CH₄_C to $t = 0.87$ for CO₂_P. Live weight and ECM were retained as control variables to ensure that the RC environment was not considerably different from that of the on-farm environment for these production traits. The means, SD and CV for LW and ECM were compared across environments with similar descriptive statistics and repeatability estimates for example LW_C and LW_RC had similar means 568.2 vs 564.5 kg and repeatability $t=0.93$ and $t=0.98$, respectively. Recognizing that in the case of the on-farm phenotypes, they are an average of many measurements from the AMS over a full week of lactation whereas RC phenotypes are the average of measurements over a 2 – 3 d period in the RC.

The individual level correlations between on-farm phenotypes and RC phenotypes are reported in Table 3.2. For the control phenotypes, ECM and LW the individual level correlations (r_l) between on-farm AMS and RC were close to unity 0.86 ± 0.15 and 0.92 ± 0.04 respectively. These results indicate very similar ranking of cows across environments and that the effect of confinement within the RC was likely small. Many authors have reported that confinement within RC alters behavior and can induce stress resulting in a drop in feed intake (Beauchemin and McGinn, 2006; Llonch et al., 2016). However, the RC in the present study were constructed from transparent polycarbonate to reduce costs and increase cow welfare, as supported by a study describing that these RC provoked no drop in DMI, when comparing before and during measurement in the RC (Hellwing et al., 2012).

3 Ranking with sniffers and respiration chambers

Table 3.1. Descriptive statistics for on farm measurements and respiration chamber measurements.

| | Unit | Mean | SD | CV (%) | t ¹ |
|--|---------|-------|--------|--------|----------------|
| On farm² | | | | | |
| CH4_P | L/d | 573 | 73.9 | 12.9 | 0.58 ± 0.11 |
| CO2_P | L/d | 6771 | 578.7 | 8.5 | 0.87 ± 0.11 |
| CH4_C | ppm | 410 | 137.0 | 33.4 | 0.53 ± 0.11 |
| CO2_C | ppm | 5746 | 1791.7 | 31.2 | 0.56 ± 0.12 |
| CH4/CO2_C | ppm/ppm | 0.071 | 0.009 | 12.7 | 0.38 ± 0.13 |
| ECM_C | Kg/d | 38.1 | 5.93 | 18.1 | 0.71 ± 0.08 |
| LW_C | Kg | 568.2 | 57.3 | 10.1 | 0.93 ± 0.02 |
| DMI_C | Kg/d | 22.2 | 4.9 | 21.5 | 0.73 ± 0.07 |
| Respiration Chamber³ | | | | | |
| CH4_RC | L/d | 521 | 56 | 10.7 | 0.61 ± 0.12 |
| CO2_RC | L/d | 6538 | 702.3 | 10.7 | 0.72 ± 0.10 |
| CH4/CO2_RC | L/L | 0.081 | 0.006 | 7.7 | 0.57 ± 0.14 |
| ECM_RC | Kg/d | 28.3 | 5.6 | 19.8 | 0.65 ± 0.12 |
| LW_RC | Kg | 564.5 | 62.3 | 11.0 | 0.98 ± 0.01 |

¹t = repeatability intraclass correlation coefficient. ²On farm phenotypes: CH4_P = predicted methane production; CO2_P = predicted carbon dioxide production; CH4_C = methane breath concentration; CO2_C = carbon dioxide breath concentration; CH4/CO2_C = ratio of methane to carbon dioxide breath concentration; ECM_C = Energy corrected milk yield; LW_C = live weight; DMI_C = dry matter intake. ³Respiration chamber phenotypes: CH4_RC = methane production; CO2_RC = carbon dioxide production; CH4/CO2_RC = ratio of methane to carbon dioxide production; ECM_RC = energy corrected milk yield; LW_RC = live weight.

Table 3.2 Individual level correlations between on farm phenotypes and respiration chamber phenotypes

| On farm ¹ | Respiration chamber ² | | | | |
|----------------------|----------------------------------|-------------|--------------|--------------|--------------|
| | CH4_RC | CO2_RC | CH4/CO2_RC | ECM_RC | LW_RC |
| CH4_P | 0.77 ± 0.18 | 0.63 ± 0.10 | 0.70 ± 0.24 | 0.68 ± 0.21 | -0.09 ± 0.29 |
| CO2_P | 0.74 ± 0.13 | 0.79 ± 0.14 | 0.41 ± 0.29 | 0.58 ± 0.22 | 0.20 ± 0.27 |
| CH4_C | 0.75 ± 0.20 | 0.80 ± 0.16 | 0.03 ± 0.39 | 0.21 ± 0.35 | 0.60 ± 0.22 |
| CO2_C | 0.62 ± 0.24 | 0.76 ± 0.18 | -0.35 ± 0.38 | 0.06 ± 0.40 | 0.69 ± 0.18 |
| CH4/CO2_C | 0.60 ± 0.27 | 0.29 ± 0.37 | 0.83 ± 0.23 | 0.68 ± 0.23 | -0.66 ± 0.24 |
| ECM_C | 0.66 ± 0.20 | 0.54 ± 0.23 | 0.52 ± 0.26 | 0.86 ± 0.15 | -0.14 ± 0.27 |
| LW_C | 0.54 ± 0.22 | 0.68 ± 0.16 | -0.32 ± 0.33 | -0.24 ± 0.28 | 0.92 ± 0.04 |
| DMI_C | 0.70 ± 0.17 | 0.64 ± 0.18 | 0.18 ± 0.33 | 0.33 ± 0.26 | 0.39 ± 0.22 |

¹On farm phenotypes: CH4_P = predicted methane production; CO2_P = predicted carbon dioxide production; CH4_C = methane breath concentration; CO2_C = carbon dioxide breath concentration; CH4/CO2_C = ratio of methane to carbon dioxide breath concentration; ECM_C = Energy corrected milk yield; LW_C = live weight; DMI_C = dry matter intake. ²Respiration chamber phenotypes: CH4_RC = methane production; CO2_RC = carbon dioxide production; CH4/CO2_RC = ratio of methane to carbon dioxide production; ECM_RC = energy corrected milk yield; LW_RC = live weight.

The predicted mass flux phenotype CH₄_P showed the highest r_1 of any on-farm phenotype with CH₄_RC 0.77 ± 0.18 which approaches the threshold of 0.80 for no significant re-ranking, however the standard errors thereof are large. This finding agrees with that of Hellwing et al (2013), who compared CH₄ production calculated from the ratio of CH₄ to CO₂ measured within the RC and predicted CO₂ production to CH₄ production measured in the RC for 157 cow measurements in 8 feeding experiments and found an R^2 of 55% which would correspond to a correlation of 0.74. The performance of CH₄_P is reliant on the accuracy of prediction of CO₂_P as a tracer gas, the r_1 between predicted mass flux (CO₂_P) and actual measured CO₂ production (CO₂_RC) was in a similar order of magnitude $r_1 = 0.79 \pm 0.14$ to that of CH₄_P and CH₄_RC. A criticism of this CO₂ prediction equation is that the metabolizable energy (ME) efficiency and mobilization of body tissues is not taken into account, which runs the risk of over predicting CO₂ production of efficient cows (increased LW and ECM at a fixed level of intake) and under predicting CO₂ production of inefficient cows (Madsen et al., 2010; Huhtanen et al., 2015). It may be possible to improve the r_1 between CH₄_P and CH₄_RC through improving the prediction accuracy of CO₂_RC by taking into account ME utilization. For instance Negussie et al (2016) compared CH₄_P predicted from the ratio of CH₄ to CO₂ concentration in the breath of 20 lactating Nordic Red cattle in concentrate feeders and CO₂_P predicted from metabolizable energy intake and found CH₄_P to have a high concordance correlation coefficient 0.70 and phenotypic correlation 0.80 with CH₄_RC.

Since all prediction equations have some level of inherent error and traits used in the prediction of CO₂_P, e.g. ECM_C, are already in the breeding goal, there is interest in assessing value of directly measured traits like CH₄_C and CO₂_C with RC traits. In this instance CH₄_C ranked animals comparatively well with CH₄_RC $r_1 = 0.75 \pm 0.20$ as compared to CH₄_P and CH₄_RC $r_1 = 0.77 \pm 0.18$ and exceeded that of both commercial control variables ECM_C, LW_C and DMI_C, which are routinely used to predict CH₄ production (Ramin and Huhtanen, 2013).

A number of authors have labelled breath gas concentration measures as imprecise (Huhtanen et al., 2015; Goopy et al., 2016; Wu et al., 2018) and this may be the case as seen by the lower repeatability estimates of CH₄_C and CO₂_C and the increased CV as compared to other phenotypes. However, the aforementioned studies often compare mass flux CH₄ production (g/day) to CH₄ breath concentrations using the coefficient of determination (R^2) with the expectation that a deviation of R^2 from unity (1.0) indicates imprecision. However, the R^2 metric reflects a combination of between cow variation and residual error (imprecision), recognizing that CH₄_C is a separate, by likely correlated trait deviations of R^2 from

1.0 are to be expected regardless of imprecision. In the present study, we made use of individual level correlations which have been corrected for non-genetic effects which can inflate between subject variability (i.e. parity, diet, lactation stage, breed) and thus phenotypic correlations. Furthermore, repeated measures per cow allows for partitioning and removal of residual variance (imprecision) which can bias phenotypic correlations downwards (Adolph and Hardin, 2007; Dingemanse and Dochtermann, 2013). Thus individual level correlations are one-step closer to genetic correlations as compared to phenotypic correlations and the coefficient of determination and are not attenuated by imprecision (Bakdash and Marusich, 2017).

3.4 Conclusions

Methane emission traits derived from breath gas measurements during milking correlated the highest with methane production in respiration chambers, exceeding that of live weight, energy corrected milk yield and dry matter intake. However, breath gas phenotypes did not exceed a correlation of 0.80 with CH₄ production in the RC and thus have potential to serve as large-scale indicator traits of CH₄ production. Whilst individual level correlations can serve as a proxy for genetic correlations, genetic correlations between RC phenotypes and breath gas phenotypes are still needed for effective use in genetic selection indices. Given the difficulties in acquiring suitably large numbers of cows in RC, the most feasible current way to achieve this is through international collaborations and incorporation of genomic information.

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Chapter 4

Can greenhouse gases in breath be used to genetically improve feed efficiency of dairy cows?

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Abstract

There is considerable interest in improving feed utilization of dairy cattle whilst limiting losses to the environment (i.e. greenhouse gases (**GHG**)). In order to breed for feed efficient or climate friendly cattle, it is first necessary to obtain accurate estimates of genetic parameters and correlations of feed intake, greenhouse gases and production traits. Reducing dry matter take (DMI) whilst maintaining production has high economic value to farmers, but DMI is costly to record and thus limited to small research or nucleus herds. Conversely, enteric methane (CH₄) currently has no economic value, and is also costly to record and is limited to small experimental trials. However, breath gas concentrations of methane (**CH4c**) and carbon dioxide (**CO2c**) are relatively cheap and have a high throughput under commercial conditions by installing sniffers into automated milking stations. The objective of this study was to combine DMI records from Denmark (**DNK**) and The Netherlands (**NLD**) and assess the genetic correlations with body weight (**BW**), fat and protein corrected milk yield (**FPCM**) and GHG related traits: CH4c and CO2c. Secondly, to assess the genetic potential for reducing DMI whilst maintaining FPCM and BW, and the added benefits of using CH4c and CO2c as indicators. Feed intake data was available on 812 primiparous cows in DNK and 2112 in NLD. Furthermore, CH4c and CO2c records were available on 478 primiparous cows in DNK and 660 in NLD. Genetic parameters were estimated using full trait animal models for DNK, NLD and combined from both countries. A restricted selection index for reducing DMI whilst maintaining FPCM and BW (genetic residual feed intake, **gRFI**) was heritable in and over both countries (DNK = 0.15, NLD = 0.17, Combined = 0.14), and genetically correlated to DMI (DNK = 0.21, NLD = 0.74, Combined = 0.64), but uncorrelated to FPCM and BW. The greenhouse gas related traits were heritable e.g. CH4c (DNK = 0.26, NLD = 0.16, Combined = 0.22), and moderately genetically correlated to DMI (DNK = 0.58, NLD = 0.29, Combined = 0.41). The relatively cheap GHG related traits showed considerable potential to improve the accuracy of breeding values of the highly valuable breeding goal traits feed intake and feed efficiency.

Key words: Breath concentration, Feed efficiency, Methane, Sniffers

4.1 Introduction

The dairy industry faces major challenges to remain profitable whilst maintaining environmental sustainability (for instance mitigating greenhouse gas emissions like methane (CH₄) and carbon dioxide (CO₂)). Feed plays a crucial role in economic and environmental performances of dairy production units, as feed constitutes the highest variable cost of production (Hemme et al., 2014). Feed composition affects not only CH₄ emissions at an individual cow level (Hristov et al., 2013), but the entire dairy production chain as well (Møller et al., 2014). The primary income from milk is volatile as milk prices often fail to increase with feed prices (Hemme et al., 2014). A promising way to increase profitability is to reduce feed costs through reducing feed intake and improving or maintaining production in the breeding objective (Veerkamp, 1998; de Haas et al., 2012). In other words, identifying cows that require less feed to give the same levels of production as contemporaries.

However, feed efficiency is not a directly observable trait, rather it is a combination of multiple traits including feed intake and production. Feed efficiency is thus defined as feed intake conditional on a combination of production traits or energy sinks (e.g. fat and protein correct milk (FPCM), body weight (BW), etc.) (Jensen et al., 1992). The most prevalent feed intake trait is dry matter intake (DMI), as this accounts for differences in moisture content of feeds and is the most prevalent across countries (Berry et al., 2014). The recording of DMI is labor intensive, expensive and limited to small research or nucleus herds, resultantly the reliability of sire breeding values are low (de Haas et al., 2012). This could in principle be overcome by including heritable and correlated predictor traits measured on a large scale in commercial and research herds in conjunction with accurate small scale research and contract herds with DMI recording (Berry and Crowley, 2013; Manzanilla-Pech et al., 2016; Wallén et al., 2017).

Identifying the cows that genetically require lower feed intake whilst giving the same levels of production as contemporaries is challenging, as DMI is highly genetically correlated to production traits like FPCM and BW (Manzanilla Pech et al., 2014). Resultantly, selection for decreased DMI is expected to have correlated reductions in FPCM and BW, which would simultaneously decrease feed costs and income from milk production. Feed efficiency traits like feed conversion ratio (FCE = FPCM/DMI) and residual feed intake (RFI), fare no better as both retain genetic correlations with production traits (Manafiazar et al., 2015). Following a restricted selection index approach (Kennedy et al., 1993) defined genetic residual feed intake (gRFI) as DMI genetically independent of FPCM and BW, which overcomes the

inherent deficiencies of ratio or phenotypic regression residual definitions of feed efficiency (Lu et al., 2015). However, the desirable statistical properties of gRFI can also hinder its uptake, since sufficiently large numbers of records on animals are required to accurately estimate the genetic correlation between DMI, FPCM and BW, which are likely only achievable through international collaborations (Berry et al., 2014). Furthermore, FPCM and BW can no longer serve as large-scale indicators of gRFI outside of their use as indicators of DMI, as they are genetically uncorrelated to gRFI. Given the difficulties in obtaining DMI records on large numbers of cows, there is a need for other easily recorded large scale traits genetically correlated to gRFI that are independent of FPCM and BW.

Studies using the intensive gold standard respiration chambers have shown high phenotypic correlations between mass fluxes of the greenhouse gas traits (GHG), CH₄ and CO₂ with DMI (Yan et al., 2010). Furthermore, the calorie value of CH₄ has been estimated as 2% - 12% of the gross energy intake of the cow (Johnson and Johnson, 1995). This has led to numerous prediction equations employing DMI and feed components to predict CH₄ production (Ellis et al., 2007; Ramin and Huhtanen, 2013), and genetic and genomic parameters for predicted CH₄ emission from DMI or metabolisable energy intake (de Haas et al., 2011; Negussie et al., 2014). Furthermore, predicted methane emission was found to be favorably genetically correlated to phenotypic RFI (de Haas et al., 2011). Thus turning this idea around, GHG traits like CH₄ production could potentially serve as indicator traits for DMI and gRFI. However, respiration chambers (RC) have high capital and labor costs and have mostly proven prohibitive to large scale genetic evaluations, with the largest studies using RC reaching approximately 1000 individuals in growing beef or sheep breeds (Pinares-Patiño et al., 2013; Donoghue et al., 2016). Furthermore, confinement within the RC is known to impact natural animal behavior and can cause a drop in DMI, leading many to question the extrapolation of these results to intensive commercial and grazing systems (Pinares-Patiño and Clark, 2008). Thus, the potential of GHG traits recorded using RC as indicator traits remains limited.

Recently, a cost effective, non-invasive and high through put method of recording methane (CH₄c) and carbon dioxide concentrations (CO₂c) in the breath of individual cattle under commercial conditions, whilst milking in the automated milking stations (AMS) (Garnsworthy et al., 2012a; Lassen et al., 2012) or concentrate feeders has emerged, collectively known as sniffers (Negussie et al., 2016). Since sniffers record gas concentrations and not mass fluxes, researchers have used scaling factors and calibrations (Garnsworthy et al., 2012b) or predicted CO₂ as a tracer gas to approximate CH₄ mass fluxes (Madsen et al., 2010). These approximations have been shown to be heritable, genetically correlated with FPCM and BW and have good

concordance with RC (Garnsworthy et al., 2012b; Lassen and Løvendahl, 2016; Negussie et al., 2016). The direct measures CH₄c and CO₂c have recently been reported as lowly heritable 0.11 and 0.12, respectively (van Engelen et al., 2018). However, the genetic correlations of CH₄c and CO₂c with feed intake and production traits are currently not known and their potential as indicator traits for DMI or gRFI is under-exploited.

The aims of this study were to estimate genetic parameters for feed intake related traits and greenhouse gas related traits in primiparous Holsteins in Denmark (DNK) and The Netherlands (NLD). Secondly, assess the use of breath concentration GHG traits as large-scale indicator traits for improving the accuracy of breeding values for DMI and gRFI.

4.2 Materials and Methods

4.2.1 Design and animals

Data on individual DMI, BW and FPCM as well as CH₄c and CO₂c was collected repeatedly over lactation in primiparous Danish Holstein and Dutch Holstein Friesian cows. Both countries have recorded DMI, BW and FPCM in indoor research herds over the last few decades, with numerous nutritional trials of various total mixed ration (TMR) feeds and expressed these as weekly averages for genetic evaluations. More recently, both countries have installed non-invasive sniffers in AMS at the research stations and in commercial herds for the repeated recording of CH₄c and CO₂c, expressed as weekly averages. The commercial herds were included to improve the representativeness of genetic parameters estimated and convergence of multitrait animal models as these herds had CH₄c, CO₂c, FPCM and BW. The data collection, editing and analysis is presented below.

4.2.2 Data Collection and Editing: Feed intake traits

Denmark. A total of 29,084 weekly averaged records for DMI, BW and milk yield (**MY**) were available on 812 primiparous Danish Holstein cows which calved from 1995 - 2015. Cows that calved between 1995 to 2000 were located at Ammitsbøl Skovgaard research herd (Skovgaard, Vejle, Denmark), from November 2001 until present the cows were relocated to the Danish Cattle Research Center (**DCRC**; Foulum, Denmark), as previously reported (Li et al., 2017). The Skovgaard barn was a tie stall system, with milking conducted twice daily and proportional samples drawn to estimate milk fat, protein and lactose content. Body weight measurements taken once a week until 3 months after calving at which point they were weighed fortnightly. Cows were fed manually 3 – 5 times a day to allow for 5%

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refusal and individual refusals recorded to determine feed intake, weekly samples of feed were taken to estimate dry matter content of the rations. The DCRC barn is a loose housing system where cows have free access to ad libitum feed in individual automated feed bins (RIC system, Insentec B.V., Marknesse, the Netherlands). Cows also had access to automatic milking stations (AMS; DeLaval AB, Tumba, Sweden) where they were offered up to 3kg of concentrate a day within the integral AMS feed bins. The AMS was fitted with a weighing platform (Danvaegt, Hinnerup, Denmark) to record BW at each milking and a device delivering and recording the amounts of concentrates and refusals.

Cows were entered into numerous nutritional trials at DCRC and have remained on a largely unchanged control TMR diet in between trials, which consists of rolled barley, corn silage, grass clover silage, rapeseed meal, and soybean meal. The DM contents of the TMR and concentrates were determined from regular analysis and combined with averaged weekly feed intake to obtain DMI per cow per week of lactation. Although multiple trials calculated, energy contents of specific feed treatments, energy intake was largely unavailable for the majority of cows. The diets within a specific treatment are expected to be the same, whilst dietary differences between treatment levels are expected. Furthermore, the dietary composition of the control diets can be expected to change with the influences of year and seasons on feed ingredients (Nielsen et al., 2003). Thus, an experiment by treatment interaction term was created which corresponds to the experimental trial and specific treatment in the case of nutritional experiments and a year by season variable for the control diet. This method of modelling the changes in ration is consistent with modelling longitudinal intake data in other internal genetic evaluations such as the global dry matter initiative (**gDMI**) (Berry et al., 2014).

When combining the data from Skovgaard where milking was conducted twice daily and DCRC where milking was voluntary in AMS, the 3 d moving average approach described by ICAR (2016) was employed. In this approach daily MY is obtained from, the sum of yield over 3 d divided by milking intervals and multiplied by 24 h. Milk composition was determined from every milking over a consecutive 48 h period each week using CombiFoss (Foss, Hillerød, Denmark) operated by Eurofins (Vejen, Denmark). Composition data from each milking were used to calculate yields of fat and protein which were smoothed by the moving average method to obtain daily yields. The calculated yields were averaged per week to get weekly records on the same time scale as DMI records. The DCRC records for BW, MY and DMI were filtered to remove weekly averages where more than two days records were missing. Cows which had less than 5 weekly records during their first lactation or an age of calving greater than 36 months were omitted. Although many of the experimental

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treatments were carefully blocked and balanced, these trials often included multiparous cows which were not the focus of the present study. Thus, a filter of a minimum of 5 cows per experimental treatment was added to ensure adequate degrees of freedom for effect estimation. The records after editing are summarized in Table 4.1 with datasets DNK1 and DNK2 corresponding to Skovgaard and DCRC, respectively.

The Netherlands. In total 47,356 weekly averaged records for DMI, BW and milk yield (MY) were included in analysis on 2112 primiparous Dutch Holstein cows which calved from 1986 - 2015. The Dutch data was comprised of multiple nutritional experiments conducted at multiple locations in The Netherlands (e.g. Aver Heino; Bosma Zathe, Ureterp; Cranendonck, Soerendonk; 't Gen, Lelystad; Minderhoudhoeve, Swifterbant; Waiboerhoeve Dairy unit 2, 3, Lelystad; Zegveld farm, Zegveld; Hoorn, Lelystad; New Waiboerhoeve, Lelystad; Dairy Campus, Lelystad). Data from these farms has been classified into four datasets (NLD1, NLD2, NLD3 and NLD4) based on previous studies which have described data collection (see Table 4.1) (Veerkamp et al., 2000; Beerda et al., 2007; Zom et al., 2012; van Knegsel et al., 2014). All herds made use of indoor housing and offered TMR that primarily consisted of grass silage, fresh grass, dehydrated grass, corn, corn silage, cereal, concentrates or beet pulp.

Differences in the recording method and frequencies exist between datasets and within datasets through time and will be mentioned briefly. Milking was conducted twice daily for most herds, with the exception of one herd (n=50) where cows were milked three times a day and the AMS herds where milking frequency is based on voluntary milking.

Initially, dry matter intake was recorded manually at 1, 2, 3 or 5 times per week depending on the experiment (Veerkamp et al., 2000; Beerda et al., 2007; Zom et al., 2012). Later, feed intake was recorded using individual automated feed bins (RIC system, Insentec B.V., Marknesse, the Netherlands). Feed was offered to achieve a 10% refusal rate, with offered feed recorded automatically but refusals recorded manually (Zom et al., 2012). In general, BW was recorded either daily, three times a week using automatic weighing platforms or in the AMS integral automatic weighing platforms. Cows were entered into numerous nutritional trials, where the diets within treatments are expected to be the same but differ between trial treatments. The DM composition of diets was used to calculate DMI, but the specific dietary composition and energy of many diets were not available. Fat and protein contents of milk were estimated at minimum once per week in all datasets. Data editing steps included only retaining cows with a Holstein breed percentage of 75%

or greater, minimum of 5 observations per experimental treatment and 5 records per cow and calved no later than 36 months of age.

Fat and protein corrected milk yield was calculated for both Denmark and The Netherlands using a single standardization equation of 4.0% fat and 3.3% protein as recommended in (FAO, 2010) using the FPCM (kg) = $(0.337 * \text{milk yield kg}) + (11.6 * \text{fat yield kg}) + (5.999 * \text{protein yield})$ (CVB, 2008).

4.2.2 Data Collection and Editing: Greenhouse gas traits

Denmark. Data on the CH₄c and CO₂c of individual Holsteins cows during milking was available on 267 primiparous cows at DCRC from mid 2013 – November 2016, with simultaneous DMI, FPCM and BW records. The non-dispersive infrared (NDIR) CH₄ sensor (Guardian NG, Edinburgh Instruments Ltd, Livingston, UK) and the NDIR CO₂ sensor (Gascard, Edinburgh Instruments Ltd, Livingston, UK) were installed in each of the AMS. The equipment installation, technical specifications, and calibration procedures for the sensors are described elsewhere (Difford et al., 2016). The time stamped AMS visit data and breath gas concentration time series on a 1 second basis were merged using a time alignment algorithm (Difford et al., 2016). Gas data from AMS milkings shorter than 300s and marked as incomplete or interrupted by the AMS were removed. A cow head-lifting algorithm was used to filter out gas readings when the cow's head is predicted to be outside of the feed bin (Difford et al., 2016). The ambient barn concentrations of CH₄c and CO₂c were estimated once daily during the morning cleaning cycle of the AMS when the AMS is free of cows and subtracted from the mean CH₄c and CO₂c of each milking for that day.

Greenhouse gas measurements were also taken in two commercial herds in Denmark from November 2015 until March 2016 as part of the REMRUM project. A portable set of four NDIR sensors as above and one portable Fourier transform infrared (FTIR) Gasmeter DX-4000 (Gasmeter; Gasmeter Technologies Oy, Helsinki, Finland), which registers gas concentrations on a 5 s basis. The use of the FTIR equipment in commercial herds is described elsewhere (Lassen and Løvendahl, 2016). The data from commercial herds was handled the same as the research herd described above. Sensors were installed in AMS (Astronaut A3, Lely Industries, NV, Maassluis, the Netherlands) equipped with weighing platforms. The milk components were taken from the national recording scheme (RYK, Skejby, Denmark) and cow fat and protein corrected milk yield (FPCM) (CVB, 2008; FAO, 2010) was estimated as in the Dutch population.

Data from the research and commercial herds were filtered to only include weekly averages where a maximum of 3 days were missing within each week of

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measurement and each cow had a minimum of 3 weekly measurements (Table 4.1). The effect of contemporary group and AMS are by necessity confounded at DCRC and are thus also confounded with the experiment by treatment interaction described for the feed intake traits. The commercial herds maintained the same diet whilst gases were recorded; a variable of year by season by AMS group nested within herd was used in place of the experiment by treatment term used in the research herds. Similarly, gas traits were filtered to ensure each experiment by treatment interaction term had at least 5 cows and a minimum of 5 weekly records per cow. The dataset from DCRC corresponding to DNK2 and commercial herds corresponding to DNK3 (Table 4.1).

The Netherlands. Data on the CH₄c and CO₂c breath gas concentrations of individual Holstein cows during milking was available on 277 primiparous cows at Dairy Campus Lelystad (NLD4) from October 2013 – November 2016, with simultaneous FPCM and BW records. The grouping of animals at NLD4 is such that the two AMS groups do not have automatic feed bins. Thus cows are tested for 70d for DMI and then moved to the AMS groups where gas recording takes place, no simultaneous records for DMI and gases are available in The Netherlands. The FTIR sensor inlet was installed in the integral feed bins of two AMS (Astronaut A3, Lely Industries, NV, Maassluis, the Netherlands), where gases were drawn from each of the AMS on alternating days. The time stamped AMS visit data and breath gas concentration time series on a 5 second basis were merged. The air sampled for the first 20 seconds of each visit was discarded to ensure no carryover of gas readings from the previous cow's milking. Data from milkings that lasted shorter than 120 seconds and data collected after 600 seconds of the start of the milking were excluded. As well as data from milkings marked as error by the AMS were excluded. The ambient barn concentrations of CH₄c and CO₂c were estimated once daily during the morning cleaning cycle of the AMS when the AMS is free of cows and subtracted from the mean CH₄c and CO₂c of each milking for that day. After editing 5904 weekly average records were available on 277 primiparous Holstein cows corresponding to dataset NLD4 (Table 4.1).

Data on the CH₄c and CO₂c on 382 primiparous Dutch Holstein Friesian cows from 11 commercial herds over November 2013 – March 2016 as part of the TiFN project 'Reduced methane emission of dairy cows' (NLD5) (Table 4.1). Four portable NDIR sensors (SenseAir LPL 113 CH₄/CO₂, Rise Acreo, Stockholm, Sweden) were installed in the integral feed bin of AMS (Lely, Astronaut A4, Lely Industries NV, Maassluis, the Netherlands) which record CH₄c and CO₂c twice per second. The sampling strategies, installation and phenotype calculation has been described elsewhere (van Engelen et al., 2018). Briefly, the time series data was aligned to the

AMS data by a function which maximizes the CH₄c and CO₂c during visits and minimizes CH₄c and CO₂c between visits. Visits which failed to exceed 90 seconds in length were excluded. The background concentrations for CH₄c was assumed to be 0 ppm and CO₂c 400ppm, these values were subtracted from the mean of 10 minimum recordings per visit and subtracted from the mean of the visit to approximate a background corrected CH₄c and CO₂c mean per milking. The CH₄c and CO₂c per milking were combined with the MY per milking. Where possible the mean BW per milking was retained, however the majority of AMS were not equipped with integral weighing platforms so BW was not available from all locations. Records on milk components during the sampling periods for each herd were acquired from the cooperative cattle improvement organization CRV (Arnhem, the Netherlands) to estimate FPCM.

Data from the research and commercial herds were filtered to only include weekly averages where a maximum of 3 days were missing within each week of measurement and each cow had a minimum of 3 weekly measurements (Table 4.1). The effect of contemporary group and AMS were by necessity confounded at Dairy Campus and are thus confounded with the experiment by treatment interaction described for the feed intake traits. The commercial herds maintained the same diet whilst gases were recorded; a variable of year by season by AMS group nested within herd was used in place of the experiment by treatment term used in the research herds. Similarly, gas traits were filtered to ensure each experiment by treatment interaction term had at least 5 cows and a minimum of 5 weekly records per cow (Table 4.1).

4.2.3 Data Collection and Editing: Pedigrees

Pedigree information of all animals was traced back to founder generations, by extracting the Danish and Dutch pedigrees from the Nordic Cattle Genetic Evaluation Database (NAV, Skejby, Denmark) and the cooperative cattle improvement organization CRV (Arnhem, the Netherlands). Pedigrees were merged and all aliases in the pedigree were removed through the use of the Interbull identification cross-reference tables and manual curation of the pedigree. Non-informative individuals in the pedigrees were removed with the DMU Trace program (Madsen, 2012). The total pedigree file consisted of 16,592 animals with a DNK pedigree of 9774 animals and NLD pedigree of 8042 animals. A total of 1074 sires were in common between the two pedigrees.

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Table 4.1. Description of available phenotypic records of dry matter intake (DMI), fat and protein corrected milk (FPCM), body weight (BW), methane breath concentration (CH4c) and carbon dioxide breath concentration (CO2c) for first parity Holstein Friesian cows in Denmark and The Netherlands.

| | | Denmark | | | The Netherlands | | | | |
|--|-----------|--------------------|---------------|-------------|-------------------|------------------|-------------|----------------|------------------------|
| Data set | | DNK1 | DNK2 | DNK3 | NLD1 | NLD2 | NLD3 | NLD4 | NLD5 |
| # of experiments | # | 2 | 24 | 1 | 13 | 1 | 3 | 61 | 1 |
| # of rations | | 2 | 112 | 2 | 47 | 27 | 27 | 268 | 10 |
| # locations | | 1 | 1 | 2 | 1 | 1 | 2 | 9 | 10 |
| # of cows with¹: | | | | | | | | | |
| DMI | | 109 | 703 | 0 | 195 | 491 | 393 | 1,033 | 0 |
| FPCM | | 108 | 702 | 267 | 193 | 474 | 362 | 1,183 | 362 |
| BW | | 111 | 702 | 267 | 195 | 490 | 361 | 1,217 | 101 |
| CH4c | | 0 | 211 | 267 | 0 | 0 | 0 | 277 | 382 |
| CO2c | | 0 | 211 | 267 | 0 | 0 | 0 | 277 | 382 |
| Total | | 111 | 703 | 267 | 195 | 491 | 393 | 1,220 | 382 |
| # weekly records for¹: | | | | | | | | | |
| DMI | | 4,090 | 24,994 | 0 | 4,105 | 11,086 | 9,270 | 15,731 | 0 |
| FPCM | | 3,472 | 23,853 | 790 | 2,503 | 10,476 | 9,141 | 19,679 | 668 |
| BW | | 2,813 | 24,812 | 800 | 3,925 | 10,460 | 8,875 | 21,935 | 679 |
| CH4c | | 0 | 5,212 | 800 | 0 | 0 | 0 | 5,904 | 2,496 |
| CO2c | | 0 | 5,212 | 800 | 0 | 0 | 0 | 5,904 | 2,496 |
| Total | | 4,355 | 25,371 | 800 | 4,105 | 11,086 | 9,270 | 22,895 | 2,496 |
| Weeks in lactation | in | 1 – 44 | 1 - 44 | 1 - 44 | 1 - 44 | 1 - 44 | 1 - 44 | 1 - 44 | 1 - 44 |
| Gas recording period | | N/A | 2013-2016 | 2014 - 2016 | N/A | N/A | N/A | 2013-2016 | 2014-2016 |
| DMI recording | | 1997-2001 | 2002-2016 | N/A | 1991 - 2004 | 1991 - 2005 | 1991 - 1998 | 1990-2016 | N/A |
| Published | | Nielsen et al 2003 | Li et al 2016 | Unpublished | Beerda et al 2007 | Veer et al. 2000 | Unpublished | Zom et al 2012 | Van Engelen et al 2018 |

¹DMI = Dry matter intake; FPCM = fat and protein corrected milk yield; BW = body weight; CH4c = methane breath concentration and CO2c = carbon dioxide breath concentration.

4.2.4 Statistical Analysis

Cows had multiple records for CH₄c and CO₂c per day made at different times per day. The starting time for each visit in the AMS was converted to 24 h angular radians for further modelling of diurnal variation in a Fourier series approach (Lassen et al., 2012). A pre-correction model within AMS group for all herds in all countries was used to estimate the daily breath gas concentration free of the effects of diurnal variation, sensor drift after calibration and test day variation between calibrations:

$$y_{ijklm} = \mu + d_i + b_j + \sum_{k=1}^3 (f_{1k} \sin \vartheta + f_{2k} \cos \vartheta) + c_l + e_{ijklm} \quad (1)$$

where y_{ijklm} is the natural logarithm of background corrected AMS visit means of CH₄c and CO₂c; d_i is the fixed effect of test day i ; b_j is the fixed effect of the first full day after each calibration j ; f_{1k} and f_{2k} are the fixed regression coefficients of Fourier series linear covariates of the time of day of measurement, modelled as harmonic pairs. Here ϑ denotes the time of day of visit expressed as 24 h angular radians. Term c_l is the random effect parameters for each cow $c_l \sim ND(0, \sigma_c^2)$, and e_{ijklm} is the random residuals $\sim ND(0, \sigma_e^2)$. In order to correct daily AMS visit means, the random residuals for each visit were summed with random cow solutions, intercept, calibration day, and the regression coefficients f_{1k} and f_{2k} multiplied by the angular radian corresponding to 12:00:00 a.m. Further, the CH₄c and the CO₂c in natural log transformed parts per million as the average per week of lactation was obtained by taking the corrected means of visits weighted by the number of seconds for each visit where the cow head was positioned correctly in the feed bin.

Univariate repeatability mixed models were conducted by use of Proc Mixed in SAS (SAS 9.3, SAS Institute Inc.) where significance of fixed effects was tested using Kenward-Roger correction for the correct denominator degrees of freedom, for all traits within country. The residuals were tested for deviations from normality by means of Kolmogorov–Smirnov normality tests and severe violations were detected for CH₄c and CO₂c which were best remedied by a natural log transformation. The means and variances of CH₄c and CO₂c from different measuring equipment and AMS models (DeLaval and Lely A3 and A4 Astronauts) differed significantly. Based on the results of a pilot study combining data from multiple measuring equipment it was decided to standardize CH₄c and CO₂c within equipment and country to a mean of zero and standard deviation of one (Difford et al., 2016).

A five trait multi-trait analysis was performed to estimate the variance components and the heritability within each country and in a combined country dataset i.e., DMI, FPCM, BW, CH₄c and CO₂c. All analyses were performed using

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DMU version 6 (Madsen and Jensen, 2014). The model used in the Danish dataset was as follows:

$$y_{ijklm} = \mu + W_i + ETC_j + a_l + pe_l + e_{ijklm} \quad (2)$$

Where y_{ijklm} is the trait of interest (DMI, BW, FPCM, CH4c or CO2c), μ is the intercept, W is the i^{th} week of lactation, ETC is the j^{th} experiment by treatment contemporary group effect, a_l is the random additive effect of the l^{th} animal $N(0, A\sigma_a^2)$ where A is the pedigree derived numerator relationship matrix, pe is the permanent environmental effect distributed following $N(0, I\sigma_{pe}^2)$ and e is the random residual $N(0, I\sigma_e^2)$. For the feed intake traits (DMI, BW and FPCM) the linear covariate *age at first calving* was significant and was included in model 2 as a fixed regression.

For the Dutch data it was necessary to place added restrictions on the permanent environmental covariance and residual covariance between CH4c, CO2c and DMI as these were recorded at different time points such that they were undefined:

$$\begin{aligned} \text{Var} \begin{bmatrix} \mathbf{a}_{DMI} \\ \mathbf{a}_{GHG} \end{bmatrix} &= \begin{bmatrix} \mathbf{A}\sigma_{a_{DMI}}^2 & \mathbf{A}\sigma_{a_{DMI}a_{GHG}} \\ \mathbf{A}\sigma_{a_{GHG}a_{DMI}} & \mathbf{A}\sigma_{a_{GHG}}^2 \end{bmatrix}, \\ \text{Var}(pe) = \mathbf{PE} &= \begin{bmatrix} \sigma_{pe_{FI}}^2 & \mathbf{0} \\ \mathbf{0} & \sigma_{pe_{GHG}}^2 \end{bmatrix}, \\ \text{Var}(e) = \mathbf{R} &= \begin{bmatrix} \sigma_{e_{FI}}^2 & \mathbf{0} \\ \mathbf{0} & \sigma_{e_{GHG}}^2 \end{bmatrix}, \end{aligned}$$

For the combined dataset the models described in (2) above were then run without placing any restrictions on the covariances between CH4c, CO2c and DMI.

Feed efficiency. The 5 x 5 variance covariance matrices estimated from the Danish, Dutch and combined country dataset were used to estimate genetic residual feed intake (gRFI) as per (Kennedy et al., 1993; Shirali et al., 2017). In which $\widehat{\mathbf{G}}$ and $\widehat{\mathbf{P}}$ are 5 x 5 additive genetic and phenotypic variance covariance matrices, respectively, for feed intake traits (FI), i.e. DMI, FPCM, BW, green house gas traits (GHG) i.e. CH4c and CO2c. Firstly, the multivariate genetic partial regression coefficients \hat{b}_1 and \hat{b}_2 of FPCM and BW on DMI are estimated respectively. The partial genetic regression coefficients were computed using genetic variance covariance matrices as $\hat{b}_{DMI|FPCM, BW} = \widehat{\mathbf{G}}_{DMI|FPCM, BW} \widehat{\mathbf{G}}_{FPCM, BW}^{-1}$ in which $\widehat{\mathbf{G}}_{FPCM, BW}$ is a 2 x 2 genetic variance covariance matrix of FPCM and BW, $\widehat{\mathbf{G}}_{DMI|FPCM, BW}$ is a 1 x 2 covariance vector of FPCM and BW on DMI, resulting in a 1 x 2 vector of genetic partial regression coefficients.

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The genetic RFI which has DMI genetically orthogonal of FPCM and BW estimated using genetic partial regression coefficients computed above in the following covariance function:

$$B = \begin{bmatrix} I_{DMI} & 0 & 0 & 0 & 0 \\ 0 & I_{FPCM} & 0 & 0 & 0 \\ 0 & 0 & I_{BW} & 0 & 0 \\ 0 & 0 & 0 & I_{CH4c} & 0 \\ 0 & 0 & 0 & 0 & I_{CO2c} \\ I_{DMI} & -\hat{b}_1 & -\hat{b}_2 & 0 & 0 \end{bmatrix}$$

Thus B is a 6 x 5 identify matrix which is pre- and post multiplied on the 5 x 5 \hat{G} and \hat{P} matrices (i.e. $B\hat{G}B'$) to obtain a 6 x 6 genetic and phenotypic variance covariance matrices including the newly computed gRFI which has the favorable properties of being genetically uncorrelated to FPCM and BW i.e. $\hat{G}_{RFI|FPCM,BW} = \hat{G}_{DMI|FPCM,BW} - \hat{b}_{DMI|FPCM,BW}\hat{G}_{FPCM,BW} = 0$.

4.2.5 Selection index calculations

Selection index theory was used to evaluate different recording strategies with FI and GHG trait combinations (I) on the selection goals (H) defined by DMI or RFI for bulls with daughter records by calculating selection accuracy ($r_{H,I}$) (Falconer and Mackay, 1996):

$$r_{H,I} = \sqrt{\frac{\mathbf{b}'\mathbf{P}\mathbf{b}}{\mathbf{v}'\mathbf{G}\mathbf{v}}}$$

Where, $r_{H,I}$ is the accuracy of the index, \mathbf{b} is the 6 x 1 vector of weighing factors for each of the records used in making selection decisions. \mathbf{P} is the 6 x 6 matrix of phenotypic (co)variance among records of each trait; \mathbf{v} is the 6 x 1 vector of relative economic values and G is the genetic (co)variance matrix among traits. To focus on recording strategies for RFI or DMI, we arbitrarily assigned a value of one to the respective selection goal and zero to all other traits in \mathbf{v} . As there is a large disparity in the throughput of different traits, combinations were evaluated assuming bulls had 100 daughters with single average weekly records of CH4c, CO2c, BW and ECM (a realistic assumption in AMS herds) and 10 daughters with single weekly average DMI records (a realistic assumption with research and nucleus herds) using the excel macro of (van der Werf, 2017).

4.3 Results

4.3.1 Descriptive statistics

Descriptive statistics of FI related and GHG related traits for DNK and NLD are shown in Table 4.2. The averages for FI related traits were marginally higher for DNK than for NLD with DMI 19.5 kg/d in DNK and 18.8 kg/d in NLD, FPCM was 28.9 kg/d in DNK and 28.3 kg/d in NLD and BW was larger 604.2 kg in DNK than in NLD 556.6. The GHG traits were likewise comparable between the countries, with CH₄c higher in NLD and 5.9 ln(ppm/d) than DNK 5.6 ln(ppm/d). Whilst, CO₂c was higher in DNK 8.5 ln(ppm/d) compared with 8.3 ln(ppm/d) in NLD.

4.3.2 Genetic parameters

The estimated genetic variances, heritability and repeatability for Danish, Dutch and the combined datasets are shown in Table 4.3. In general, all estimates were higher for traits in DNK as compared to NLD and the combined country dataset, FI related traits had higher heritability and repeatability than GHG related traits.

Feed intake related traits. Estimated heritability for DMI was 0.41 in DNK, 0.21 in NLD and 0.29 in the combined datasets with SE ranging from 0.03 to 0.06. The repeatability of DMI was more consistent with 0.64 in DNK and 0.59 in NLD and the combined dataset. Heritability estimates for FPCM followed a similar pattern as DMI across datasets with the highest estimate in DNK of 0.48, lowest in NLD 0.20 and intermediate in the combined dataset 0.35. Body weight was the most heritable trait in all datasets with 0.57 in DNK, 0.49 in NLD and 0.45 in the combined dataset.

Greenhouse gas related traits. The h^2 estimates of CH₄c and CO₂c were similar within and across datasets. The heritability of CH₄c was highest in DNK 0.26, followed by the combined dataset 0.22 and NLD 0.16. Similarly, heritability of CO₂c was 0.23 in DNK and the combined dataset and 0.15 in NLD. The SE of h^2 was reduced in the combined country data (0.06) such that the 95% confidence intervals for h^2 of CH₄c and CO₂c did not include zero. Both weekly average CH₄c and CO₂c were highly repeatable in DNK 0.72 and 0.73, in NLD 0.67 and 0.61 and in the combined dataset 0.69 and 0.70, respectively.

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Table 4.2. Descriptive statistics (No. of weekly records, mean value, standard deviation (sd), minimum (min), maximum (max)) of feed intake related traits and greenhouse gas related traits for each country.

| Trait ¹ | Unit | Denmark | | | | | The Netherlands | | | | |
|--------------------|---------|--------------------|-------|------|-------|-------|--------------------|-------|------|-------|-------|
| | | No. weekly records | mean | sd | min | max | No. weekly records | mean | sd | min | max |
| DMI | kg/d | 24,994 | 19.5 | 3.1 | 6.6 | 37.2 | 39,464 | 18.8 | 3.2 | 5.0 | 31.9 |
| FPCM | kg/d | 23,906 | 28.9 | 5.8 | 3.0 | 53.0 | 41,454 | 28.3 | 5.0 | 2.9 | 54.8 |
| BW | kg | 24,908 | 604.2 | 64.9 | 387.0 | 869.0 | 44,649 | 556.6 | 56.9 | 321.8 | 851.6 |
| CH4c | Ln(ppm) | 4888 | 5.6 | 0.49 | 3.4 | 7.1 | 8400 | 5.9 | 0.36 | 3.2 | 7.4 |
| CO2c | Ln(ppm) | 4888 | 8.5 | 0.30 | 7.4 | 9.68 | 8400 | 8.3 | 0.42 | 6.5 | 9.3 |

¹DMI = Dry matter intake; FPCM = fat and protein corrected milk yield; BW = body weight; CH4c = natural logarithm of methane breath concentration and CO2c = natural logarithm of carbon dioxide breath concentration.

Genetic Residual Feed Intake. For gRFI, the genetic partial regression coefficients of FPCM and BW on DMI differed between the countries with 0.403 and 0.03 in NLD and 0.211 and 0.016 for DNK, respectively. However the relative ratio of regression coefficients was similar across countries. Thus the h^2 of gRFI was similar in DNK ($h^2 = 0.15 \pm 0.04$) to NLD ($h^2 = 0.17 \pm 0.04$) (Table 4.3). In all datasets the heritability of gRFI is lower than that of DMI and significant from zero. The estimates of G and P from the combined country increased certainty around the resulting regression coefficients of 0.295 and 0.017 for FPCM and BW on DMI respectively and a low but significant heritability ($h^2 = 0.14 \pm 0.03$) (Table 4.3).

4.3.3 Genetic and phenotypic correlations

Genetic and phenotypic correlations between FI related traits and GHG related traits within DNK, NLD and the combined dataset are presented in Table 4.4. In general, the genetic correlations between FI traits and GHG traits were similar in all datasets with the exception of gRFI as is discussed below. Furthermore, the standard errors of genetic correlations were far larger in DNK than in NLD and substantially reduced in the combined country dataset, following a trend of increases in numbers of cows.

Feed Intake related traits. DMI was highly genetically and phenotypically correlated with FPCM and moderately with BW. With genetic correlations for FPCM in DNK ($r_g = 0.79 \pm 0.05$), NLD ($r_g = 0.68 \pm 0.05$) and combined ($r_g = 0.67 \pm 0.05$). Genetic correlations between DMI and BW were somewhat different across datasets with (0.09 ± 0.16) in DNK and (0.38 ± 0.07) and combined (0.39 ± 0.07), although the magnitude of standard errors is limiting inferences into differences. Genetic correlations between FPCM and BW ranged from negative (-0.51 ± 0.17) in DNK close to zero in NLD and in the combined dataset (-0.05 ± 0.08). Whilst phenotypic correlations between FPCM and BW were very similar ranging from (-0.02 ± 0.05) in DNK to (0.08 ± 0.02) in NLD.

Greenhouse Gas related traits. The genetic correlations between GHG traits and FI traits were very similar for both countries. DMI was moderately genetically correlated to CH₄c, ranging from 0.29 in NLD to 0.58 in DNK, similarly for CO₂c estimates ranging from 0.34 in the combined dataset to 0.42 in NLD. The GHG traits also showed similar patterns for genetic correlations with FPCM and BW in both countries, where genetic correlations were positive with FPCM ranging from (0.24 – 0.54) and low with BW ranging from (-0.13 – 0.00). The genetic and phenotypic correlations between CH₄c and CO₂c were consistently very high with r_g ranging from 0.95 in DNK and NLD to 0.96 in the combined dataset.

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Table 4.3. Estimated genetic (σ^2_a) and phenotypic (σ^2_p) variances, heritability (h^2) and repeatability of weekly averages (t^2) with corresponding standard errors (se) in parenthesis from within country and combined country multitrait models.

| Trait ¹ | Denmark | | | The Netherlands | | | Combined Country | | |
|--------------------|--------------|------------------------|------------------------|-----------------|------------------------|------------------------|------------------|------------------------|------------------------|
| | σ^2_a | h^2 (se) | t^2 (se) | σ^2_a | h^2 (se) | t^2 (se) | σ^2_a | h^2 (se) | t^2 (se) |
| DMI | 2.38 | 0.41 _(0.06) | 0.64 _(0.02) | 1.08 | 0.21 _(0.04) | 0.59 _(0.01) | 1.63 | 0.29 _(0.03) | 0.59 _(0.01) |
| FPCM | 18.87 | 0.48 _(0.07) | 0.82 _(0.01) | 3.73 | 0.20 _(0.02) | 0.76 _(0.01) | 8.57 | 0.35 _(0.01) | 0.73 _(0.01) |
| BW | 1220.0 | 0.57 _(0.13) | 0.89 _(0.01) | 107 | 0.49 _(0.00) | 0.86 _(0.00) | 1162.7 | 0.45 _(0.00) | 0.86 _(0.01) |
| CH4c | 0.25 | 0.26 _(0.11) | 0.72 _(0.02) | 0.10 | 0.16 _(0.16) | 0.67 _(0.01) | 0.24 | 0.22 _(0.06) | 0.69 _(0.01) |
| CO2c | 0.23 | 0.23 _(0.12) | 0.73 _(0.02) | 0.11 | 0.15 _(0.15) | 0.61 _(0.02) | 0.23 | 0.23 _(0.06) | 0.70 _(0.01) |
| gRFI | 0.86 | 0.15 _(0.04) | 0.72 _(0.02) | 0.60 | 0.17 _(0.03) | 0.43 _(0.01) | 0.53 | 0.14 _(0.03) | 0.48 _(0.01) |

¹DMI = Dry matter intake; FPCM = fat and protein corrected milk yield; BW = body weight; CH4c = natural logarithm of methane breath concentration and CO2c = natural logarithm of carbon dioxide breath concentration.

Genetic Residual Feed Intake. Within and across countries gRFI was genetically uncorrelated to FPCM and BW (Table 4.4). However, in DNK the genetic correlation between DMI and gRFI was far smaller ($r_g = 0.21 \pm 0.61$) than in The Netherlands ($r_g = 0.74 \pm 0.06$). The genetic correlations between gRFI and GHG traits were negative in Denmark with ($r_g = -0.17 \pm 0.85$) and ($r_g = -0.22 \pm 0.95$) between CH4c and CO2c with gRFI respectively. Similarly, in The Netherlands ($r_g = -0.52 \pm 0.32$) and ($r_g = -0.44 \pm 0.32$) for CH4c and CO2c with gRFI respectively. In the combined country dataset gRFI remained genetically uncorrelated to FPCM ($r_g = 0.05 \pm 0.09$) and BW ($r_g = 0.00 \pm 0.09$). Whilst a strong positive genetic correlation with a substantially reduced standard error persisted with DMI ($r_g = 0.64 \pm 0.04$) (Table 4.4). In the combined dataset the genetic correlations between both CH4c and CO2c with gRFI were positive ($r_g = 0.33 \pm 0.15$) and ($r_g = 0.32 \pm 0.15$).

4.3.3 Prediction accuracy of selection indices for DMI and gRFI

The accuracy of bull selection goals for DMI or gRFI based on the combined country estimated G and P (co)variance matrices for several recording strategies for daughters using different combinations of FI and GHG traits are presented in Table 4.5. For both DMI and gRFI indices, including DMI daughter records gave higher accuracies than indices including only GHG or FPCM and BW. For DMI, accuracy with GHG traits was lower (0.43) than FPCM and BW (0.75) although, the accuracy of FPCM, BW and GHG traits (0.78) was comparable to the same index including DMI (0.83). Conversely, for accuracy of gRFI the contribution of FPCM and BW was negligible (0.05) as compared to the modest contributions of GHG traits (0.30). Furthermore, the accuracy of gRFI based on DMI and GHG traits (0.47) exceeded the accuracy of DMI alone (0.42) and was equivalent to DMI, FPCM and BW (0.47). The highest accuracy for gRFI (0.53) was achieved using all recoding traits including DMI.

Table 4.4. Genetic (below diagonal) and phenotypic (above diagonal) correlations between feed intake-related traits¹, greenhouse gas related traits² and feed efficiency restricted indices³, with corresponding standard errors in parenthesis for Denmark, The Netherlands and the combined dataset.

| Traits | Denmark | | | | | |
|------------------|-------------|--------------|--------------|--------------|--------------|--------------|
| | DMI | FPCM | BW | gRFI | CH4c | CO2c |
| DMI | | 0.56 (0.02) | 0.30 (0.05) | 0.20 (0.03) | 0.35 (0.05) | 0.23 (0.05) |
| FPCM | 0.79 (0.05) | | -0.02 (0.05) | -0.47 (0.02) | 0.31 (0.05) | 0.17 (0.05) |
| BW | 0.09 (0.16) | -0.51 (0.17) | | -0.41 (0.02) | 0.12 (0.06) | 0.09 (0.06) |
| gRFI | 0.21 (0.64) | 0.002 (0.58) | 0.004 (0.71) | | -0.06 (0.04) | 0.04 (0.04) |
| CH4c | 0.58 (0.19) | 0.54 (0.20) | 0.00 (0.26) | -0.17 (0.85) | | 0.91 (0.00) |
| CO2c | 0.36 (0.23) | 0.37 (0.20) | -0.02 (0.28) | -0.22 (0.95) | 0.95 (0.04) | |
| The Netherlands | | | | | | |
| DMI | | 0.47 (0.01) | 0.31 (0.02) | 0.85 (0.00) | -0.02 (0.08) | -0.01 (0.08) |
| FPCM | 0.68 (0.05) | | 0.08 (0.02) | 0.06 (0.02) | 0.11 (0.03) | 0.11 (0.03) |
| BW | 0.38 (0.07) | -0.05 (0.08) | | -0.06 (0.02) | 0.06 (0.04) | 0.08 (0.03) |
| gRFI | 0.74 (0.06) | 0.00 (0.15) | 0.00 (0.11) | | -0.09 (0.09) | -0.10 (0.09) |
| CH4c | 0.29 (0.14) | 0.24 (0.13) | -0.13 (0.14) | -0.52 (0.32) | | 0.87 (0.00) |
| CO2c | 0.42 (0.14) | 0.35 (0.12) | -0.12 (0.13) | -0.44 (0.32) | 0.95 (0.02) | |
| Combined Dataset | | | | | | |
| DMI | | 0.59 (0.01) | 0.29 (0.02) | 0.72 (0.01) | 0.29 (0.03) | 0.22 (0.03) |
| FPCM | 0.67 (0.05) | | 0.05 (0.02) | -0.11 (0.01) | 0.20 (0.03) | 0.14 (0.03) |
| BW | 0.39 (0.07) | -0.05 (0.08) | | -0.07 (0.02) | 0.03 (0.04) | 0.03 (0.04) |
| gRFI | 0.64 (0.04) | 0.05 (0.09) | 0.00 (0.09) | | 0.20 (0.03) | 0.16 (0.04) |
| CH4c | 0.41 (0.12) | 0.35 (0.11) | -0.06 (0.11) | 0.33 (0.15) | | 0.92 (0.00) |
| CO2c | 0.34 (0.12) | 0.27 (0.11) | -0.08 (0.11) | 0.32 (0.15) | 0.96 (0.00) | |

¹Feed intake-related traits: DMI = dry matter intake; fat and protein corrected milk yield = FPCM; and body weight = BW. ²Greenhouse gas related traits: natural logarithm of methane breath concentration = CH4c; and natural logarithm of carbon dioxide breath concentration = CO2c. ³Feed efficiency restricted indices: genetic residual feed intake = gRFI.

4.4 Discussion

4.4.1 Genetic parameters

Feed intake related traits. Investigations into the FI related traits DMI, FPCM and BW; have been an area of intense research in both countries with both being members of international collaborative consortia like gDMI (Berry et al., 2014). The heritability estimate for DMI (0.41) was slightly higher than reported previously for Nordic Holsteins (0.20 – 0.40) by Li et al. (2016). However, Li et al. (2016) estimated h^2 for 1 – 24 weeks of lactation with an increasing trend from 0.20 to 0.40, whilst we reported entire lactation h^2 from weeks 1-44. Similarly, the h^2 estimate for DMI in NLD was within the lower range of reported values from previous studies (0.15 – 0.41) (Koenen and Veerkamp, 1998; Vallimont et al., 2010). In the present study, parameters are estimated over the course of the entire first lactation, recognizing that previous findings from both populations demonstrate that the genetic parameters of DMI, FPCM and BW change over the course of lactation (Manzanilla Pech et al., 2014; Li et al., 2016). The genetic correlations between both countries have been estimated as close to unity (0.80 – 0.99) (de Haas et al., 2012), indicating the potential to estimate parameters jointly over the two populations. Indeed, for the combined population DMI had ($h^2 = 0.29 \pm 0.03$) comparable to a large study with over 10,000 animals from 9 countries in the gDMI consortium (Berry et al., 2014).

The h^2 estimates for BW are well within the reported ranges in DNK for multiparous Nordic cattle breeds (0.52 – 0.77) (Sondergaard et al., 2002; Sloniewski et al., 2005) and primiparous Dutch Holstein Friesian cows (0.22 – 0.74) (Koenen and Veerkamp, 1998; Manzanilla Pech et al., 2014). Similarly, h^2 estimates for FPCM were within range of previous findings in NLD (0.22 – 0.43) (Manzanilla Pech et al., 2014). However, the h^2 estimate for FPCM was larger than those reported for ECM in DNK (0.27- 0.39) (Sondergaard et al., 2002; Lassen and Løvendahl, 2016). Recognizing that the two trait definitions differ due to different linear combinations of milk yield, fat yield and protein yield, with FPCM 1.9% MY, 64.7% fat yield and 33.4% protein yield (CVB, 2008) and ECM 1.2% MY, 60.5% fat yield and, 38.2% protein yield (Sjaunja et al., 1991). The choice of FPCM was driven by need to make findings of this study comparable to international research efforts for GHG mitigation (FAO, 2010).

Greenhouse gas related traits. The h^2 estimates of CH₄c and CO₂c were consistent across datasets with estimates in the combined country dataset of 0.22 and 0.23, both significant from 0, respectively. These are substantially larger than those reported for primi- and multiparous Dutch Holsteins Friesians in commercial

herds CH₄c ($h^2 = 0.11 \pm 0.02$) and CO₂c ($h^2 = 0.12 \pm 0.02$) (van Engelen et al., 2018), of which the primiparous cows are part of this study. However, the estimates of (van Engelen et al., 2018) are based on the mean CH₄c and CO₂c per milking, whilst the present study makes use of an average per week of lactation and thus far more repeated measures per animal as was shown to be an effective error reduction technique with sniffers (Difford et al., 2016). The present estimates are also larger than the ratio of CH₄c to CO₂c (CH₄c/CO₂c) in multiparous Danish Holstein ($h^2 = 0.16 \pm 0.04$) (Lassen and Løvendahl, 2016) and Dutch Holstein Friesians cows ($h^2 = 0.03 \pm 0.01$) (van Engelen et al., 2018). Although CH₄c and CO₂c are short term breath concentration measures, the h^2 estimates are comparable to that of 24 hr total mass flux CH₄ production from respiration chambers in over 1000 growing Aberdeen Angus ($h^2 = 0.22 \pm 0.06$) (Donoghue et al., 2016) and 1225 dual purpose mixed and composite sheep ($h^2 = 0.29 \pm 0.05$) using portable accumulation chambers (Pinares-Patiño et al., 2013).

4.4.2 Genetic correlations

Feed Intake related traits. DMI was highly genetically and phenotypically correlated with FPCM and moderately with BW. With genetic correlations for FPCM in DNK ($r_g = 0.79 \pm 0.05$), NLD ($r_g = 0.68 \pm 0.05$) and combined ($r_g = 0.67 \pm 0.05$) in range of estimates in repeatability model studies (0.57 – 0.86) (Huttmann et al., 2009; Manzanilla Pech et al., 2014; Manzanilla-Pech et al., 2016). Genetic correlations between DMI and BW were somewhat different across datasets with (0.09 ± 0.16) in DNK and (0.38 ± 0.07) and combined (0.39 ± 0.07), although the magnitude of standard errors is limiting inferences into differences, the estimates were close to the ranges reported in literature (0.17 – 0.45) (Huttmann et al., 2009; Manzanilla-Pech et al., 2016). Genetic correlations between FPCM and BW ranged from negative (-0.51 ± 0.17) in DNK close to zero in NLD and in the combined dataset (-0.05 ± 0.08) falling within ranges of estimates in literature (ECM and FPCM) ($-0.36 - 0.04$) (Huttmann et al., 2009; Manzanilla Pech et al., 2014; Lassen and Løvendahl, 2016).

Greenhouse gas related traits. These are the first reported genetic correlations for breath gas concentration measures and FI related traits which makes comparisons and validations from literature challenging. Recently, genetic correlations have been reported for CH₄ production using the SF₆ method in 314 multiparous Australian Holsteins under grazing conditions (Breider et al., 2018). The genetic correlations between CH₄c and mass flux CH₄ production is not currently known and would be of value to assess the merit of CH₄c as an indicator trait for CH₄

production. The phenotypic correlation between CH₄c and a mass flux of CH₄ production is unlikely to be unity, based on R² values from comparisons of using the artificial reference cow or GreenFeed (C-Lock Inc, SD, USA) (Huhtanen et al., 2015; Wu et al., 2018). However, it is likely positive and in the range of ($r_p = 0.29 - 0.97$) depending on extraneous factors like cow muzzle position from the sniffer inlet and wind speed (Huhtanen et al., 2015; Wu et al., 2018). Moreover, the correlation between CH₄ production and FI related traits can provide some insight into the expected nature of CH₄c and FI traits. Breider et al (2018) found CH₄ production to be moderate genetically correlated to DMI ($r_g = 0.34 \pm 0.22$), MY ($r_g = 0.26 \pm 0.27$) and BW ($r_g = 0.42 \pm 0.16$). Similarly, (Lassen and Løvendahl, 2016) reported genetic correlations between the ratio of CH₄c/CO₂c with ECM of ($r_g = 0.04 \pm 0.06$) and ($r_g = -0.12 \pm 0.08$) with BW in multiparous Danish Holsteins. Supporting the notion that genetic correlations between FI related traits and CH₄c are in the expected range.

Interestingly, the genetic and phenotypic correlations between CH₄c and CO₂c were close to unity with (0.95 and 0.91, in DNK respectively and 0.95 and 0.87 in NLD). Studies using the intensive gold standard method have reported phenotypic correlations as high as 0.96 between mass flux CH₄ production and CO₂ production (Aubry and Yan, 2015), however no studies have reported genetic correlations between these traits. With such high genetic correlations between the two gases, it is imperative that these results are confirmed in future studies, preferably employing as alternative methods of gas measurement to ensure these correlations are not an artifact of the gas sampling method.

4.4.3 The relationship between GHG traits and feed efficiency

Both GHG traits had weak to moderate negative genetic correlations with gRFI in DNK and NLD datasets with large SE (Table 4.4). In DNK, 478 primiparous cows had repeated weekly records for GHG traits, 44% of which had full lactation concurrent records for GHG traits and FI traits. In NLD 660 primiparous cows had repeated weekly records for GHG traits for part of lactation, 42% of which had DMI records in the first 70 days in milk, resultantly no cows had concurrent GHG and FI traits. Combining the two datasets greatly reduced the standard errors on genetic correlations between GHG and FI traits, likely due to having larger numbers of cows with full lactation concurrent records with DMI and GHG traits. Interestingly, in the combined county dataset moderate positive genetic correlations between gRFI and CH₄c ($r_g = 0.33 \pm 0.15$) and CO₂c ($r_g = 0.32 \pm 0.15$) which were significant from zero.

The relationship between CH₄ emission and feed efficiency is a hotly contended topic, with inconsistent results between phenotypic RFI and CH₄

production. Positive phenotypic correlations ($r_p = 0.10 - 0.56$) between residual energy intake and CH_4 production predicted from energy intake and the ratio of $\text{CH}_4\text{c} / \text{CO}_2\text{c}$ from sniffers in concentrate feeders were reported for 412 lactating Nordic red cattle in a long term study (Negussie et al., 2014). Similarly, a study using the GreenFeed system found efficient phenotypic RFI groups to have significantly reduced DMI, CH_4 and CO_2 production as compared to inefficient RFI groups (Hailemariam et al., 2016). To date only a single genetic correlation of ($r_g = 0.72$) between CH_4 production predicted from gross energy intake and phenotypic RFI has been reported (de Haas et al., 2011). Breider et al (2018) investigated the responses to selection in DMI, MY, BW and CH_4 production from SF6, using restricted selection indices in Australian Holsteins on pasture. However, the index for reducing DMI and restricting MY and BW (gRFI) showed a correlated increase in CH_4 production, indicating genetically feed efficient cows produce more CH_4 production. However, using the estimated parameters from Breider et al (2018) and the methods described in Kennedy et al (1993) it is possible to estimate genetic correlation between gRFI and CH_4 production which was small $r_g = -0.004$ (Breider et al., 2018).

The inconsistencies in identifying the genetic correlation between GHG traits and feed efficiency in the aforementioned studies could be due to differences in environment and diet (grazing vs TMR), trait definitions (gRFI vs phenotypic RFI) and CH_4 emission (CH_4 production vs CH_4c) as well as CH_4 measurement methods (Sniffers, GreenFeed and SF6). Within the present study however, the differences in genetic correlations between gRFI and GHG traits within countries and the combined dataset are due to the uncertainty around estimating gRFI and GHG traits due to limited concurrent full lactation records within country, as genetic correlations between FI traits and GHG traits did not differ substantially from the within country estimates. From the findings in the present study genetic improvement in gRFI will likely have favorable correlated responses in CH_4c and CO_2c , but far larger studies are needed to obtain more accurate estimates. Moreover, the genetic correlation between CH_4c recorded during milking and full 24hr mass flux CH_4 production is not currently known (Pickering et al., 2013). Should the genetic correlation between CH_4c and CO_2c traits prove to be positive with mass flux CH_4 production, there is scope to use breath gas concentration measures as large scale indicator traits for reducing CH_4 emissions in dairy cattle. The findings of these studies and the relative instability of gRFI within countries highlights the difficulties in estimating feed efficiency in small research or nucleus herds and this impacts the certainty around correlations with CH_4 emissions. However, they do suggest that reducing DMI whilst

maintain FPCM and BW is feasible and will likely result in decreases in CH₄c and CH₄ production.

4.4.5 Defining feed efficiency as a restricted selection index

Kennedy et al (1993) demonstrated that defining genetic partial regression coefficients of energy sinks (gRFI) derived from multi-trait analyses could overcome the limitations of missing data and lactation stage bias inherent in RFI by defining gRFI as the portion of DMI breeding values that are genetically independent of energy sinks. This multi-trait approach has further advantages of allowing extension to random regression models over lactation (Lu et al., 2015; Moraes et al., 2015; Shirali et al., 2015), and extended to horizontal modelling which allows for adjustment of traits measured infrequently for example traits recorded during the dry periods (Shirali et al., 2017). However, this method assumes that the genetic (G) and phenotypic (P) (co)variance matrices are estimated without error, which is rarely true in practice and thus the partial genetic regression coefficients can be biased.

Estimating G and P and subsequently gRFI within country was challenging as cows with records for DMI were limiting, particularly in DNK with 812 primiparous cows with DMI records, as compared to NLN with near triple this number. Combining datasets reduced the uncertainty around genetic correlations between gRFI and FI related traits. Despite the near unity genetic correlation in DMI in the two countries (de Haas et al., 2012), the possible presence of genotype by environment interaction between the two countries for gRFI and the possible introduction of bias to estimates from unknown extraneous factors in spatial and temporal data with additive genetics cannot be discounted. Neither country had sufficient numbers of cows with DMI records to estimate gRFI within country with high levels of certainty. The combined dataset allowed for estimation of feed efficiency (gRFI), which remained genetically uncorrelated to FPCM ($r_g = 0.05 \pm 0.09$) and BW ($r_g = 0.00 \pm 0.09$) but highly correlated to DMI ($r_g = 0.64 \pm 0.04$). Thus selecting for improved (reduced) gRFI will have a correlated reduction in DMI and likely have little to no unfavorable correlated genetic responses in FPCM and BW. Thus, it is feasible to select for reduced DMI without reducing FPCM and BW, effectively breeding for cows that require less feed to maintain production levels of contemporaries.

Prediction accuracy of selection indices for DMI and gRFI. When assessing the potential of a large scale indicator trait to a scarcely recorded and expensive goal trait, it is important to not only account for the genetic correlations but also the value of increased numbers of records of the indicator trait. The accuracy of selection indices for DMI and gRFI were assessed for bulls using different combinations of FI

and GHG traits based on the estimated combined country genetic and phenotypic (co)variances and numbers of daughters with single weekly averaged records. For DMI, including GHG traits had limited gains as compared to including FPCM and BW. However, for accuracy of gRFI the contribution of FPCM and BW was negligible (0.05) as compared to the modest contributions of GHG traits (0.30). The addition of DMI records to GHG traits increased accuracy further (0.47) and the highest accuracy of gRFI was achieved recording all traits (0.53). Similarly, Manzanilla-Pech et al (2016) evaluated the accuracy of selection indices for DMI and RFI using conformation traits (stature, chest width and body depth) along with BW and milk energy output in the USA and NLD Holstein populations. The bull accuracies for DMI using conformation traits were 0.43 in NLD and 0.63 in USA and further increased to 0.79 in NLD and 0.97 in USA with the addition of BW and milk energy output (Manzanilla-Pech et al., 2016). Although, Manzanilla-Pech et al (2016) predicted bull EBVs for phenotypic RFI, thus some genetic correlations between RFI and component traits remain, the maximum accuracies obtained using all available traits were low for NLD 0.17 but substantial for USA 0.97, the discrepancy likely due to differences in genetic correlations between RFI and predictor traits. In the present study, gRFI remained largely genetically uncorrelated to BW and FPCM which resulted in low accuracies of bull EBVs only using these traits (0.05), however, only using GHG traits increase accuracies 6 fold to 0.30 demonstrating the value of these traits as indicators of feed efficiency. In a recent study, Zetouni et al (2018) found modest correlations (-0.20 – 0.01) between stature, chest width and body depth) and CH₄ production predicted from the ratio of CH₄c/CO₂c and production traits, indicating CH₄c and CO₂c are likely not co-explaining the same genetic variation in DMI and RFI as body conformation traits. Given these findings, the achieved accuracy of bull EBVs for gRFI in the present study maybe be further increased by the addition of body conformation traits.

Table 4.5. Selection accuracies for bulls for dry matter intake (DMI) and genetic residual feed intake (gRFI) using differing recording schemes with feed intake related traits¹ and greenhouse gas related traits².

| Selection Indices* | Accuracy (DMI) | Accuracy (gRFI) |
|---|----------------|-----------------|
| DMI ₁₀ | 0.66 | 0.42 |
| CH4C ₁₀₀ + CO2C ₁₀₀ | 0.43 | 0.30 |
| BW ₁₀₀ + FPCM ₁₀₀ | 0.75 | 0.05 |
| DMI ₁₀ + CH4C ₁₀₀ + CO2C ₁₀₀ | 0.70 | 0.47 |
| DMI ₁₀ + BW ₁₀₀ + FPCM ₁₀₀ | 0.81 | 0.47 |
| BW ₁₀₀ + FPCM ₁₀₀ + CH4C ₁₀₀ + CO2C ₁₀₀ | 0.78 | 0.31 |
| DMI ₁₀ + BW ₁₀₀ + FPCM ₁₀₀ + CH4C ₁₀₀ + CO2C ₁₀₀ | 0.83 | 0.53 |

*Subscripts denote the number of daughters with a single weekly average record for a given trait within a selection index. ¹Feed intake-related traits: DMI = dry matter intake; fat and protein corrected milk yield = FPCM; and body weight = BW. ²Greenhouse gas related traits: natural logarithm of methane breath concentration = CH4c; and natural logarithm of carbon dioxide breath concentration = CO2c.

Implications. GHG traits, FPCM and BW show considerable potential as indicator traits for improving the accuracy of bull EBVs for DMI or gRFI. However, these indicator traits are no substitution for DMI traits, which remain the most costly and limiting trait. In a simulation study Wallén et al., (2017) estimated that genotyped contract herds with 4000 primiparous cows per year phenotype for RFI was feasible for a national breeding scheme giving genomic reference population accuracies comparable to progeny testing schemes with 250 daughter per sire. In the context of the present study; collating data from multiple research herds in DNK and NLD, over a period of three decades approached 3000 primiparous cows with DMI records. Resultantly, historical and geographical structure within the dataset which could potentially bias the estimation of genetic and phenotypic variation covariance matrices on which gRFI is estimated. This potential problem would be circumvented if suitably large numbers of cows are recorded over lactation for DMI and GHG traits concurrently each year under commercial conditions. By installing sniffer sensors in commercial AMS herds which had integral weighing scales, it was possible to obtain records on over 1000 primiparous cows with GHG traits, BW and FPCM in a tenth of the time it took to obtain 3000 primiparous cows with DMI records. This demonstrates a proof of concept as to how these predictor traits can readily be generated on a large scale under commercial conditions and can supplement DMI records recorded on a small scale in contract herds. Recent, promising developments in 3D camera technologies for measuring feed intake under commercial conditions

could further supplement availability of DMI and GHG data over lactation, but are still under development (Lassen et al., 2018). Furthermore, should the genetic correlation between CH₄c and mass flux CH₄ production using accurate and intensive methods be positive and there is an economic incentive to reduce CH₄ emission, there is scope for CH₄c to act as a large scale indicator trait for genetic selection of climate friendly cows.

4.4 Conclusion

This study showed significant genetic variation for the concentration of methane and carbon dioxide in the breath of lactating Holstein cattle. Furthermore, a restricted selection index approach demonstrated genetic variation for more feed efficient cows with reduced DMI but little to no decreases in FPCM and BW. The genetic correlations between GHG traits, FI traits and feed efficiency open up the possibility for GHG traits as large scale indicator traits for feed efficiency.

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Chapter 5

Host genetics and the rumen microbiome jointly associate with methane emissions in dairy cows

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Abstract

Background. Cattle and other ruminants produce large quantities of methane (~110 million metric tonnes per annum), which is a potent greenhouse gas affecting global climate change. Methane (CH₄) is a natural by-product of gastro-enteric microbial fermentation of feedstuffs in the rumen and contributes to 6% of total CH₄ emissions from anthropogenic-related sources. The extent to which the host genome and rumen microbiome influence CH₄ emission is not yet well known.

Results. This study confirms individual variation in CH₄ production was influenced by individual host (cow) genotype, as well as the host's rumen microbiome composition. Abundance of a small proportion of bacteria and archaea taxa were influenced to a limited extent by the host's genotype and certain taxa were associated with CH₄ emissions. However, the cumulative effect of all bacteria and archaea on CH₄ production was 13%, the host genetics (heritability) was 21% and the two are largely independent.

Conclusion. This study demonstrates variation in CH₄ emission is likely not modulated through cow genetic effects on the rumen microbiome. Therefore, the rumen microbiome and cow genome could be targeted independently, by breeding low methane-emitting cows and in parallel, by investigating possible strategies that target changes in the rumen microbiome to reduce CH₄ emissions in the cattle industry.

Key words: Host microbe interaction, Methane emission, Microbiability, Ruminotypes

5.1 Introduction

Methane (CH₄) is a potent greenhouse gas (GHG) with a climate change potential ~32 times greater than carbon dioxide (CO₂) (Holmes et al., 2013) and an atmospheric half-life of 12 years, which is substantially shorter than CO₂ (> 100 years) (IPCC, 2014). Therefore, reducing CH₄ emissions from anthropogenic-related sources has been identified as a key area for mitigating climate change with immediate effects (Gerber et al., 2013; IPCC, 2014). Livestock accounts for 14.5% of anthropogenic-related GHG emissions and enteric CH₄ emissions from ruminants accounts for 5.8% (Gerber et al., 2013). Furthermore, CH₄ emissions from livestock is predicted to markedly increase due to an expected doubling in the global milk and meat demand by 2050 (Alexandratos et al., 2006).

Ruminants, the most widespread livestock species, can digest a wide variety of high fiber feedstuffs due to the distinct microbiome in their rumen. Methane is a natural by-product of gastro-enteric fermentation of high fiber plant biomass by microbial enzymatic activity in the rumen (Hill et al., 2016). Bacteria, protozoa, and fungi in the rumen produce CO₂ and hydrogen (H₂), which are converted to CH₄, primarily by archaea known as methanogens. Approximately 99% of CH₄ emitted from cattle is released in the breath by eructation and respiration (Janssen and Kirs, 2008). The emission of CH₄ is also a crucial pathway for maintaining H₂ balance and ruminal pH, as the optimal conditions for anaerobic fermentation by the rumen microbial community is limited to a narrow range of partial pressure of H₂ and pH (McAllister and Newbold, 2008). Hydrogenase-expressing bacteria convert metabolic hydrogen from anaerobic fermentation into H₂ which is then converted to CH₄ via methanogenesis (McAllister and Newbold, 2008). Furthermore, emitted CH₄ has a caloric value and represents a 2-12% net loss of a cow's gross energy intake (Johnson and Johnson, 1995; Yan et al., 2010). Consequently, cattle and other ruminants with increased efficiency to digest high fiber feedstuffs but reduced CH₄ production could in principal benefit the global climate and concurrently improve the profitability and sustainability of cattle production.

Mitigation to decrease CH₄ production by cattle to date has been largely unsuccessful, as the available measures are temporary and not cumulative. Large international research approaches target the rumen microbial communities through feed additives (chemical or biological), feed formulations, and anti-methanogen vaccines (Moss et al., 2000). However, rumen microbial species rapid adaptation to changes in the substrate results in resistance to treatments and CH₄ production returns to pre-treatment levels (Hristov et al., 2013). Conversely, rumen

transplantation studies (transfaunation) show that the rumen bacterial community recovered to near pre-transfaunation composition after a short period of time (Weimer et al., 2010). This indicated the existence of a degree of host influence on rumen microbial composition (Weimer et al., 2010). Host genotype in cattle was reported to explain inter-animal differences in CH₄ production (Donoghue et al., 2016; Lassen and Løvendahl, 2016) and the rumen microbial community influenced CH₄ production (Roehe et al., 2016). However, empirical evidence linking the host's genetic influence over the rumen microbial community and CH₄ production is rather limited (Roehe et al., 2016).

A promising strategy is genetic selection for low CH₄ emitting cows, as it is sustainable, persistent, and cumulative over subsequent generations. Whether the host influences the rumen microbial community, and consequently CH₄ production, or the two interact to affect CH₄ production is currently unknown. If reduced CH₄ production in cows is a consequence of poor symbiosis with rumen microbes and thus fiber digestibility, there is a risk selection for reduced CH₄ production will act against the very symbiosis which has aided ruminants and rumen microbes' coexistence. Thus, the extent to which the rumen microbiome is under the host genetic influence needs elucidation. If host genetics impose a strong influence on rumen microbial composition, traits influenced by rumen microbes could be improved by using rumen microbial composition as indicator traits in selection. However, should host genetics impose a strong influence on rumen microbial composition and selection for CH₄ production proceed without cognizance of rumen microbial composition, there is a risk of unfavorable correlated responses in rumen microbial composition.

We hypothesized that: 1) the relative composition of the microbiome in the rumen is heritable i.e. controlled by host genome and 2) variation in methane emission from rumen is influenced by both the cow genome and rumen microbial content.

5.2 Materials and Methods

5.2.1 Experimental design

All handling of animals was conducted according to a protocol approved by The Animal Experiments Inspectorate, Danish Veterinary and Food Administration, Ministry of Environment and Food of Denmark (Approval number 2016-15-0201-00959).

Methane emissions from 750 lactating Holstein cows in five commercial herds were recorded using a portable Fourier Transform Infrared unit (FTIR; Gasmeter

DX-4000, Gasmot Technologies, Helsinki, Finland)(Lassen et al., 2012; Lassen and Løvendahl, 2016) and one research herd using a permanently installed non-dispersive infrared (NDIR; Guardian NG/Gascard Edinburgh Instruments Ltd., Livingston, UK) (Difford et al., 2016a). Briefly, the FTIR and NDIR equipment were installed within the feed bins of automated milking systems (AMS) in each commercial herd with the FTIR for seven consecutive days and the NDIR were permanently placed in the research herd. The FTIR and NDIR device inlets were installed in the AMS feed bins and methane (CH₄) and carbon dioxide (CO₂) gas concentrations (ppm) sampled continuously every 5 s and 1 s, respectively(Lassen et al., 2012; Difford et al., 2016a). Cows were milked individually in the AMS and milked on average (18.2 ± 3.4) times during the seven-day period, for durations ranging from five minutes to 12.2 minutes. Mean CH₄ and CO₂ gas concentrations were corrected for environmental factors, including diurnal variation and day to day differences using a linear mixed model following (Difford et al., 2016a) to approximate daily averages. Measurement stability was assessed by model repeatability and used as data quality control. All herds practiced indoor feeding strategies with *ad libitum* access to feed and water. A total mixed ration (TMR) was provided, consisting primarily of rolled barley, corn silage, grass clover silage, rapeseed meal, soybean meal and up to 3 kg of concentrate supplement given during milking. Although all commercial herds employed a standardized TMR recipe, ingredient-specific differences among farms were expected to contribute to differences in TMR dietary values over herds.

Weekly mean values for milk yield and body weight were combined with weekly gas concentrations, as described in (Lassen et al., 2012) and applied to predict cow heat production(CIGR, 2002). During each week of CH₄ and CO₂ recording at different herds, milk samples were collected to estimate milk fat and protein percentages. Cow fat and protein corrected milk yield (FPCM) was estimated following the national recording scheme (RYK, Skejby, Denmark) (Sjaunja et al., 1991). Methane production (L/day) was estimated using the CH₄ to CO₂ ratio and predicted CO₂ emission (Pedersen et al., 2008) from the conversion of cow heat production units to CO₂ production, following (Madsen et al., 2010a) and then converted to (g/d) using CH₄ density at standard temperature and pressure.

Holstein cow pedigree records were traced in the Danish national database (NAV, Skejby, Denmark) as far back as 1926 to construct a pedigree-based relationship matrix for the quantitative genetic analysis.

5.2.2 Sampling rumen liquid fraction

Immediately following the CH₄ recording period, rumen content samples were drawn from individual cows by oral insertion of the probe “Flora Rumen Scoop” (Geishauser et al., 2012). Approximately 40 mL of the liquid fraction containing particulate matter was drawn from the rumen using this method. Trained technicians conducted the sampling to ensure correct probe insertion into the rumen following a previously established protocol (Geishauser et al., 2012), recognizing that the location of the flora rumen scoop may differ somewhat from sampling to sampling. The entire “Flora Rumen Scoop” was rinsed vigorously between animal sampling to minimize cross-contamination. Samples were labeled, immediately placed on ice, and transferred to the laboratory within two hours for further processing. Each 40 mL sample was mixed vigorously, a subsample of 1.2 mL rumen fluid was collected, and transferred to a 1.5 mL vial, then snap frozen in liquid nitrogen, before storing at -80 °C, until shipped on dry ice to a commercial sequencing company (GATC Biotech, Constance, Germany) for analysis.

5.2.3 DNA extraction, bacterial and archaeal 16S rRNA gene amplification, and sequencing

DNA extraction, sequencing library construction and sequencing were conducted by GATC Biotech (Constance, Germany). Rumen samples were defrosted at 4 °C overnight and vortexed until homogenous. A representative sample (500 µl) containing rumen liquid and solids was used for DNA isolation using the Qiagen QIAamp stool kit (Valencia, United States of America) following the manufacturer’s instructions, modified for the larger sample size (Johnson et al., 2005).

Two primer sets were used to create 16S rRNA libraries, one set for all bacteria and one set for all archaea. Universal bacterial 16S rRNA gene primers (covering the V1-V3 variable regions) 27F: 5'-AGAGTTTGATCCTGGCTCAG-3' and 534R: 5'-ATTACCGCGGCTGCTGG-3' were used to generate the bacterial amplicon libraries (expected amplicon size 508 bp) (Ward et al., 2012). Universal archaeal 16S rRNA gene primers (covering the V4-V6 variable regions) S-D-Arch-0519-a-S-15 5'-CAGCMGCCGCGGTAA-3' and S-D-Arch-1041-a-A-18 5'-GGCCATGCACCWCCTCTC-3' were used to generate the archaeal amplicon libraries (expected amplicon size 542 bp) (Klindworth et al., 2013). Following protocols standardized by GATC Biotech, PCR amplifications were conducted with GoTaq Green polymerase (Promega, Madison, USA) with 30 PCR cycles and a 60 °C annealing temperature for the archaeal amplicon libraries and 25 PCR cycles with a 60 °C annealing temperature for the bacterial amplicon libraries. The 16S rRNA amplicons were purified using the Axyprep Fragment Select bead purification system (Axygen Biosciences, New York, USA), according to the manufacturer’s instructions. The size and purity of the PCR product

was verified on a Fragment Analyzer using a High Sensitivity NGS Fragment Analysis Kit (Advanced Analytical Technologies, Ankeny, USA). Multiplex indices and Illumina overhang adapters were added to both amplicon libraries in a second PCR amplification round (six cycles), followed by Fragment Analyzer analysis to confirm the correct size of the amplicons (Advanced Analytical Technologies, Ankeny, USA). Ninety-six libraries were pooled in equimolar concentrations and sequenced with an Illumina sequencing instrument using the 300 bp paired-end read mode, according to the manufacturer's specifications. Approximately half the samples were run using the Illumina MiSeq platform and half with the HiSeq platform. The 300 bp paired end protocol was adapted to HiSeq by GATC Biotech. The specific samples entered into sequencing batches within each sequencing platform were recorded for subsequent significance testing to examine possible differences between sequencing batches and sequencing platforms in statistical analyses.

5.2.4 Bacterial and archaeal 16S rRNA gene sequence processing and OTU table construction.

Bacterial and archaeal sequence reads underwent quality control, processing and were clustered into operational taxonomic units (OTUs) using the LotuS pipeline (Hildebrand et al., 2014) with the following options: Sequence truncation length and minimum sequence length after barcode and primer removal was 230 bp. Minimum average sequence quality score was 27, the maximum number of ambiguous bases was 0, maximum homonucleotide run was set to 8. Sequences were filtered away if any of the 50 bp segments in a sequence had average scores below 25 or if the expected number of errors exceeded 2.5 in the binomial error model. The low-quality sequence ends were trimmed by applying a sliding window quality filter with a width of 20 bp and a minimum average quality score within the window of 25. Sequences were truncated if the probabilistic accumulated error exceeded 0.75. The reads were de-replicated and sequences with a minimum of 10 replicates were retained for OTU clustering within the LotuS pipeline. Sequence pairs were merged with Flash (Magoč and Salzberg, 2011) and clustered into OTUs based on sequence similarity (97%) with UPARSE (Edgar, 2013) and chimeric sequences removed with UCHIME reference-based chimera detection (Edgar et al., 2011). Representative sequences from each OTU were aligned with ClustalO (Sievers et al., 2011) and a phylogenetic tree built with FastTree2 (Price et al., 2010). Representative sequences, the OTU table, and phylogenetic trees were transferred to QIIME (version 1.9.0) (Caporaso et al., 2010), where further analyses were performed. Taxonomy was assigned to each OTU using the RDP classifier with a confidence level of 0.8 (Wang et al., 2007) using greengenes (gg_13_8_otus) as the reference database.

Unclassified OTUs and OTUs classified to non-target kingdoms were filtered from the OTU tables, i.e. only OTUs classified as k_Bacteria were maintained for the bacterial primer set and similarly OTUs classified as k_Archaea maintained for the archaeal primer set. Finally, samples with < 50,000 sequences were removed and OTUs containing < 10 sequences were filtered out of the OTU table.

5.2.5 Statistical models: Additive Genetic Variance Estimation

The linear mixed model utilized to estimate additive genetic variance is as follows:

$$y_{ijkl} = \mu + h_j + p_k + b_1(dim_i) + b_2(e^{-0.065 \times dim_i}) + a_i + e_{ijkl} \quad (1)$$

where y_{ijklm} is the observed phenotype, e.g. methane emission in grams/day; μ is the model intercept; h_j is the herd fixed effect ($j = 6$ levels); p_k is the parity fixed effect ($k = 4$ levels); b_1 is days in milk fixed regression coefficient ($dim = 1-350$); and b_2 is the Wilmink term fixed regression coefficient generated on dim to account for non-linearity in early lactation (Wilmink, 1987). Term a_i is individual animal random additive genetic effects $\sim NID(0, A\sigma_a^2)$, where σ_a^2 is the additive genetic variance and A is the pedigree derived numerator relationship matrix ($i = 750$ animals); and e_{ijkl} is the random residual $\sim NID(0, \sigma_e^2)$, where σ_e^2 is the error variance. The additive host genetic effects on relative rumen bacterial and archaeal abundance was estimated applying the same general equation as model 1 above, with the addition of the sequencing batch fixed effects nested within the sequencing platforms (11 levels). The analyses were performed using the DMU software (Madsen et al., 2010b).

5.2.6 Statistical models: Rumen Microbial Variance Estimation

The relationship among cows based on their similarity in rumen microbiome composition was estimated by constructing a microbial relationship matrix (M) inspired by Ross et al. (2012), where a metagenomic relationship matrix was created from a vector of aligned rumen microbial contig sequences. The matrix was computed as a variance-covariance matrix from rumen bacterial and archaeal abundance as follows:

$$M = \frac{XX'}{n} \quad (2)$$

where X is the matrix of natural log transformed bacterial and archaeal relative abundance for all animals and n is the number of bacterial and archaeal OTUs within the population. Matrix X is derived from OTU tables after filtering out OTUs, which were absent from more than 50% of the samples and were homogeneous. The matrix X was subsequently scaled and centered within sequencing instrument

(Miseq or Hiseq) to account for differences between instruments and recombined into a single matrix prior to the calculation of \mathbf{M} .

The variance explained by microbial composition was estimated employing models similar to equation 1, where the random effect of m_i was fit separately and jointly with a_i , i.e. random additive genetic effects. Term m_i is the rumen microbial effect for the i^{th} animal $\sim \text{NID}(0, \mathbf{M}\sigma_m^2)$, where σ_m^2 is the rumen microbial variance and \mathbf{M} is the microbial relationship matrix, described in (2), $i = 750$ animals.

5.2.6 Association between rumen bacterial and archaeal OTU's and host methane emission

The association between the relative abundance of each bacterial and archaeal OTU abundance with host methane production was conducted using linear mixed model analyses as proposed by Yu et al. (2006), with the exception that OTU effects were estimated in place of allele substitution effects for genetic variants, as performed in genetic association analysis. The significance threshold was calculated using a Benjamini Hochberg false discovery rate correction for multiple testing. There were 189 archaeal and 3894 bacterial OTUs tested, and the microbiome wide significant threshold at FDR of 15% was in $-\log_{10}(P)$ scale 2.17.

5.2.7 Microbial Community Analysis

A principal coordinate analysis (PoCA) was conducted to investigate similarities or dissimilarities using a distance matrix from the archaeal and bacterial rumen community composition. The Bray-Curtis coefficient was employed separately for the archaeal and bacterial OTU tables to create sample-summary matrices, which were further explored using non-metric multidimensional scaling (NMDS)(Bray and Curtis, 1957). The effects of environmental and genetic parameter effects on community structures were evaluated using the following model:

$$y_{ijkl} = \mu + sb_j + p_k + b_1(dim_i) + b_2(e^{-0.065 \times dim_i}) + a_i + e_{ijkl} \quad (3)$$

where y_{ijklm} is the observed phenotype, e.g. PCoA 1 and PCoA2 for bacteria or archaea; μ is the model intercept; sb_j is the sequencing batch run fixed effect nested within the sequencing platform ($j = 11$ levels); p_k , b_1 , dim_i , b_2 , a_i , and e_{ijkl} are as described in equation (1). Additive genetic effects of host could not be detected from the bacterial and archaeal community structures for PCoA 2 and only herd environmental effects were significant. The distribution of *a priori* defined high and low emitters along PCoA1 was tested for bacterial and archaeal community structures, respectively by means of Mann-Whitney tests.

5.3 Results

5.3.1 Variation in methane emission and its heritability in lactating dairy cattle

Methane concentration in the exhalation-breath of 750 lactating Holstein dairy cows from farmer herds in Denmark was measured individually during automated machine milking for one week. Within-week methane measurements had a high repeatability coefficient of 0.70 ± 0.02 (estimate \pm SE). Estimated average daily methane emission was 395.8 ± 63.5 g/d (mean \pm SD), which was consistent with reports from the literature (Lassen et al., 2016). Considerable variation in estimated CH₄ emission among cows was observed. The top 10% methane emitting cows (519.28 ± 28.5 g/d) had a 41% mean difference from the low 10% emitting cows (303.8 ± 11.9 g/d) (Supplemental Fig S5.1). Results from linear mixed model with pedigree records indicated methane emission was moderately heritable, 0.19 ± 0.09 (heritability coefficient, $h^2 \pm$ S.E), which was consistent with previous findings in lactating Holstein cows in Denmark (Lassen and Løvendahl, 2016).

5.3.2 Rumen bacterial and archaeal community composition

We identified 3,894 bacterial operational taxonomic units (OTUs, $\geq 97\%$ identity) and 189 archaeal OTUs, which were present in a minimum of 50% of the cow samples (50% threshold maximizes the variation in a binary trait i.e. presence or absence). Taxonomic classification revealed generic bacterial and archaeal composition. The predominant bacterial phylum found was *Bacteroidetes* $72.2\% \pm 6.5$ (mean \pm SD), followed by *Firmicutes* ($18.3\% \pm 5.6$) and *Tenericutes* ($2.8\% \pm 1.0$). *Abconditabacteria*, *Spirochaetes*, *Fibrobacteres*, and *Proteobacteria* each comprised less than 2%, and another 20 phyla constituted 1% of all sequence reads.

The archaeal community was dominated by two families, *Methanobacteriaceae* and *Methanomassiliicoccaceae* ($35\% \pm 22.1$) of the orders *Methanobacteriales* ($64.2\% \pm 22.2$; mean \pm SD) and the recently proposed order *Methanomassiliicoccales* and class *Thermoplasmata* (Iino et al., 2013), respectively. The remaining archaeal community was comprised of 10 families, which were low in abundance, cumulatively accounting for less than 1% of all archaeal sequence reads.

5.3.3 Additive genetic variance estimates of rumen microbiota

OTU abundance and OTU abundance collapsed at genus and family levels were used as microbial phenotypes. The heritability thereof was estimated using a linear mixed model with pedigree records (known as ‘animal models’), which partitions total variance into additive genetic and environmental variance (Henderson, 1975). We calculated 95% confidence intervals for OTU h^2 estimates and found for 6% of bacterial and 12% of archaeal OTUs, the estimates

were significantly higher than zero ($P < 0.05$), ranging from 16 - 44% (Fig. 5.1) and 18 - 33% (Fig. 5.2), respectively. Due to the high number of independent tests, we calculated false discovery rate (FDR) corrected P - values for h^2 estimates with a FDR threshold of 15% (Supplementary Table 5.1).

Heritability of bacterial and archaeal abundance was further estimated at the genus level. In total eight bacterial genera out of 144 showed significant h^2 estimates ranging from 0.17 to 0.25 (Table 5.1). Only a single archaeal genus, *Methanobrevibacter*, had a h^2 estimate significantly different from zero (0.22 ± 0.09). However, *Methanosphaera* and *Methanomicrococcus* might also be under host additive genetic control with heritability estimates approaching significance thresholds (Table 5.1).

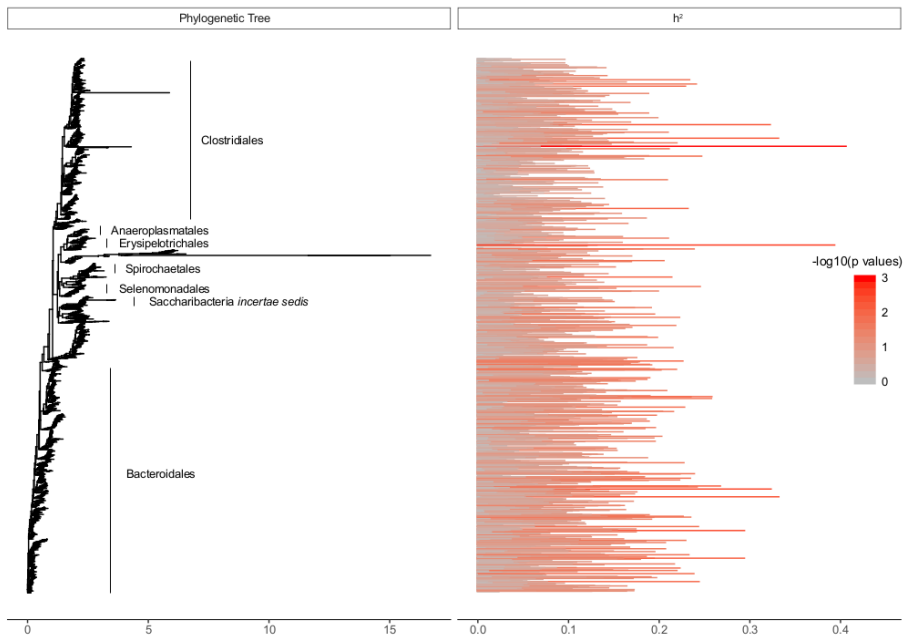


Figure 5.1 Phylogeny of 3,894 rumen bacterial OTU. Branch lengths represent substitution number per site calculated by FastTree2(81). Heritability estimates (h^2) for each OTU abundance are plotted with a horizontal bar and colored by taxonomic group classification.

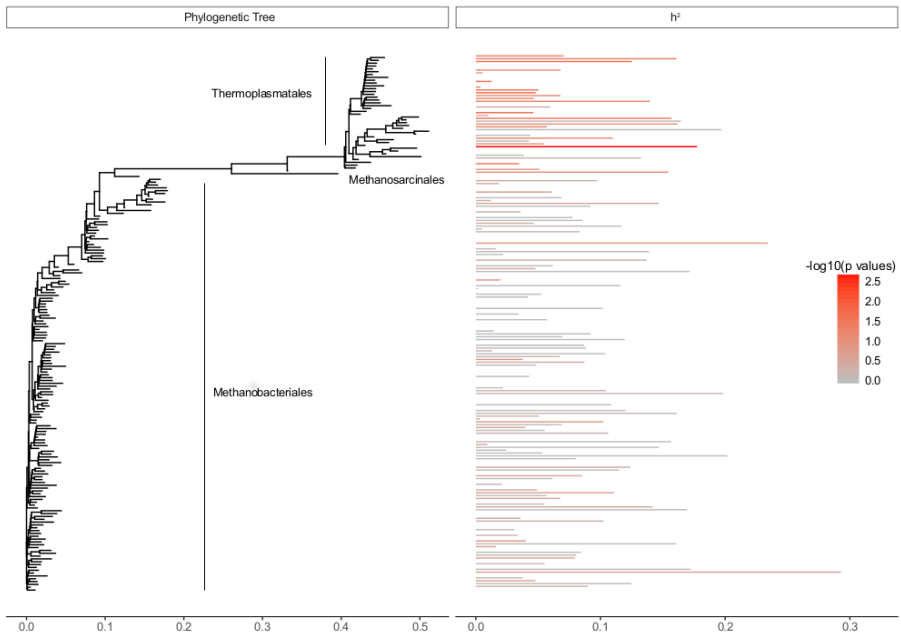


Figure 5.2 Phylogeny of 189 rumen archaeal OTU abundance. Branch lengths represent substitution number per site calculated by FastTree2(81). Heritability estimates (h^2) for each OTU abundance is plotted with a horizontal bar and colored by taxonomic group classification.

5.3.4 Association between microbiota abundance and methane production

Associations between relative bacterial and archaeal OTUs, genera abundance, and host CH_4 emissions were tested, while simultaneously controlling for environmental factors and familial structures common in livestock due to relatedness among study samples (Yu et al., 2006; Tsilimigras and Fodor, 2016). The OTU or genera log-transformed abundance present in > 50% of cows were fit as an explanatory variable in a linear mixed model for CH_4 production. Numerous significant OTUs were detected but failed to pass the threshold for multiple testing ($\text{FDR} \leq 0.15$) (Supplementary Table 5.1). This was a hypothesis-generating analysis and not directed at specific hypothesis testing therefore we reported the significance and FDR corrected values (Supplementary Table 5.1). Seven genera in total were detected, which exceeded the significance threshold at FDR of 15%. The $-\log_{10} P$ -values are plotted in Figure 5.3.

Table 5.1 Estimated heritability (h^2) and P-value for the relative abundances of bacterial and archaeal genera.

| Genus | Relative abundance (%) | $h^2 \pm SE$ | 95% Confidence Interval | P-value |
|------------------------|------------------------|-----------------|-------------------------|---------|
| Bacteria | | | | |
| Paludibacter | 0.01 | 0.25 ± 0.10 | (0.05 – 0.45) | 0.015 |
| Unclassified | 0.01 | 0.25 ± 0.09 | (0.04 – 0.41) | 0.08 |
| <i>Spirochaetaceae</i> | | | | |
| R4-45b | 0.01 | 0.23 ± 0.09 | (0.05 – 0.41) | 0.014 |
| F16 | 0.8 | 0.22 ± 0.09 | (0.04 – 0.40) | 0.018 |
| Unclassified | 0.04 | 0.21 ± 0.09 | (0.02 – 0.40) | 0.027 |
| <i>Endomicrobia</i> | | | | |
| Unclassified | 0.08 | 0.20 ± 0.09 | (0.01 – 0.39) | 0.36 |
| <i>Victivallaceae</i> | | | | |
| Unclassified | 0.02 | 0.19 ± 0.09 | (0.01 – 0.37) | 0.042 |
| <i>Proteobacteria</i> | | | | |
| Sporobacter | 0.01 | 0.17 ± 0.08 | (0.00 – 0.34) | 0.046 |
| Archaea | | | | |
| Methanobrevibacter | 55.8 | 0.22 ± 0.09 | (0.04 – 0.42) | 0.02 |
| Methanosphaera | 8.1 | 0.18 ± 0.10 | (-0.00 – 0.36) | 0.055 |
| Methanomicrococcus | 0.7 | 0.18 ± 0.09 | (-0.02 – 0.38) | 0.08 |

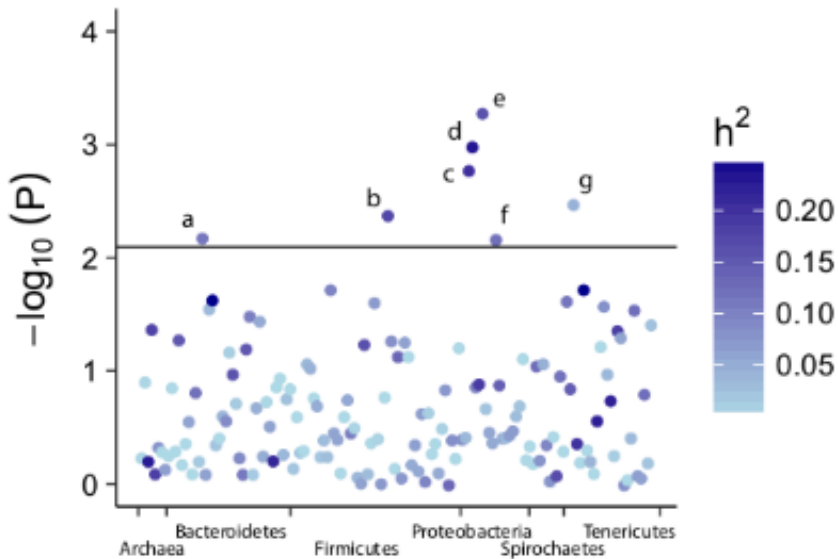


Figure 5.3. Manhattan plot of rumen bacterial and archaeal genera associations with methane emissions (g/day) colored by heritability (h^2) estimates. Color gradient indicates genera h^2 with light blue ($h^2 = 0$) ranging to dark blue ($h^2 = 0.30$). The y-axis is $-\log_{10}(P)$ for association tests. The horizontal line represents the Benjamini-Hochberg FDR 15% for multiple testing significance thresholds. Genera above the threshold are a) Unclassified *BS11* group; b) *Sporobacter*; c) Unclassified *Victivallaceae*; d) Unclassified *Lentisphaeria*; e) Unclassified *Alphaproteobacteria*; f) Unclassified *Rickettsiales*; and g) *Sphaerochaeta*.

5.3.6 Variation in methane emission attributed to cows' additive genetics and rumen microbiome

The relative proportion of variation in CH_4 emissions due to rumen microbial composition and host additive genetic components was estimated individually and jointly using linear mixed models. Likelihood ratio tests revealed that fitting either random effect of rumen microbial composition or individual cow's polygenic component fitted the data significantly better than the null model i.e. including only fixed effects ($P < 0.001$). The model fitting both random effects (microbial composition and polygenic component) was significantly better ($P < 0.001$) than models including only one random effect. The proportion of variance in CH_4 production explained by the microbiome, here defined as microbiability (m^2), was calculated in analogy to the heritability (h^2) (Difford et al., 2016b; Camarinha-

silva et al., 2017). The contrast between the two intra-class correlation coefficients h^2 and m^2 with their respective standard errors for all models are depicted in Figure 5.5. The m^2 of CH_4 emission estimated individually was 0.15 ± 0.08 (estimate \pm S.E) and the h^2 estimated individually was 0.19 ± 0.09 . Simultaneous estimates of both effects indicated slightly lower microbiability (0.13 ± 0.08), whereas h^2 exhibited a corresponding increase (0.21 ± 0.09) as compared to the preceding models fitting only one of the random effects. The combined microbial abundance and additive genetic effects were responsible for $\sim 34\%$ of the total phenotypic variation in CH_4 emissions.

5.4 Discussion

5.4.1 Additive genetic variation in methane emissions and bacterial and archaeal taxa abundance

The results of this study show that estimated CH_4 emissions from a dairy cow were partially under the influence of host (cow's) additive genetics, which explained 19% of the total variation. Of the rumen bacterial OTUs, a modest $\sim 6\%$ were associated with host additive genetics exhibiting significant heritability estimates (16 - 44%) (Fig. 5.1). Similarly, only $\sim 12\%$ of archaeal OTU abundance was influenced by host additive genetics, with heritability estimates ranging from 18 - 33% (Fig. 5.2). However, bacterial and archaeal heritability estimates failed to pass the threshold for multiple testing. Our test was conservative as a large number of taxa were analyzed with many OTUs having little or no influence by the host genome. Studies with larger sample sizes would give more reliable estimates of the heritabilities, especially for lower heritable OTUs. The h^2 estimates observed in this study were consistent with findings of intestinal microbiota in mice (Leamy et al., 2014; Org et al., 2015) and humans (Goodrich et al., 2014, 2016) and confirm that the majority of variation in rumen microbial abundance is due to factors other than host additive genetics (Rothschild et al., 2017). Interestingly, the patterns of h^2 with phylogeny differed between the bacteria and the archaea (Fig.5.1 and Fig.5.2). Heritable OTUs were distributed throughout the bacterial microbiome whereas archaea showed increased heritability within the *Thermoplasmatales*. This highlights the value of collating phylogeny with heritability estimates to focus research into possible mechanisms which predispose differential relative abundance of certain taxa across genetically related cows. The method employed to sample rumen contents is high-throughput and less invasive than surgical procedures, making it better suited to sampling large numbers of cows under commercial farm conditions. Large sample size is critical in genetic evaluations. However, it is important to note

that the floral rumen scoop is inserted into an undefined portion of the rumen and likely samples the liquid phase. Recognizing that rumen microbial communities differ between liquid, solid and epimural phases (Kong et al., 2010), studies testing the repeatability and representativeness of sampling are needed.

5.4.2 Associating rumen microbial taxa abundance with CH₄ emissions

We utilized linear mixed model analysis to test for associations between bacterial and archaeal OTUs, genera and families with estimated CH₄ emissions, while concurrently accounting for effects such as parity, lactation stage, herd of origin and familial structure from the pedigree. Several bacterial genera associated with CH₄ emission were detected. Out of these, four were found either to be affected by methane inhibitors or related to H₂ production and other methanogenesis substrates. Three were moderately heritable (0.17 – 0.25) (Supplementary Table 5.1). One of the identified bacteria, *Sporobacter*, with a mean relative abundance of 0.01% (*Ruminococcaceae*, *Clostridiales*, *Firmicutes*), belongs to a group with only a single cultured representative, *Sporobacter termitidis*, isolated from the intestine of wood-feeding termites (*Nasutitermes lujae*), also known for producing large amounts of CH₄. However, when this isolate was co-cultured with an archaea species, *Methanospirillum hungatei*, CH₄ was not produced. *S. termitidis* was found to generate acetate and methylsulfides, but not H₂ or CO₂, therefore interspecies H₂ transfer did not occur and facilitate CH₄ production (Grech-Mora et al., 1996). The recent discovery and proposed archaeal order *Methanomassiliococcales* species found to utilize methylsulfides and H₂ in methanogenesis (Borrel et al., 2014), provides a possible mechanism for methylsulfide producers to contribute to CH₄ production when H₂ producers are present. *Methanomassiliococcales* was prevalent in our samples (mean relative abundance 35%); therefore, *Sporobacter* could potentially be contributing to CH₄ production via a similar pathway.

We also detected *Sphaerochaeta* with a mean relative abundance of 0.01%, associated with estimated CH₄ production. Genomes from cultured *Sphaerochaeta* isolates revealed acetate, formate, ethanol, H₂, and CO₂ were potential fermentation end products (Caro-Quintero et al., 2012), many of which are methanogenic archaea substrates (Knapp et al., 2014). Furthermore, seed extracts from *Perilla frutescens* (*Lamiaceae*), a medicinal herb, decreased CH₄ production *in vitro* from rumen samples of lactating dairy cows and decreased *Sphaerochaeta* abundance (Wang et al., 2016). Interestingly, Caro-Quintero et al. (2012) reported up to 40% of the genes from *Sphaerochaeta* species were exchanged with members of *Clostridiales* (*Firmicutes*) and this inter-order-species horizontal gene transfer was

most extensive in mesophilic anaerobic bacteria, such as the conditions found in termite and ruminant guts (Caro-Quintero and Konstantinidis, 2014). Here 16S rRNA gene sequencing is used as a proxy for metabolic activity but cannot account for inter-order-species horizontal gene transfer. Therefore, full metagenome sequence may have an advantage over the 16S rRNA gene to describe rumen microbial contents.

One bacterial genus detected in the present study, which is positively associated with estimated CH₄ production, is classified in the yet uncultured BS11 gut group of the *Bacteroidales* (mean relative abundance 1.4%). The relative abundance of the BS11 group reportedly decreased concomitantly with CH₄ production by dietary methanogenic inhibitors, such as *P. frutescens* seed extract, mentioned previously (Wang et al., 2016), monesin and essential oil supplementation in dairy cattle (Boadi et al., 2004; Schären et al., 2017), and bromochloromethane in Japanese goats (Denman et al., 2015). Thus, supporting our finding of a positive association between BS11 and CH₄ production. Solden et al. (2016) employed metagenomics sequencing and shotgun proteomics approaches to phylogenetically and metabolically resolve the BS11 gut group. They resolved two genera within the group and both exhibited multiple pathways to ferment hemicellulose, a capability previously unknown for BS11. The resulting fermentation end products included acetate, butyrate, propionate, CO₂, H₂ (Solden et al., 2016) the latter two being methanogenesis substrates. Genes encoding 'fucose sensing' pathways were found for only one of the proposed BS11 genera, offering a possible mechanism for interaction between genes in the BS11 group and the host (Roehe et al., 2016). However, further studies are needed to elucidate the links between CH₄ inhibitors, host genes and CH₄ production.

Due to the absence of cultured rumen bacteria isolates, an understanding of the metabolic function in many bacterial genera remains in its infancy. However, from the isolates discussed above, results suggested CH₄ emissions depend on abundance of bacterial taxa that produce substrates for methanogenesis, such as H₂. Remarkably, associations between archaeal relative abundance and estimated CH₄ production were not detected in the present study, despite the knowledge that archaea are directly responsible for CH₄ production. A meta-transcriptome study in sheep found archaeal transcription pathways and not simply abundance, contributed to inter-animal differences in CH₄ production (Shi et al., 2014). This study was congruent with conclusions reached in two recent reviews, which examined results from dairy cattle and other ruminant studies employing 16S rRNA (Tapio et al., 2017) and 'meta-omics' approaches (Wallace et al., 2017), where bacteria

abundance produced and utilized H₂ or stabilized pH, which affected CH₄ emissions and feed efficiency and archaeal activity matched substrate availability.

5.4.3 Associating rumen community structure with methane emissions

The combined effects of the bacterial and archaeal community structure (beta diversity) on estimated CH₄ emissions were investigated by conducting PCoA on the archaeal and bacterial communities, which revealed 2-3 clusters for archaea (Fig. 4a) and two clusters for bacteria (Fig. 4b). Beta diversity is a non-parametric distance measure used in microbiology and ecology to assess the differences between environments or samples (in this case cows) as opposed to alpha diversity which takes into account the diversity within cows. Clusters of a similar nature were first reported in intestinal bacterial community types in humans (Wu et al., 2011; Arumugam et al., 2013), chimpanzees (Moeller et al., 2012), mice (Hildebrand et al., 2013) and pigs (Ramayo-Caldas et al., 2016), referred to as “enterotypes”, and found associated with specific host phenotypes. This concept was extended to sheep rumen bacterial communities and referred to as “ruminotypes” (Kittelman et al., 2014). The ruminotypes observed herein followed a continuous gradient and did not form discrete clusters, which is consistent with the latest findings in microbiome stratification (Knights et al., 2014). Importantly, we found that animal and farm factors like herd of origin, parity and lactation stage, as well as technical factors, i.e. sequencing batch, contributed to the observed variation and stratification in ruminotypes. Similar findings were reported in rumen bacterial richness at different lactation stages and over different parities (Jewell et al., 2015), suggesting later parities (higher parity cows are older) decreased bacterial richness and increased production (Lima et al., 2015). We detected a moderate heritable genetic component acting along PCo1 axis, with h² of 20% for bacterial and 39% for archaea, when controlling for lactation stage and parity, demonstrating the first evidence of host additive genetic influence on rumen bacterial and archaeal community structure (beta diversity). All the above-mentioned factors contribute to microbiome structure and associations with host phenotypes.

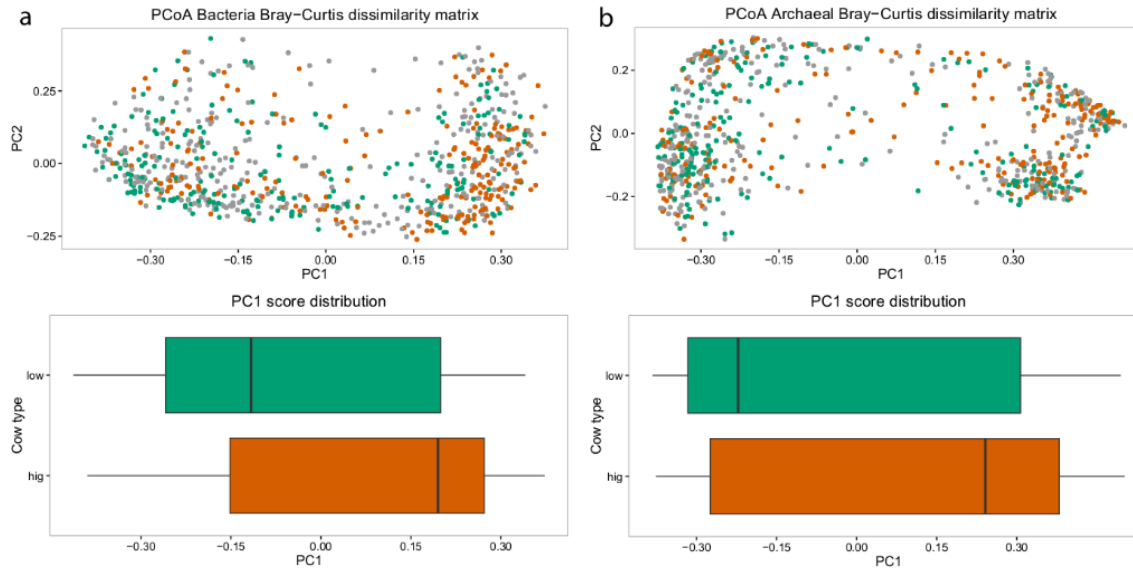


Figure 5.4 Principal coordinate analysis (PCoA) of rumen bacterial community diversity (a) and archaeal community diversity (b) based on 16S rRNA amplicon sequencing contrasting 10% highest methane emitters (orange), 10% lowest methane emitters (blue), and 80% intermediate emitters (grey). Distribution of high and low emitters along PCo1 showed significant differences ($P < 0.001$) for both figures.

An association was detected between the highest and lowest CH₄ emitters and bacterial and archaeal ruminotypes along PCo1, however, ruminotype cluster memberships were not exclusive to high and low emitters. This suggested ruminal bacterial and archaeal community structure provided a modest contribution to CH₄ emission. Kittelmann et al. (2014) surveyed microbial community composition in multiple sheep cohorts with low and high CH₄ yield (methane emission per kg dry matter intake, CH₄/DMI). A ruminotype “S” associated with low CH₄ yield and enriched with *Sharpea azabuensis* was reported. A follow up study in sheep also found low CH₄ yielding sheep to be associated with ruminotype “S”, enriched with *Sharpea* spp. It was hypothesized a smaller rumen size and higher turnover rate promoted faster growing bacteria, such as *Sharpea*, which favor hetero-fermentative growth on soluble sugars, resulting in lower H₂ production and subsequently decreased CH₄ formation by hydrogenotrophic methanogens (Kamke et al., 2016). Smuts et al. (1995) reported passage rate (and consequently turnover rate) in sheep was heritable, indicating a possible mechanism for host genetics to influence ruminotypes. Methane emission phenotypes differed between the sheep and the present study. Kittelmann et al. (2014) assessed the amount of CH₄ production per unit of DMI but not CH₄ production directly. DMI measurements are not currently recorded on dairy cattle under commercial farms due to the high costs and therefore, CH₄ emissions in the present study could not be corrected for feed intake. In light of the differences in phenotype definitions and similarities in ruminotypes between studies, it would be of interest in future work to obtain DMI records on cows and test if the ruminotypes observed show an increased relationship with CH₄ yield. The heritability estimates for PCo1 and PCo2 indicates these measures could potentially be used as indicator traits in genetic selection should they be highly correlated to a trait of interest, however PCo1 and PCo2 (beta diversity) does not account for the total rumen microbial variation within and between individuals.

The method employed to measure CH₄ production in the present study is high throughput and non-invasive, making it practically viable for measuring large numbers of animals under commercial farm conditions. However, the cost trade off of this method is that it makes use of milk yield and body weight in the estimation of CH₄ production. Validation of this method with the ‘gold standard method’ (climate respiration chambers) has yielded highly correlated ($r = 0.8 - 0.89$) and concordant (concordance correlation coefficient = 0.84) results in dairy cattle (Garnsworthy et al., 2012; Negussie et al., 2016). However, the effects of body weight and milk yield on estimation of CH₄ cannot be discounted and further research into the relationships between these variables and the rumen microbiome would be of value.

5.4.4 Quantifying variation in CH₄ due to cow additive genetic effects and rumen microbiome

In this study, we quantified the combined effects of all rumen bacterial and archaeal OTUs simultaneously on estimated host CH₄ emissions using a microbial relationship matrix among cows. This is a parametric approach similar to assessing both alpha and beta diversity, as total rumen microbial variation within and between individuals is taken into account simultaneously. We expressed the combined effects as the variance ratio due to microbial composition to the total variance in estimated CH₄ emissions (m^2 , microbiability), an analogy to h^2 . Estimated CH₄ emissions had 15% m^2 , indicating the combined rumen bacteria and archaea abundance of dairy cattle was associated with a considerable amount of variation in estimated CH₄ emissions among animals. Ross et al. (2012) first proposed the generation of metagenomic relationship matrices in dairy cattle and reported a CH₄ emission prediction accuracy of 0.47, explaining 22% of the total variation in CH₄ production (Ross et al., 2013). However, Ross et al. (2013) did not have sufficient data to estimate h^2 or microbiability (m^2) in CH₄ production. A study with 207 pigs employing 16S rRNA sequencing of gut microbes, found eight of the 49 bacterial genera to be heritable and estimated m^2 and h^2 for feed intake ($m^2 = 0.16$, $h^2=0.42$), daily gain ($m^2 = 0.28$, $h^2=0.11$) and feed conversion ratio ($m^2 = 0.21$, $h^2=0.19$) (Camarinha-silva et al., 2017). Only daily gain had higher h^2 compared with m^2 . These findings suggest agreement with holobiont theory, where variation in the genome and microbiome can cause variation in some complex traits, on which artificial, natural selection and genetic drift can act (Zilber-Rosenberg and Rosenberg, 2008; Bordenstein and Theis, 2015). However, the aforementioned study did not have adequate numbers of animals to estimate m^2 and h^2 simultaneously to assess the relative interactions between additive genetics and the microbiome. Thus, it was unable to assess if host additive genetics co-influences the microbiome and variation in phenotypes.

In contrast, we estimated m^2 and h^2 concurrently to examine the shared information between the two effects. Microbiability of estimated CH₄ production decreased by two percentage points to 13% and h^2 exhibited a corresponding increase from 19 to 21%. This result indicated host genetic effects do interact with the microbial community composition but are not the primary mechanism for host genetic effects on estimated CH₄ emissions. A possible explanation for the negligible amount of shared influence between the two relationship matrices might be the small percentage of heritable bacterial and archaeal OTUs. This implies that the rumen bacterial and archaeal communities affected estimated host CH₄ emissions independently and host genetics influenced a small portion of these bacteria and

archaea. The combined host additive genetics and rumen microbial community composition explained $\sim 34\%$ of the total variance in estimated CH_4 emissions in dairy cattle. Thus, breeding for low CH_4 production can be expected to result in limited correlated genetic responses to shape the rumen microbiome and breeding can likely proceed without taking cognizance of the rumen microbiome for this trait. However, larger studies estimating genetic correlations between rumen microbiota and CH_4 emissions and better functional annotation of rumen microbiota are needed to confirm this.

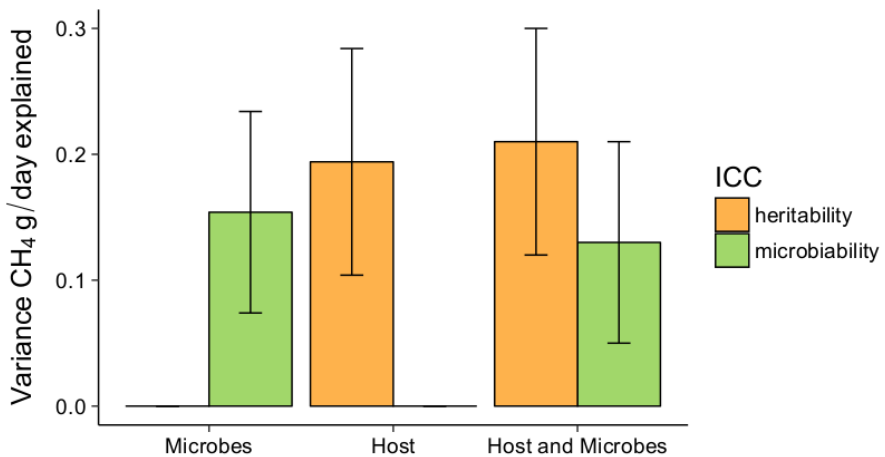


Figure 5.5. Proportion of variance in CH_4 explained by different sources (Intra-class correlation coefficients, ICC) due to additive genetic effects (heritability) and rumen microbe content (microbiability), with respective standard errors when fit separately or jointly.

Microbiability estimates can be used as a tool for quantifying the cumulative effects of microbial abundance on phenotypes, e.g. complex diseases and quantitative traits. However, further research is required to elucidate the biological mechanisms shaping microbiability. For example, animal factors known to affect CH_4 production and rumen microbial populations, such as passage rates or individual differences in feed intake might influence microbiability estimates. Human intestinal microbiome studies finds that numerous disease phenotypes are associated with microbial richness, species abundance, and microbial community structure (Le Chatelier et al., 2013; Falcony et al., 2016). Subsequent work using stool consistency and opaque markers as proxies for colonic transit time found all three metrics and disease phenotypes are partially confounded with colonic transit time (Roager et al.,

2016; Vandeputte et al., 2016). Similarly, in sheep studies, low CH₄ yielding sheep are associated with lower retention time and smaller rumens (Goopy et al., 2014), relationships with specific rumen microbial clusters (Kittelmann et al., 2014) and different bacterial and archaeal species (Kamke et al., 2016). Therefore, studies are needed to determine if microbial differences among subjects associated with phenotypic differences are causative or are consequences of unknown extraneous factors. It is also necessary to clarify the mechanisms which allow rumen microbes to be passed on to successive generations, to assess the efficacy of perturbations of the rumen microbiome such as probiotics and rumen transplants aimed at desired changes to the rumen microbiome and associated changes in phenotypes (Bickhart and Weimer, 2017). Regardless of the underlying biology, quantifying the relative contribution of rumen microbes and additive genetics to complex phenotypes helps characterize whether the host genome and microbiome are acting jointly as a holobiont and highlights the merits of targeting microorganisms to achieve a specific change in a phenotype or selective breeding. Furthermore, providing additional information, such as relative abundance of rumen fungi and protozoa, or ‘meta-omics’, including meta-transcriptomics or meta-proteomics data can be readily adopted and incorporated into this methodology, offering insights into economically important livestock and disease traits in humans.

5.4 Conclusion

Methane production by dairy cows is not only influenced by factors such as feed intake and composition among others, but also the cow’s individual genetic composition and rumen microbial composition. Each cow’s additive genetic effects influence a modest amount of variation in the abundance of a small percentage of rumen bacterial and archaeal taxa, and thereby contribute to variation in rumen microbiome composition and function. We detected associations between CH₄ emissions and rumen bacteria abundance, which are known to produce methanogenesis substrates, suggesting bacteria driven CH₄ production pathways. Although we detected a heritable component to ruminotypes, the association to CH₄ production was weak. Concurrently, host additive genetic effects and rumen microbes contributed to inter-animal differences in CH₄ production, however negligible interaction was observed between microbiability and heritability. Consequently, cow additive genetic effects on CH₄ emissions were largely unmodulated by cow additive genetic effects on rumen bacteria and archaea abundance. Strategies to reduce CH₄ emissions in ruminants can be optimized by a

multifaceted approach, for instance, selective breeding to unlock host's genetic potential and strategies which may effect desired changes in the rumen microbiota like rumen transplantation, and probiotics.

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6

General discussion

6.1 Overview

The main research area of this PhD project is to investigate the genetic variation contributing to methane emission, feed efficiency and the rumen microbiome. The use of multidisciplinary approaches such as method comparison, quantitative genetics and microbiology was proposed to better understand the genetic basis of the abovementioned fields.

The following general discussion is organised into several sections. The first section is dedicated to the discussion of the primary findings of Chapter 2 – 3 on the development, comparison and performance of sniffers for CH₄ emission recording. The second section is dedicated to the findings of Chapter 4, where we investigate the use of sniffers for CH₄ evaluation in practice and their genetic correlations to feed efficiency. Lastly, the findings of Chapter 5 investigated genetic variation in rumen microbial abundance, and integrating this information into better understanding biology and genetic basis of CH₄ production. At the end of the chapter, conclusions are drawn and some recommendations for future research are made.

6.2 Development of non-invasive sniffer methane phenotypes

The use of sniffers for recording breath gas concentrations and estimating CH₄ production is, like other methods, under continuous development. For genetic evaluations, defining the genetic correlation between the alternative method and the gold standard is crucial to evaluating the relative merit of the alternative method. This can be seen by the breeders equation where response to selection in the gold standard (R_{gs}) is a function of the heritability of the alternative method h^2_A , their genetic correlation r_g and intensity of selection (I) ($R_{gs} = I \cdot h_A \cdot r_g$). If the genetic correlation is suitably high ($r_g > 0.80$) (Robertson, 1959), and the alternative method is as heritable or higher, the alternative method can replace the gold standard. If the genetic correlation is moderate $r_g > |0.3|$, it may still be useful as an indicator trait. However, measurements using both methods on 10^3 - 10^4 related individuals are needed to estimate a genetic correlation with meaningful standard errors (Visscher, 1998). Even larger numbers are required if measurements are made on different animals or animals at different points in time, or environments (Bijma and Bastiaansen, 2014). The high cost, limited throughput and time or invasiveness of the gold standard drives the need for alternative methods. Other ways of assessing the merit of an alternative method are therefore needed prior to the investment in recording thousands of animals usually required for estimating genetic correlations.

Text Box 1. Accuracy and precision of phenotypes and random solutions

The statistical terms accuracy and precision are not synonymous at the phenotypic level (measurement) and genetic level (random solutions), see Barnhart et al, (2007a) for a discussion of definitions on agreement with continuous measurements. Phenotypic accuracy is a measure of ‘systematic error’ (also commonly called mean or intercept bias) and is the fixed difference between what is measured by a method and the known or ‘true value’. The gold standard method is taken to be the method, which measures the ‘true value’ with the least systematic error, and by definition is the most accurate method. Precision is a measure of the ‘random error’ and is the closeness of multiple measurements taken using the same method on a subject under the conditions the ‘true value’ has not changed or the change in ‘true values’ is accounted for statistically or by careful design. The phenotypic variance of measurements can then be decomposed into the between-subject variation (σ^2_b) (the variation between ‘true values’ in a population) and the within-subject variation (σ^2_e) (the random error or residual variation around ‘true values’). An alternative method can therefore be more precise than the gold standard method if the residual error of the alternative method is smaller than that of the gold standard. Precision cannot be determined in the absence of repeated measures per subject. The repeatability (t) of a method under the conditions of assessing precision described above, is a metric for comparing precision between methods. Defined as the ratio of between-subject variation to the total subject variation ($t = \sigma^2_b / (\sigma^2_b + \sigma^2_e)$). The accuracy of random solutions such as estimated ‘true values’ from repeated measures per subject or estimated breeding values from records on relatives is the correlation between the true random solutions and estimated random solution from a linear mixed model. This definition is somewhat confusingly similar to precision of phenotypes as is seen by the equation for random solution accuracy $r_{ay} = \sigma^2_b / \sqrt{(\sigma_b^2 (t + (1-t)/n), \sigma_y^2)}$ see Mrode, R., 2003 for proofs and extension to estimated breeding values. From this equation, it can be seen that the accuracy of random solutions are directly proportional to the repeatability or phenotypic precision and increases as the number of records per individual (n) increases.

Instrument or method comparison and validation studies, which identify and quantify phenotypic sources of (dis)agreement such as accuracy, precision and correlation (Barnhart et al., 2007a) are a viable first step. Knowledge of sources of (dis)agreement is indispensable as it facilitates further corrective steps, continued development of the method and informs the scope of the usage prior to investment in recording thousands of animals.

Precision can only be assessed if repeated measures are made per individual, which allows for partitioning of variance into between- and within-subject variance (precision). This facilitates the estimation of individual level correlations using between-subject (co)variances. Importantly, the repeatability serves as the upper threshold for heritability estimates and individual level correlations between methods serve as proxies for genetic correlations (Wolak et al., 2012; Dingemans and Dochtermann, 2013).

The accuracy of a method refers to measurements being close to the 'true value' of interest, a prerequisite of which is the gold standard, which is assumed to record the 'true value' (Hartnack, 2014). In the case of CH₄ emission, this is indirect calorimetry in RC. In the general introduction, we challenge the recording of the 'true value' of CH₄ emissions from dairy cattle, as it changes within day and from day to day. Furthermore, the RC often compare poorly within and between research facilities (Gardiner et al., 2015) and animal behaviour changes due to confinement within RC can change the true value of CH₄ emission due to changes in feed intake and may not necessarily be representative of the true value of CH₄ emission under grazing conditions (Pinares-Patiño et al., 2013). Regardless of the shortcomings of the RC method, assessment of method accuracy cannot be conducted without defining a gold standard. Precision is the variability between repeated measures on the same subject, thus, it is a measure of the random error (precision) of a method (ISO, 1994). Repeatability is the ratio of between subject variance to the total variance and used as a metric for precision but also the correlation between repeated measures of the same method. To this end, the repeatability of a method can be informative because if a method is not repeatable, it will not agree with another method (Hanneman, 2008).

Text Box 2. Correlations at phenotypic, individual and genetic levels

Pearson's product moment correlation is a popular method for assessing the linear association between two measurements. This metric assumes independence of error between measurements. In the case of single measurements per subject with two methods the phenotypic correlation is the ratio of covariance the product of standard deviations of each method ($r_p = \sigma_{p1,p2} / (\sigma_{p1} \cdot \sigma_{p2})$). However, if multiple measurements are taken per subject, as is the case when assessing precision, then the random errors on each subject are likely not independent. In this case, partitioning the variation into between- (σ_b^2) and within-subject variation (σ_e^2) can account for non-independence. To this end, the individual level correlations is more appropriate, as this estimates the correlation between estimated random subject solutions for each method $R_i = \sigma_{b1,b2} / (\sigma_{b1} \cdot \sigma_{b2})$. In the case that one or both methods are measured with imprecision ($t < 1$) and the random errors are not correlated between methods, then r_p in the presence of repeated measures is biased downwards and a bias predictor of the correlation between 'true values'. In the case of genetic correlations, the genetic relationships between individuals also affects the independence of measures. For repeated measures, the between-subject variation is further partitioned into additive genetic variation (σ_a^2) and permanent environmental variation (σ_{pe}^2) which includes non-additive genetics, maternal, common environmental and learned effects etc. If the permanent environmental variation is the same sign as the genetic variation (not always the case), then the individual level correlation is a better predictor of the genetic correlations than phenotypic correlations.

A prerequisite of method agreement is a positive linear association between measurements from each method. The most common metric for assessing linear associations is correlation, or Pearson's product-moment correlation. However, a high correlation ($r = 1$) does not ensure perfect agreement as it fails to take into account differences in accuracy ($\mu_1 \neq \mu_2$) or precision ($\sigma_{e1}^2 \neq \sigma_{e2}^2$) (Altman and Bland, 1983). Conversely, measurements of two methods can have the same means (accuracy $\mu_1 = \mu_2$) and the same magnitude precision ($\sigma_{e1}^2 = \sigma_{e2}^2$) but have no linear association ($r = 0$), thus high accuracy and precision is meaningless. Furthermore, accuracy, precision and correlation are interrelated facets of agreement. For instance, as imprecision increases, the statistical power to detect significant differences in accuracy decreases. Moreover, the correlation between

two variables is biased downwards in the presence of repeated measurements when one or both methods is measured with error (i.e. $t < 1$) (Spearman, 1904; Adolph and Hardin, 2007). The implications are that method comparison studies must use repeated measures per subject per method to estimate precision, but must also account for the attenuating effects of random error on the phenotypic correlation by estimating the individual level correlations. This is readily done in bivariate linear mixed effect models, commonly known in genetics as bivariate repeatability animal models (Hamlett et al., 2004; Roy, 2009; Downs and Dochtermann, 2014).

Text Box 3. Assessing agreement with CCC and CIA

Agreement between measurements from two methods requires that all three facets of (dis)agreement: accuracy, precision and correlation are assessed. Failing to assess all three facets can lead to miss-leading conclusions. Scaled agreement indices such as the concordance correlation coefficient (CCC) and the coefficient of Individual agreement (CIA) summarise all three facets into a single value ranging from 0 – 1, where 0 indicates no agreement and 1 indicates full agreement (Barnhart et al., 2007b). The CCC for repeated measures is:

$$CCC = (2 \cdot \sigma_{b1} \cdot \sigma_{b2} \cdot r_i) / (2\sigma_{b1} \cdot \sigma_{b2} + (\mu_1 - \mu_2)^2 + (\sigma_{b1} - \sigma_{b2})^2 + \sigma_{e1}^2 + \sigma_{e2}^2)$$

The CIA can only be estimated for repeated measures as:

$$CIA = (2 \cdot \sigma_{e1} \cdot \sigma_{e2}) / (2(1-r_i)\sigma_{b1} \cdot \sigma_{b2} + (\mu_1 - \mu_2)^2 + (\sigma_{b1} - \sigma_{b2})^2 + \sigma_{e1}^2 + \sigma_{e2}^2)$$

In general for both CCC and CIA, as the individual level correlations increase and difference in means ($\mu_1 - \mu_2$) (accuracy), differences in the between subject variation ($\sigma_{b1} - \sigma_{b2}$) decrease, the resulting agreement increases. However, for the CCC as the between subject variation increases so too does the CCC. As the between subject variation in population specific, comparing methods on a heterogeneous population may inflate the CCC between methods. Conversely, for the CIA, if one or both of the methods has higher imprecision (i.e. within-subject variation increases $\sigma_{e1}^2 + \sigma_{e2}^2$), the CIA maybe inflated. The two metrics are therefore complementary in assessing method agreement. Using the threshold of bioequivalence from the Food and Drug Administration, the threshold for good agreement with CIA > 0.455 and CIA > 0.80 for excellent agreement (Barnhart et al., 2007b). McBride, G.B. (2005) suggested a threshold of CCC < 0.90 as poor, $0.95 < CCC < 0.99$ as substantial and CCC > 0.99 as almost perfect.

A prerequisite of genetic and many other studies is the individual identification of cows for linking phenotypic information to familial relationship information (Falconer and Mackay, 1996). To this end, it is useful to know the sensitivity, or the true positive identification rate (where the cow is measured by sniffers and identified) and specificity, or the true negative identification rate (where the cow is not being measured and not identified). In Chapter 2, we developed and validated a time alignment algorithm and detected a systematic time series difference of 1 hour and 1 minute and a constant drift of two seconds per day between the digital clocks time stamping the gas recording and the AMS visitation. The 1 hour systematic time difference was due to inconsistent application of daylight saving time, which would result in a specificity of 50% and sensitivity of 50% over a 1 year period. After the application of the time alignment algorithm, both metrics were restored to 100% during the recording period. To the best of our knowledge, this potentially large source of error has not been reported elsewhere in CH₄ recording or the integration of time series in other automated recording devices including AMS, feed bins and activity recording devices. Sensitivity and specificity assessments are needed before method comparisons can proceed because inaccurate subject identification will result in poor comparisons, regardless of the methods compared.

6.2.1 Predicting CH₄ production from sniffers with the CO₂ ratio method

In Chapter 3, we compare CH₄ emissions using sniffers with subsequent measurements in RC for 10 Danish Holstein and 10 Danish Jersey cows, with the express purpose of assessing individual level correlations. Although time elapsed between the two methods, as well as a change in diet (underlying biology of the 'true values' will have changed), we can still compare the accuracy of the two methods, recognising the limitations of our study will cause an underestimation of agreement.

The CH₄ production estimated from sniffers using the CO₂ ratio method differed significantly ($P < 0.05$) by an overestimation of 9.9% corresponding to 52 L CH₄/day. Negussie et al. (2016) compared 21 Nordic red cows in the RC, with estimated CH₄ production, using a sniffer installed in a concentrate feeder and predicted CO₂ tracer gas from estimated metabolisable energy intake and found a non-significant 6% under prediction of 36.4 L/day. In the case of the comparison in Chapter 3, we cannot determine the cause of inaccuracy due to the confounding effects of diet and time but together with the results of Negussie et al. (2016), we can infer that sniffer methods using the CO₂ ratio method are less accurate (6 – 10%) or potentially as accurate as the RC.

Comparing sub-components of the sniffer CO₂ ratio method for accuracy, the predicted CO₂ tracer gas in Chapter 3 is 3.5% higher than recorded CO₂ production in the RC and not significantly different ($P>0.05$). In a study evaluating the accuracy of the predicted CH₄ production using CO₂ tracer gas against measured CH₄ production in the RC, Hellwing et al. (2013) found a significant systematic difference in accuracy of 16.6% or 96 L/day under prediction in 157 records on lactating Danish Holstein cows in 8 experiments covering 30 experimental diets. Similarly, a study comparing predicted CH₄ production from predicted CO₂ tracer gas and recorded CO₂ tracer gas in the RC, a non-significant under prediction of 3.9% or 5.8 L/day was observed in Dexter Heifers (Haque et al., 2017). From the findings of these studies, it appears that the prediction of CO₂ tracer gas contributes a negligible amount of deviations from accuracy across a range of breeds and diets. Differences due to the predicted CO₂ tracer gas were systematic and can be corrected for by the addition of a constant correction factor or through statistical modelling. However, prediction equations based on the relationship between CO₂ production and other phenotypes, such as milk production, can change over generations due to genetic selection and thus requires periodic re-evaluation. Furthermore, the prediction equations for CO₂ production, described in Madsen et al. (2010), do not take into account mobilization of body reserves and thus remain a potential source of bias, particularly in early lactation.

The recorded ratio of CH₄ to CO₂ in the AMS was however, 12% under that of the recorded ratio in the RC ($P>0.01$) (Chapter 3). In a comparison study using the GF system with or without active airflow (similar to sniffers), Huhtanen et al. (2015) found that the recorded ratio of CH₄:CO₂ production (mass flux measure) inconsistently differed from that of the ratio of CH₄:CO₂ concentration (concentration measure 'sniffer'). In the first experiment, a standard GF concentrate feeder system was used and the ratio of CH₄:CO₂ concentration under-recorded significantly as compared to the ratio of CH₄:CO₂ production ($P<0.001$) by 1.3% in Nordic red cattle (Huhtanen et al., 2015). However, in the second experiment, the GF system was retrofitted to AMS and significantly over-recorded as compared to the ratio of CH₄:CO₂ production ($P<0.001$) by 7.8% in Holstein Friesian cows (Huhtanen et al., 2015). Unfortunately, it is not possible to quantify from these studies whether one or both gases is contributing to the inaccuracy. From these studies, it is clear that deviations in the accuracy of the CH₄:CO₂ ratio will have consequences for the accuracy of predicted CH₄ production. Further work is required to ascertain the direct causes of inaccuracy due to partial confounders such as breed, installation device, sensor and diet. Particularly if the deviation in the CH₄:CO₂ is due to inaccuracy of the sensor for CH₄, CO₂ or both. This point is illustrated by the

findings in Chapter 2, where significant differences between CH₄ and CO₂ concentrations were detected using two different sniffer sensors in the same AMS feed bin.

As is reported in Chapter 3, the repeatability of average weekly predicted CH₄ production ($t = 0.58 \pm 0.11$) is not significantly different to subsequent average daily CH₄ production in the RC ($t = 0.61 \pm 0.12$). However, it is important to note that the time window is larger for the predicted CH₄ production as compared to the RC, which is shown in Chapter 2, to reducing residual error and thus increasing repeatability. The implications of this finding are that, in order to achieve adequate precision using sniffers, a longer recording period is required.

The correlation between predicted CH₄ production and recorded CH₄ production in Chapter 3 yielded individual level correlations of 0.77 ± 0.18 and similar ranking. In the comparisons of Negussie et al., (2016), a phenotypic correlation of 0.80 was back calculated from the CCC estimated between the two methods. In the correlation between sniffer using the eructation peak method and RC, CH₄ production was 0.89 (Garnsworthy et al., 2012). Although, deviating from unity, these correlations are substantial when placed in context of other methods of recording CH₄ production. For instance, an individual level correlation of 0.81 between CH₄ production from SF₆ and RC in beef heifers (McGinn et al., 2006), and a phenotypic correlation of 0.83 in Holstein x Norwegian red (Muñoz et al., 2012). Comparisons from the GF compared to RC have yielded phenotypic correlations of -0.6 – 0.24 in dairy cattle (Garnett, 2012; Hammond et al., 2015), but considerably higher in shorthorn cattle (0.85) (Velazco et al., 2015). Phenotypic correlations between SF₆ and GF have been reported in the range of 0.4 – 0.77 in dairy cattle (Hammond et al., 2015; Hristov et al., 2016). These results place the correlation between the sniffer method using CO₂ ratio and RC well within the range of more accepted methods.

Assessing the agreement through calculating the CCC and the CIA, scales the correlation between the methods based on the deviations from accuracy and precision. Both metrics have inherent shortcomings, which can be overcome by reporting both simultaneously, something which is rarely done in literature. The CCC is population specific and thus influenced by population heterogeneity, which can result in inflation when a highly heterogeneous population is measured with both methods (Barnhart et al., 2007b). The CIA is less influenced by population heterogeneity, but evaluates agreement relative to imprecision, thus a method with a low repeatability can have an inflated CIA (Barnhart et al., 2007b; Haber et al., 2010). From the comparison in Chapter 3 we find a (CCC = 0.56) and (CIA = 0.11), which indicates some agreement between methods, but the sniffer method using

CO₂ tracer gas is not equivalent to the RC. However, it is important to note that the time elapsed between measurements and the change in diets in the results from Chapter 3 will bias these metrics downwards. In the study of Negussie et al. (2017), the concordance between methods was CCC = 0.65 – 0.93, but in the absence of repeated measures from the RC it was not possible to calculate the CIA. Studies comparing SF₆ and RC found the CCC ranging from 0.10 (Garnett, 2012) to 0.68 (Muñoz et al., 2012) in dairy cattle and CCC = 0.79 in beef heifers (McGinn et al., 2006), GF and RC CCC = 0.10 and the GF and SF₆ CCC = 0.60 (Hammond et al., 2015). None of the studies reported on precision, within-cow coefficient of variation or repeatability and, thus, it was not possible to calculate the CIA to determine if high CCC values are due to population heterogeneity or true method agreement. What is clear from these findings is that there is some agreement between different methods but none of the methods agree sufficiently well with the RC at the phenotype level to be deemed equivalent. It is also clear that the sniffer method performs comparably, and in some cases, better than the SF₆ and GF methods reported. The implications of this are that if the intended purpose of the methods are to quantify individual cow CH₄ production for comparing for instance two feed additive or for taxing farmers, then none of the alternative methods agree suitably well at the phenotype level to replace the RC (see Text Box 3).

If the intended purpose of a method is genetic evaluations, then the genetic correlation between methods is the most informative for assessing how best to incorporate an alternative method into a selection index. However, individual level correlations, when methods are imprecise (i.e. $t < 0.80$), or single measure phenotypic correlations, when precision is high (i.e. $t > 0.80$), from carefully controlled method comparison studies are good indicators of the sign and magnitude of genetic correlations (see Text Box 2), but at a fraction of the cost. This point is illustrated in sheep where portable accumulation chambers (PAC) are a short-term total emission flux method alternative to RC. In earlier studies, Goopy et al. (2011) compared the two methods in 39 sheep, by measuring for 22 hours in the RC and then measuring 1 and 2 hours immediately after, in the PAC. The phenotypic correlation between methods was reported as 0.67 (from supplementary material on 13 sheep we find CCC = 0.32). In a subsequent genetic study on 3601 lambs with 4733 records in PAC and 8655 in RC, Jonker et al., (2018) found a genetic correlation of 0.67 ± 0.11 for CH₄ production between methods. In this case, the method comparison on 39 sheep and the genetic evaluation on 3601 sheep both show that PAC is not equivalent to the RC, but is a highly correlated but separate trait. Importantly, Jonker et al. (2018) did not report individual level correlations, but they did report phenotypic correlations in the presence of repeated measurements,

which were substantially lower $r_p = 0.27 \pm 0.02$ and a biased predictor of the genetic correlation due to attenuating effects of permanent environments and residual error. To this end, the individual level correlations found in Chapter 3 (0.77 ± 0.11) are very promising as proxies for genetic correlations between sniffers with the CO₂ ratio methods and RC.

6.2.2 Using CH₄ breath concentrations directly

The use of predicted CO₂ tracer gas poses some challenges for practical implementation in genetic evaluations, as predicted CH₄ production is a linear combination of ECM and BW, both of which are heritable traits and genetically correlated to CH₄ production (Lassen and Løvendahl, 2016). Moreover, ECM is heavily under selection. For this reason, the direct measurement of CH₄ concentration is attractive as a potential indicator trait. Garnsworthy et al. (2012) proposed the use of CH₄ concentration from sniffers installed in the AMS with a gas recovery correction factor to approximate CH₄ production and found this to be highly correlated (0.89) with CH₄ production in RC for 12 Holstein Friesians. Methane breath concentration is a separate but likely correlated phenotype to CH₄ production in the RC and, thus, evaluating agreement outside of individual level or genetic correlations is of limited use.

As the RC records the concentration of gases, it is possible to install alternative methods within the RC and make method comparisons for discerning sources of error. However, the RC are under very stable and high rates of airflow, if the airflow rate is constant and recovery rates are 1, the correlation between background corrected CH₄ concentration in the RC and CH₄ production from the RC will be 1. This means that comparing CH₄ concentration from alternative methods installed within the RC are, in effect, sensor comparisons and will not reflect the realized correlations between CH₄ breath concentration under variable conditions on farm and CH₄ production in the RC. However, these comparisons do give an impression of the effects of variable head movement on the sniffer methods under controlled atmospheric conditions.

As an unpublished sub-experiment of Chapter 3, we installed the sniffer instrument directly into the feed bin of the RC for 19 days, culminating in 2-3 measuring days on 5 Holstein and 4 Jersey cows. The objective of this comparison was to assess the performance of the CH₄ and CO₂ sensor under the carefully controlled conditions of the RC. Data was analysed using the models in Chapters 2-3, the results of which are presented in table 6.1 below:

Table 6.1 Sources of (dis)agreement¹ between sniffer sensors and RC sensors for daily CH₄ and CO₂ concentrations

| Measure ² | Mean (SE) | CV% | σ_b^2 | σ_w^2 | Rep | r_p | CCC | CIA |
|-------------------------|-------------------------|------|----------------------|---------------------|-------------------|-------|------|------|
| CH ₄ sniffer | 220 ^a (17.4) | 1.7 | 2,014 ^a | 202 ^a | 0.90 ^a | | | |
| CH ₄ RC | 214 ^a (12.9) | 1.6 | 1,408 ^b | 183 ^a | 0.87 ^a | 0.95 | 0.93 | 0.52 |
| CO ₂ sniffer | 2733 ^a (195) | 22.6 | 310,579 ^a | 69,408 ^a | 0.82 ^a | | | |
| CO ₂ RC | 2562 ^a (118) | 14.1 | 120,927 ^b | 10,135 ^b | 0.92 ^b | 0.8 | 0.66 | 0.56 |

¹Sources of (dis)agreement: CV% = coefficient of variation %, σ_w^2 = Inter-cow variance, σ_b^2 = Intra-cow variance i.e. precision, Rep = repeatability indicates the ratio of inter cow variation to total variation, not to be confused with repeatability coefficient used in agreement studies, r_p = Pearson's correlations coefficient, CCC = Lin's concordance correlation coefficient, CIA = Coefficient of individual agreement.² Gas concentrations in ppm. Superscripts denoted significant differences between methods.

From this sensor comparison, we find high agreement for both CH₄ and CO₂ sensors and no detectable difference in accuracy for both gases and precision for CH₄. However, the CO₂ sensor from the sniffer was significantly less precise than the CO₂ sensor from the RC. This validation study gives a high level of confidence in the performance of the sniffer sensors under stable and controlled conditions. Based on the high phenotypic correlations, CCC and CIA of CH₄ ($r_p = 0.95$, CCC = 0.93 and CIA = 0.52), and that the correlation between CH₄ concentration and CH₄ production in the RC is ~ 1 , one might incorrectly conclude that sniffer breath concentrations are the same trait as mass flux. Interestingly, for both sensor comparisons, the sniffer gave significantly higher between-animal variations, indicating differential animal to sensor interaction. The RC sensor samples from the inflowing and outflowing air and does not interact with the cow, but the sniffer was installed inside the feed bin of the RC. The larger between-animal variation between the two sensor sets could be due to variation in the distance from the cow head to the sniffer sensor sampling point or the exhalation and eructation rate between cows. Due to the constant airflow rates within the RC, it is unlikely due to variable air velocity.

The effects of cow head movement, variable proximity to the sensor sample inlet and dilution due to exhalation rate and air velocity have been assessed under controlled and barn conditions in two studies using dummy cows, the results of which are summarised in Table 6.2 below (Huhtanen et al., 2015; Wu et al., 2018). Huhtanen et al., (2015) released a known amount of CO₂ with an exhalation rate controlled using a foot pump and recorded the response using the GF as a flux or sniffer (GF study). Wu et al. (2018) used the artificial reference cow (ARC), a cylinder and piston system which controls exhalation rate, temperature and a mass flow controller for releasing known amounts of CH₄ at differing levels (Wu et al., 2015) and recorded the response using a sniffer (Gasmeter as used in Chapter 2) (ARC study). Both studies found that under controlled conditions, without variable air movement, cow head movement and exhalation rate, the captured fraction of gas (CFG) (relative difference in accuracy) was 100% (no difference in accuracy) and the CV less than that of RC 12% (an indicator of precision). However, when the distance from sensor sampling inlet to exhalation point was increased to 30 cm, the CFG decreased to 78% and 62% and the CV increased to 29.8% and 8.8% in the GF and ARC study, respectively. In general, from Table 6.2, it can be seen that air velocity in the feed bin increased from 0.09 to 0.24 m/s which further decreased the CFG from 26 to 35 and increased the CV from 17.1 to 26.9 %, respectively. Cow head movement had limited effects on CFG (102-109%) but increased the CV (10.4 - 20 %). Increasing exhalation rate systematically decreased CFG but had limited effects on CV.

Table 6.2: Effects of variable muzzle position, head movement, exhalation rate and air velocity on the captured fraction of gas (CFG) and coefficient of variation (CV) in GF study Huhtanen et al. 2015 and ARC study Wu et al. 2018.

| Muzzle position (cm) | Tidal Volume (L) | Breath (Breath/min) | Head movement | Sampling (L/min) | Exhalation (L/min) | Air velocity (m/s) | CFG (%) | CV (%) | Study |
|----------------------|------------------|---------------------|---------------------|------------------|--------------------|--------------------|---------|--------|-------|
| 0 | 3.5 | 30 | static | 1 | 111 | 0 | 100 | 9.8 | GF |
| 30 | 3.5 | 30 | static | 1 | 111 | 0 | 78 | 29.8 | GF |
| 0 | 4.4 | 30 | static | 4 | 132 | 0 | 100 | NA | ARC |
| 30 | 4.4 | 30 | static | 4 | 132 | 0 | 62 | 8.8 | ARC |
| 0 | 3.5 | 30 | static | 1 | 111 | 6 | 33 | 28.9 | GF |
| 30 | 3.5 | 30 | static | 1 | 111 | 6 | 13 | 4 | GF |
| 30 | 4.4 | 30 | static | 4 | 132 | 0.09 | 35 | 17.1 | ARC |
| 30 | 4.4 | 30 | static | 4 | 132 | 0.24 | 26 | 26.9 | ARC |
| NA | 3.5 | 30 | Moving ¹ | 1 | 111 | 0 | 102 | 20 | GF |
| NA | 3.5 | 30 | Moving ² | 1 | 111 | 0 | 109 | 10.4 | GF |
| 30 | 4.4 | 30 | Moving ³ | 4 | 132 | 0.09-0.024 | 24.8 | 25 | ARC |
| 0 | 3.5 | 60 | static | 1 | 141 | 0 | 100 | 12 | GF |
| 30 | 3.5 | 60 | static | 1 | 141 | 0 | 48 | 33.1 | GF |
| 0 | 4.7 | 60 | static | 1 | 282 | 0 | 93 | 11.1 | GF |
| 30 | 4.7 | 60 | static | 1 | 282 | 0 | 48 | 33.2 | GF |
| 0 | 4.4 | 30 | static | 4 | 174 | 0 | 62 | NA | ARC |
| 0 | 3.8 | 25.1 | static | 4 | 95 | 0 | 95 | NA | ARC |
| 0 | 5 | 34.9 | static | 4 | 174 | 0 | 36 | NA | ARC |
| 0 | 4.18 | 28.5 | static | 4 | 120 | 0 | 74 | NA | ARC |
| 0 | 4.62 | 31.5 | static | 4 | 145 | 0 | 54 | NA | ARC |

¹ side to side movement; ² up and down movement; ³ combined uncertainty based on direct observation of 20 cows 76% in feed bin, 16% out of feed bin and 8% above feed bin.

Key to these findings are that increases in CV are due to increased random error (i.e. imprecision) and decreased systematic error (i.e. accuracy) and a decrease in CFG reflects a systematic error. Random errors caused by variable muzzle distance, air velocity and head movements can be reduced through the partitioning of variance into between-subject variance and within-subject variation. Systematic errors can be accounted for through statistical modelling, for example, including air velocity as a linear covariate or through a correction factor if compared to a gold standard. However, cow level systematic differences such as head lifting, which was reported by Huhtanen et al. (2015) to be highly repeatable (0.74 – 0.82), or exhalation rate, which results in systematic differences between cows independent of CH₄ production (Wu et al., 2018), cannot be accounted for through statistical measures. All of these sources of error will reduce the correlation between CH₄ concentration and CH₄ production from the high value we report in Table 6.1 ($r_p = 0.95$) and under the controlled conditions in Wu et al., (2018) ($R^2 = 0.97$, $r_p = 0.99$). What can be seen from the R^2 reported under barn conditions in these studies is that the phenotypic correlation in the presence of repeated measures is, in practice, in the range of (0.30 - 0.61). Unfortunately, neither study reported individual level correlations between CH₄ production and CH₄ breath concentration, so it is not possible to determine the correlation between CH₄ production and CH₄ breath concentration when random and systematic errors are accounted for in the best fitting statistical model.

In Chapter 2 we developed a cow head lifting algorithm which filters out data when the gas concentrations are no longer significantly different from background gas concentrations and no longer changing dynamically. This is a prediction of when the cow head has left the feed bin, but cannot account for the variation of muzzle position within the feed bin. We installed an NDIR gas sensor (sampling rate 1 L/min), comparable to that of Huhtanen et al., (2015), and a FTIR gas sensor (sampling rate 4 L/min), same as that of Wu et al. (2018), in a single feed bin of an AMS, applied the time alignment and cow head lifting algorithm and compared instruments. Significant systematic differences were observed in the means between instruments for both gases, likely due to differences in CFG due to different sampling rates. This systematic difference was corrected for by calibrating one of the instruments onto the other. We furthermore, modelled the repeated measurements per day and per week to remove associated random error. In the case of CH₄ concentration, the intra-cow variation or imprecision decreased by more than 50% in both instruments from per visit to per day and more than 80% from per visit

to per week. The repeatability estimates correspondingly increased from $t = 0.38 - 0.33$ per AMS visit, to $t = 0.58 - 0.51$ per day and $t = 0.84 - 0.77$ per week for the NDIR and FTIR, respectively. These results show how increasing the number of records per cow can be an effective noise reduction method. However, the means remained largely unchanged and thus, systematic differences (i.e. decreased CFG) cannot be remedied through repeated measures. In principle, if we had run concentration recovery tests over each day of the experiment, it may have been possible to correct the drop in accuracy due to low and variable capture fractions, but repeatable cow factors like exhalation rates, would remain a source of bias. Regardless, the values of CH₄ and CO₂ breath concentrations per visit agreed sufficiently for interchangeable use (CIA = 0.78 and CIA = 0.65). Regrettably, we did not report the CCC or individual level correlations in Chapter 2, these were $r_I = 0.70$ and $CCC = 0.69$ per visit for CH₄ concentration. The individual level correlations from Chapter 3 for weekly CH₄ concentrations were slightly lower than CH₄ production (0.75 ± 0.20). Although the standard errors thereof are large, the correlation still indicates that CH₄ breath concentrations are a good candidate as a large scale indicator trait for CH₄ production in the RC.

6.2.3 Genetic variation in sniffer methane phenotypes

Sniffer methane phenotypes have been dismissed off hand by many authors for the great variation (implied imprecision), inaccuracy or deviations in linear association from flux methods (Huhtanen et al., 2015; Goopy et al., 2016; Hammond et al., 2016; Hristov et al., 2018). For example, in their review Hammond et al. (2016) stated ‘...the need for high throughput methodology, e.g. for screening large numbers of animals for genomic studies, does not in itself justify the use of methods that are inaccurate, imprecise, or biased.’. Whilst the accuracy of the breeding values starts with the accuracy and precision of the phenotype, the validity of this statement is not absolute when applied to the specific use of sniffer CH₄ phenotypes as large scale indicator traits for genetic evaluations. As demonstrated in Chapters 1-2, random bias or imprecision can be overcome through repeated measurements (see Text Box 1) and systematic bias or inaccuracy is of lesser importance as breeding values are centred (expressed relative to the base population with a mean of 0). As will be discussed below, slope or correlation bias (deviation from unity in linear association with the gold standard) can still be of use

in a selection index, provided there is a linear genetic correlation significant from zero.

In Chapter 4, we find significant heritability estimates of CH₄ ($h^2 = 0.16 - 0.26$) and CO₂ ($h^2 = 0.15 - 0.23$) from repeated weekly records of breath concentrations in first parity cows in Denmark and The Netherlands. In Chapter 5 we report a heritability of 0.19 ± 0.09 for CH₄ production using the CO₂ ratio method from single weekly records of multiparous Danish Holstein cows, in line with the findings of others using this method ($h^2 = 0.21 - 0.25$) (Lassen and Løvendahl, 2016; Zetouni et al., 2018). Using the individual level correlations estimated in Chapter 3, it is possible to estimate accuracy of breeding values for a bull with multiple daughter records (Figure 6.1). From Figure 6.1, it can be seen that the accuracy of bull EBVs fast approached 0.7 as the number of daughters with records exceeds 100, which is practically feasible with the high throughput of sniffers installed in robotic milking herds. This assumes CH₄ production in the RC is as heritable in lactating Holsteins on commercial mixed ration or grazing Holsteins using the SF₆ method 0.33 (0.15) (Breider et al., 2018) and our individual level correlations are reasonable predictors of the genetic correlations between sniffer and RC CH₄ emission. The resulting co-heritability (h_x, h_y, r_{xy}) is 0.19 using predicted CH₄ production and 0.21 using CH₄ breath concentrations, indicating considerable potential for sniffer CH₄ phenotypes as large-scale indicator traits. In this instance, the slope or correlation bias which results in individual level correlations between sniffer phenotypes and RC CH₄ production deviating from 1, does not preclude the use of sniffer phenotypes for breeding. However, this only demonstrates a proof of concept. Accurate estimates of genetic correlations between CH₄ production in the RC and sniffer phenotypes, as well as all other traits currently included in the selection index, are needed before implementation can proceed with any certainty. Recording thousands of cows in RC requires a substantial financial investment and the largest deterrent is that CH₄ production currently has no economic value, inhibiting the bio-economical modelling required to gauge the cost returns expected from selective breeding. This could be remedied by international collaborations to form a RC genomic reference population. Niu et al. (2018) recently collated RC data on 5,233 lactating dairy cattle, 68% of which are Holstein for phenotypic prediction equations, had all these cows been genotyped it would be feasible to estimate genetics correlations between the RC and other methods on genotyped Holstein populations.

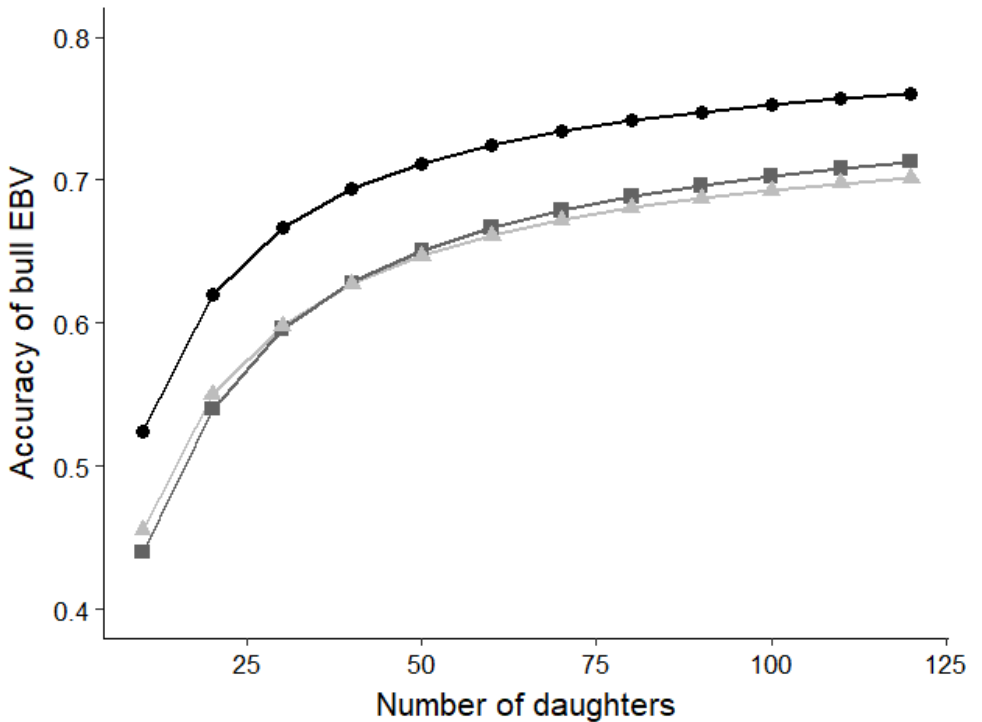


Figure 6.1. Accuracy of bull estimated breeding value (EBV) for CH₄ production in the RC ($h^2 = 0.33$), based on weekly averaged records for predicted CH₄ production from sniffers ($h^2 = 0.19$) (dark grey squares), CH₄ breath concentration ($h^2 = 0.22$) (light grey triangles) or both (black circles). Assuming $r_g \sim r_l$ and $rl = 0.77$ for CH₄ in RC and CH₄ production from sniffers and $r_l = 0.75$ for CH₄ production in the RC and CH₄ concentration from sniffers.

6.3 Genetic variation in feed efficiency of lactating cows

Unlike CH₄ emission, optimising feed utilisation in dairy cattle has a high economic value, but recording individual feed intake is challenging and costly and, thus, limits throughput to small research and specialised contract herds (Berry et al., 2014). As feed efficiency is not a directly observable animal metric, measuring feed efficiency relies on recording the feed intake and energy sinks and combining these into a single metric or index. In animal breeding, it has long been known that the most efficient index is a selection index based on genetic (co)variances and economic

weights (Hazel, 1943). The implications of this are that appropriate weighting of DMI and energy sinks in the breeding goal will result in a genetic feed efficiency sub index, which requires records on 10^4 - 10^5 animals (Veerkamp et al., 2013). However, for research purposes, a feed efficiency dependent variable is needed to determine the correlations between feed efficiency and other traits, as well as gauge responses to different interventions.

In its simplest form, feed efficiency can be defined as RFI, where DMI is regressed on multiple energy sinks in a multiple regression model and the resulting random residuals (RFI) are the feed efficiency phenotype analysed in an animal model (two-step RFI)(Koch et al., 1963). This method requires suitably large numbers of simultaneous records of DMI and energy sinks to accurately estimate the phenotypic regression coefficients. An extension of this model is to include the random additive genetic effects of animal directly into the multiple regression (single step RFI), which has been shown to be genetically correlated to two step RFI (0.80 – 0.991) (Saviotto et al., 2014; Tempelman et al., 2015). An advantage of both methods is that heterogeneity in regression coefficients due to different diets or stages of lactation, can be accounted for by including a regression by diet or lactation stage interaction term (Tempelman et al., 2015; Li et al., 2017). A disadvantage of both methods is that by defining phenotypic correlation coefficients, RFI is phenotypically independent of energy sinks, but genetic correlations remain and, thus, correlations to other traits are not representative (Huttmann et al., 2009; Spurlock et al., 2012). A further disadvantage of the single step model is that we cannot estimate genetic correlations between single step RFI and DMI in a multitrait model due to collinearity problems, making estimating the correlated response to selection for DMI impossible.

Kennedy et al. (1993) demonstrated a method using selection indices for estimating regression coefficients from (co)variances, post-hoc from multitrait models, which include DMI and energy sinks. By defining a restricted gains index of DMI conditional on the genetic (co)variances with energy sinks, it was possible to define RFI genetically independent of energy sinks (RFI_g). Under the specific set of circumstances where the genetic partial regression coefficients are equivalent to the phenotypic partial regression coefficients, RFI_g will be equivalent to phenotypic RFI models. A further advantage of this approach is that sparsely recorded energy sinks are not excluded from the model analyses. A disadvantage is that large numbers of records are required to accurately estimate genetic (co)variances. Furthermore, heterogeneity of partial regression coefficients due to different rations cannot be

estimated outside of the more cumbersome hierarchical Bayesian structural equation modelling (Lu et al., 2016).

In Chapter 4, we estimate gRFI for first parity lactating Holstein cattle in Denmark ($h^2 0.15 \pm 0.04$) and The Netherlands ($h^2 0.17 \pm 0.03$) and the combined country dataset ($h^2 0.15 \pm 0.04$), confirming that gRFI is significantly heritable (Lu et al., 2015). However, defining gRFI in Denmark with 813 cows, with 29084 weekly records, was challenging, as can be seen from the standard errors on the genetic correlations with component traits DMI (0.21 ± 0.64), ECM (0.002 ± 0.58) and live weight (0.004 ± 0.71). The Netherlands had 2112 first parity cows with 39,464 average weekly records for DMI, which reduced the uncertainty around the genetic correlations with DMI (0.74 ± 0.06), ECM (0.00 ± 0.15) and live weight (0.00 ± 0.11). Whilst, historical data, different experimental rations and different lactation stages can contribute to the heterogeneity of genetic partial regression coefficients and, thus, the standard errors of genetic correlations, these are two of the largest DMI datasets in the world at present. This implies that to define gRFI in practice, substantially larger numbers of cows are required to obtain more current and homogenous data. These results do lend confidence to the notion that breeding for increased metabolic efficiency is possible in lactating dairy cattle and these parameters can be utilized to estimate the cost benefit of large scale recording of DMI under commercial conditions. An important consideration is that we only estimate gRFI as a measure of metabolic efficiency during the first lactation, recognising that the productive life of cows spans multiple lactations and the life time efficiency spans lactating, growing and dry periods. In this respect, a considerable portion of the cow's life span is largely ignored; considerable efforts are required to define feed efficiency during all lifetime stages to truly increase the profitability and sustainability of the dairy industry. Regardless, metabolic efficiency during the first lactation is a logical first step in this progression.

To this end, new developments in the field of image analysis technology, providing an opportunity for the volumetric analysis of feed intake on dairy cows under commercial barn conditions, are promising. This technology is still under development, however, initial results show image analysis individual identification of cows show a sensitivity of 95% and specificity of 95% (Thomassen et al., 2017), which are comparable to current automated feeding systems reviewed in the general introduction. In a pilot study conducted over 14 days in a commercial Jersey barn in Denmark, Lassen et al. (2018) was able to obtain daily DMI ($t = 0.65$) and weekly averaged DMI ($t = 0.84$), which are directly comparable to those found in

Danish Jerseys used in automated feed bins in a research barn (Li et al., 2016). Although further validation of this technology is required, it poses a significant opportunity to increase the throughput of feed intake recording and, thus, feed efficiency throughout a cow's life.

6.3.1 The relationship between Feed efficiency and CH₄ emissions

As CH₄ emission currently has no economic value but, is estimated to constitute a loss of energy from the cow (2-12%) of gross energy intake (Johnson and Johnson, 1995), many have suggested that improving feed efficiency will concomitantly improve CH₄ emissions (Beauchemin et al., 2008; Waghorn and Hegarty, 2011; Lassen and Løvendahl, 2016). It follows that, since CH₄ production in the RC is directly proportional to DMI (Yan et al., 2010), breeding cows which eat less DMI, for contemporary levels of ECM and BW (the very definition of feed efficiency), will reduce CH₄ production. A counter argument is that increased diet digestibility (D) reflects the inherent ability of ruminants to transform ingested feed into metabolisable energy. Digestibility has been shown to be repeatable ($t = 0.28 - 0.37$) in dairy cattle (Huhtanen et al., 2015; Cabezas-Garcia et al., 2017). However empirical modelling evidence shows that D is a function of rate of digestion (kd) and rate of passage (kp): $D = kd / (kd + kp)$. Since kd is feed, specific and independent of the animal, only kp can influence digestibility. It follows that decreasing kp, will increase D, increasing feed efficiency but also increasing CH₄ emissions. Furthermore, studies in sheep divergent for CH₄ yield (CH₄/DMI), show that high emitters had significantly higher digestibility (Pinares-Patiño et al., 2011). Unfortunately, the authors did not compare CH₄ production in absolute terms and instead chose a ratio trait as their response variable.

As has been discussed previously herein, adequate high throughput phenotyping of traits required for estimating feed efficiency and CH₄ production with the RC is challenging. As a result, inconsistent definitions of feed efficiency are used, often with alternative methods of recording CH₄ emissions on a small scale at the phenotypic level. Furthermore, feed efficiency phenotypes and CH₄ emission phenotypes are often expressed as ratios, the limitations of ratios in terms of their non-linear relationships between the actual ratio and its component traits and the narrow range of parameters where these correlations are linear are discussed elsewhere (Sutherland, 1965; Veerkamp and Emmans, 1995). Not surprisingly, defining the relationship between feed efficiency and CH₄ emissions in dairy cattle has been inconsistent. A favourable phenotypic correlation was detected between

RFI and CH₄ production ($r_p = -0.55$) in Angus cattle on pasture using the GF (Velazco et al., 2017). No relationships were detected between RFI, CH₄ yield or CH₄ intensity in lactating Holstein, Simmental and Jerseys cows fed forage diets in the RC (Münger and Kreuzer, 2008). No differences in CH₄ production or CH₄ yield was detected on non-lactating Holstein-Friesians divergent for RFI, fed alfalfa cubes in the RC (Waghorn and Hegarty, 2011). Unfavourable relationships were found between RFI and CH₄ production in beef heifers on feedlot, recorded using RC (Montanholi et al., 2016). However, all of these studies had limited numbers of cows (16 - 39).

Studies using CH₄ production estimated from metabolisable energy intake yielded favourable correlations between feed efficiency and CH₄ production. de Haas et al., (2011) estimated CH₄ production from MEI in 665 first parity Holstein-Friesian cows and found RFI to be favourably correlated both phenotypically ($r_p = 0.72$) and genetically ($r_g = 0.38$) to RFI. In 429 lactating Nordic red cattle, CH₄ production was estimated using sniffers and the CO₂ ratio method, with predicted CO₂ tracer gas estimated from MEI, and found that the phenotypic correlation between residual energy intake (REI) (similar to RFI except energy intake is used in place of DMI) and CH₄ production was favourable ($r_p = 0.60$) (Negussie et al., 2014). In the same study, the ratio of ECM/EI was also favourably correlated ($r_p = -0.62$), both ratio and residual traits indicating that more efficient cows will produce less CH₄ in absolute terms. However, the phenotypic correlation between REI and CH₄ yield ($r_p = 0.38$) indicates that more efficient animals produce more CH₄ production per kilogram of energy intake. This finding could potentially reconcile the two opposing arguments about feed efficiency and CH₄ emissions, where more efficient animals produce less total CH₄, but more CH₄ per unit of feed intake through increased digestion. However far larger studies are required to confirm this on a genetic level. For instance, Breider et al (2018) estimated genetic parameters for DMI, MY, BW and CH₄ production (SF₆ method) in 314 Holstein cows on pasture. The restricted gains index for decreasing DMI, whilst maintaining MY and BW (definition of gRFI) showed a predicted correlated increase in CH₄ production. In Chapter 4, we applied the methods described by Kennedy et al. (1993) and found $r_g = -0.004$ between gRFI and CH₄ production.

In Chapter 4, we estimated genetic parameters for gRFI and CH₄ concentration from the sniffers directly, but found non-significant negative genetic correlations in Denmark ($r_g = -0.17 \pm 0.85$) and The Netherlands (-0.52 ± 0.32). This could, in part, be due to difficulties in acquiring adequate numbers of cows to define gRFI and estimate genetic correlations with CH₄ concentration. However, in the

combined population dataset we found a significant favourable genetic correlation ($r_g = 0.33 \pm 0.15$). The change of sign from positive to negative in the genetic correlation in combined population could be due to increased numbers of records. It could, however, also be an artefact of population heterogeneity in genetic partial regression coefficients or other unknown factors, like genotype by environment interactions. Further work is needed to validate these results, either through alternative modelling strategies for RFI, or increased within country records for both traits. In the absence of a genetic correlation between CH₄ concentration and CH₄ production in the RC, it is not possible to infer the relationship between gRFI and CH₄ production. What is clear, regardless of the sign of the genetic correlations, is that from their magnitude in all datasets, CH₄ concentration is a potential large-scale indicator trait for gRFI. Therefore, there is an economic incentive to record CH₄ concentrations until a genetic correlation with CH₄ production using the gold standard method is forthcoming.

6.4 Genetic variation in the rumen microbiome of lactating cows

In Chapter 5, we investigated both the genetic variation in rumen bacteria and archaea relative abundance and composition to assess the level of host genetic control on the rumen microbiome. At the individual OTU level, we identified 3,894 bacteria and 189 archaea present in at least 50% of the cows. From these, 4083 animals models were run and we found 6% of the rumen bacteria were significantly heritable (h^2 0.16 – 0.44) and 12% of the archaeal OTUs were significantly heritable (h^2 0.18 – 0.33). These results provide evidence from a good-sized cohort of cattle for host additive genetic effects on the rumen microbiome. However, the number of significantly heritable OTUs is alarmingly close to what would be expected under the null hypothesis. Although, it is not common practice for heritability estimates, we set a false discovery rate of 15% and none of the OTUs tested, passed this. This could, in part, be due to the sheer number of OTUs tested, many of whom are not heritable and are, relatively speaking, low number of cows ($n=750$).

There are limited heritability estimates for rumen microbes from which to draw comparisons. Sasson et al., (2017) assessed the heritability of relative abundance of rumen bacteria in 47 genotyped Holsteins, and found 22 out of 85255 OTUS to be significantly heritable ($h^2 > 0.70$). These estimates are considerably higher than our own, however, heritability estimates are notoriously biased in small numbers of cows. The same authors reported in their supplementary material a

significant heritability of 1 for DMI and RFI, 0 from milk yield (Sasson et al., 2017). Conversely, there is a wealth of literature on heritability of gut microbes in monogastrics. In 412 pairs of humans twins, Goodrich et al. (2014) found 5.3% of 909 OTUs to be significantly heritable (h^2 0.20 – 0.35), and in a follow up study with 1126 pairs of twins, 8.8% of 945 OTUs were found to be significantly heritable (h^2 0.15-0.42) (Goodrich et al., 2016). From these results, the authors conclude that human genetics shape the gut microbiota. In a study with 207 pigs, Camarinha-silva et al., (2017) found 8 genera out of 49 to be significantly heritable (h^2 0.32 – 0.52). Although the gastro-intestinal tract of monogastrics and ruminants are vastly different, these results together do suggest that a portion of the gut microbiota are heritable and, are thus, influenced by host genetic composition. In both cases, the host is heavily reliant on the activity of microbes for the derivation of energy from feed, a crucial requirement for survival. It is, thus, in the host's best interest to provide an environment suitable for microbes which fulfil key digestive processes. As discussed by Shapira, (2016), evolutionary processes rarely leave biological pathways essential to survival up to chance.

The variation in the rumen microbiome can be partitioned into other sources of host variation, other than additive genetics. If repeated measures are available, it is possible to partition phenotypic variation into permanent environmental effects of host (also called general environment) and residual variation (also called special environment)(Falconer and Mackay, 1996). If repeated measures are made over very short periods of time, such that underlying biology has not changed, then the residual variation is 'method imprecision', as discussed in previous sections of this thesis. The permanent environment includes host additive genetics, non-additive genetics, maternal effects, common environmental effects and other host traits, like learnt behaviour. The permanent environment expressed as a proportion of the total variation is the repeatability, as mentioned in previous sections of this thesis. In a subset of 57 cows from Chapter 5, we had repeated rumen samples 122 days apart, on the same diet, and ran repeatability animal models at the genus level (Zhu, 2016). We combined this with the genus level heritability estimates reported in Chapter 5 for comparison in Figures 6.2-6.3 below. Importantly, the repeatability estimates are a measure of persistence of ranking between cows through time, which had considerable range ($t = 0 - 0.80$). A proportion of genera 11.2% were not repeatable ($t = 0$), indicating the host explains no variation in their abundance and 12% has $t < 0.50$, indicating more than 50% of the variation in their abundance was explained by host.

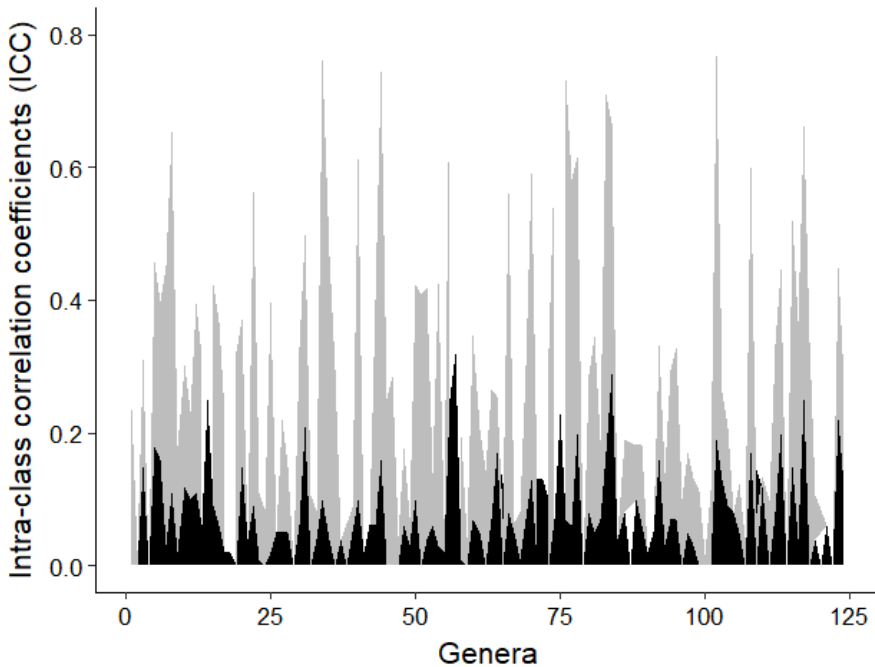


Figure 6.2 Intraclass correlations coefficients for rumen genera, repeatability (Grey) for 5 repeat sampled Holsteins and Heritability (Black) for 750 single sampled Holsteins.

From Figure 6.3 there is a general positive trend between magnitude of repeatability and heritability, although for some genera they were repeatable but not genetic and vice versa this is partly due to the smaller sample size with repeated measurements as $t = 0$ and $h^2 < 0$ is not possible in practice. Similar results were reported for human gut microbiota using the intra class correlation coefficient (ICC) between pairs of dizygotic and monozygotic twins and heritability estimates (Goodrich et al., 2014, 2016). This ICC between identical twins is equivalent to the repeatability from repeated measures, however, the ICC between dizygotic twins reflects shared common and maternal effects (the difference between ICC for monozygotic twins and dizygotic twins is half the additive genetics). Together, these results indicate that the host does influence the abundance of certain rumen microbiota and that additive genetics is one of the contributing host effects. Further work and samples from larger numbers of cows are required to assess the effects of maternal and common environmental effects.

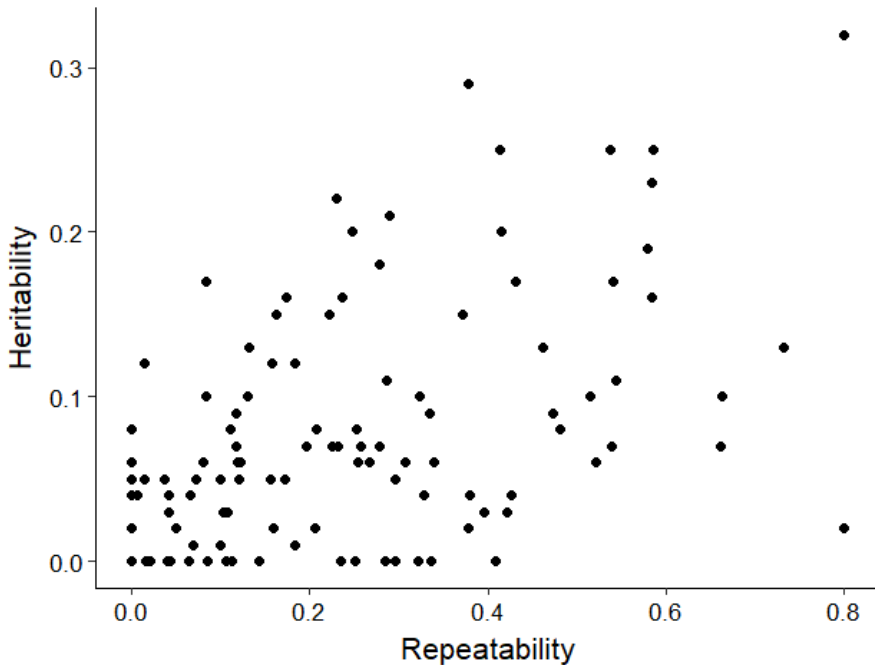


Figure 6.3 Repeatability of rumen genera in 57 Holsteins cows against the heritability of rumen genera in 750 Holstein cows.

6.4.1 Genetic variation in rumen bacterial and archaeal abundance

Some authors have subsequently challenged the notion that the relatively small proportion of heritable microbes can truly shape the entire community composition. For instance, Rothschild et al. (2017) weighted the heritability estimates of Goodrich et al. (2016a) by their relative abundance to estimate heritability of the entire microbiome and found 1.9% - 8.1%, depending on the use of a false discovery rate threshold for heritability estimates. For our dataset, the respective weighted average heritability of the rumen bacteria and archaea from Chapter 5 was 10.9%. However, simply weighting the heritability by relative abundance assumes all microbiota are of equal functional importance and do not interact with each other.

Examining microbial community composition through multidimensional scaling, such as principal coordinate analysis or beta diversity, has identified stratified clusters in humans referred to as 'enterotypes' (Arumugam et al., 2013) and ruminants referred to as 'ruminotypes' (Kittelmann et al., 2014). Diversity

estimates are used in microbial ecology to assess the diversity within samples (alpha diversity) and between samples (Beta diversity). These clusters found using beta diversity, are thought to represent altered host-microbial states (Arumugam et al., 2013) and in the case of sheep, one of the clusters was associated with smaller rumen size, high rumen passage rate and lower CH₄ emissions (Kittelmann et al., 2014). In Chapter 5, we investigated ruminotypes in 750 Holstein cattle and found clustering, which was partly explained by non-genetic host parameters, like lactation stage and parity, technical factors, like sequence batch and environmental variables, like herd. However, the largest sources of variation was additive genetics (h^2 0.20 – 0.39). These findings suggest that host additive genetics influences the rumen microbial communities to a larger extent than was suggested using the average heritability. It is important to note that clustering methods are data set specific and are, thus, affected by population and environmental heterogeneity. In our study, the herds were selected partly for similar management and dietary regimes, as well as breed uniformity (Danish Holsteins), to minimize these potential sources of variation. In the UK, in the human twins dataset, Goodrich et al., (2016) investigated different alpha and beta diversity metrics and found strong correlations between heritable microbiota, as well as host SNPs with diversity measures, indicating heritable microbes and host genetics influences community structures. Conversely, in a study with 1142 genotyped individuals in Israel from diverse genetic backgrounds (grandparents originating from 72 different countries), Rothchild et al, (2017) failed to detect a heritable component to gut microbial diversity estimates or significantly associated SNPs. Thus, it is important to consider the genetic and environmental heterogeneity when assessing their relative contributions to microbial communities, as both sources are associated with microbial variations.

6.4.2 Genetic variation in rumen bacterial and archaeal community composition

As genetic variation in CH₄ production and rumen microbes is detected in Chapter 5 and methanogenic archaea and bacteria are directly responsible for CH₄ production, a logical progression is to investigate associations between rumen microbiota and CH₄ production. We ran a microbiome wide association study (MWAS) for OTUs and genera against CH₄ production, whilst simultaneously controlling for familial relationships and non-genetic factors, and found eight genera to exceed an FDR threshold of 15%. Of these eight, four were found to be affected by methane inhibitors or be associated with the production of methanogenic

substrates; none were archaea, further strengthening the hypothesis that bacteria drive CH₄ production through methanogenic substrate production. Of the eight, three were found to be moderately heritable (h^2 0.17 – 0.20), a positive indicator that breeding or rumen microbial abundance could result in changes to CH₄ production, but genetic correlations from larger datasets are required to confirm this. Single OTU or genera associations fail to capture the complexity of microbial communities, so we also investigated associations between community structure and CH₄ production and found the first principal component of both bacterial and archaeal beta diversity significantly associated with CH₄ production, indicating there is an optimal rumen microbial community composition for CH₄ production.

Due to the moderate heritability estimates of rumen microbiota and CH₄ production, as well as the detection of heritable and non-heritable rumen microbiota associated with CH₄ production, it was necessary to quantify the relative contributions of rumen microbiota and host genetics to variation in CH₄ production simultaneously. To do this, we create a microbial kinship matrix, analogous to genomic or pedigree relationship matrices used in the estimation of heritability. We defined microbiability (m^2) as the proportion of phenotypic variance due to rumen microbial kinship, in analogy to heritability. As with heritability, microbiability is narrow sense, in that it does not account for microbe-microbe interactions or microbe-environment interactions. In Chapter 5, we found significant microbiability of CH₄ production ($m^2 = 0.15$) comparable to that of heritability ($h^2 = 0.19$). Crucially, when comparing these two sources of variation, first separately and then jointly, we found minor changes to the estimates, indicating that in the case of CH₄ production, the two sources are largely independent. This implies that selection for microbes associated with CH₄ production may have limited impact, as the majority of the variation in m^2 is due to rumen microbiota less influence by the host genetics and vice versa. This also opens up the possibility for a largely under exploited source in host variation.

6.4.3 Prospective for using rumen microbiota to obtain desired changes in the host phenotype

It is proposed in the general introduction that the successful symbiosis due to coevolution of ruminant and rumen microbiome suggests that both be viewed together as a single unit (i.e. as a holobiont). It was discussed how ignorance of the microbial contribution to host phenotypes can result in suboptimal or unpredictable response, as was seen by rapid microbial adaption to methanogenic inhibiting

additives. For traits of interest, we propose determining the microbiability through the use of microbial relationship matrices derived from OTUs, full metagenomic or transcriptomic sequences to quantify the microbial association with host phenotypic variation. If a trait of interest is not associated with the microbiome, further exploration of the rumen microbiome is not needed. However if a trait of interest is associated with the rumen microbiome it is necessary to determine the relative independence from diet or the host additive genetics. Particularly, the genetic correlations between associated microbiota and the host trait of interest to ensure unwanted correlated responses to selection in the rumen microbiome are not realized. Conversely, if genetic correlations are sufficiently greater than 0 there is a potential to incorporate rumen microbial phenotypes into the selection index as indicator traits to improve the accuracy of selection and thus the response.

By estimating the microbiability we determined that differences in the rumen microbiome are associated with differences in the host CH₄ production. Furthermore, some of the specific rumen microbiota are likely under the influence of host additive genetics, but to a large extent the two are independent. In the case of CH₄ production where the two sources are largely independent, selective breeding is one method to reduce CH₄ emissions but this will not fully exploit the variation observed due to the rumen microbiome and will take generations to illicit a response. Better understanding of whether microbiability is causative or merely co-associated with differences in the host CH₄ production and other traits are needed. One potential method is to inoculate cows with poor microbial estimated values (random solutions from microbiability models) with the rumen contents of cows which have favourable microbial estimated values and assess the response in host CH₄ production. Inoculation of cattle with the cud of healthy donor cows has long been an effective and rapid treatment of cows with digestive or metabolic disorders (DePeters and George, 2014). This is an example of using rumen microbiota to affect a desired change to the host cow phenotype. Recent work with repeated inoculation of beef heifers with bison rumen contents altered the rumen microbiome and metabolic activity eliciting increased protein and nitrogen retention on a nutritionally poor straw diet (Ribeiro et al., 2017). Another example is the discovery of *Synergistes jonesii* in goats which conveys the ability of the host to degrade the otherwise toxic pyridinediol compounds from the fodder plant *Leucaena leucocephala*, which was successfully conveyed to Australian beef cattle and sheep (Graham et al., 2013). Both of these examples demonstrate the rapid adaption of

host to dietary changes acquired through the incorporation of beneficial microbes into the rumen community.

6.5 Conclusions and Recommendations:

The aims of this thesis were to improve and validate the high throughput non-invasive sniffer method of recording CH₄ emissions under commercial conditions. Furthermore, to investigate the genetic relationships between CH₄ emission, feed efficiency and the rumen microbiome. Through instrument comparisons and validation, sources of error in sniffer CH₄ emission phenotypes were identified and where possible corrected. The resulting CH₄ production and CH₄ breath concentration phenotypes were found to be significantly heritable but were low-moderately in agreement at the phenotypic level with RC CH₄ production. However, the individual level correlations (as proxies for genetic correlations) between sniffer CH₄ emission phenotypes were high (0.75 – 0.77) demonstrating the value of sniffer CH₄ emission phenotypes as large scale indicator traits in a selection index for reduced RC CH₄ production.

Feed efficiency was found to be significantly heritable in first parity Holsteins in Denmark and The Netherlands. Furthermore favourable genetic correlations were estimated between CH₄ breath concentration and feed efficiency, to this sniffer CH₄ emission phenotypes can be of use as large scale indicator traits. A proportion of the relative abundance of the rumen bacteria and archaea were found to be heritable, demonstrating the host additive genetics can influence the rumen microbiome. Some heritable bacterial genera were found to be significantly associated with CH₄ production, but the majority of variation in CH₄ production explained by rumen microbial composition was largely independent of host additive genetics. The aims of this thesis were largely met; however, caveats and future research considerations remain.

The first priority for improving the profitability and sustainability of the dairy industry is the development of a life-time metabolic and maintenance efficiency index. The single most limiting factor to breeding for more feed efficient cattle is an accurate, precise, non-invasive, cost effective and high throughput method of recording feed intake in dairy cattle under commercial conditions. Small research studies have demonstrated proof of concept in genetic variation for metabolic and maintenance efficiency during lactation. Until methane production has a direct economic or societal value, its value to the farmer will lie in being an

indicator trait for feed intake and feed efficiency. To this end, the sniffer method shows considerable promise, but any further development to the method that increases accuracy and precision without compromising throughput and non-invasiveness, will be of value to improving accuracy of EBVs. Should CH₄ production get an economic or societal value, obtaining genetic correlations with meaningful standard errors between RC and sniffers will be crucial to their use as a large scale indicator trait for CH₄ production. Simulations show combining genomic information with feed efficiency records (and likely CH₄ production as well) on 4000 cows a year will yield adequate genomic prediction accuracies to obtain genetic responses for these traits (Wallén et al., 2017). With considerable investment, these numbers are achievable in the RC, which will yield meaningful genetic correlations between feed intake, efficiency, production traits and CH₄ production within the RC and other traits under commercial conditions like sniffer breath concentration measures. The first breeding company to publish EBVs for these traits will have a competitive advantage which must be considered in cost benefit analyses.

Further work is needed in linking rumen microbial associations to causation. Whilst some rumen microbes are heritable their cost and throughput limits their use as indicator traits for already costly and limiting breeding goal traits like CH₄ production and feed efficiency. Estimating microbiability and estimated microbial values for host phenotypes is the first step to identifying desirable rumen microbial communities associated with host phenotypes. However, validation of transmittable desired changes to host phenotypes through transfaunation and inoculation is still required. This could be investigated by *in vitro* batch fermentation studies before attempting *in vivo* investigations. Should improved feed efficiency and CH₄ production prove transmittable through inoculation with rumen contents from donor animals this could open up new possibilities for directing rapid changes to host phenotypes using quantitative genetics methods to identify optimal microbial communities.

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Gareth Difford
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About the author

Gareth Frank Difford was born 31 October 1988 in East London South Africa. He completed a B.S.c in Molecular Biology and Biotechnology from the Stellenbosch University, South Africa in 2010. During his bachelor he developed an interest in aquaculture and genetics. Shortly thereafter, he obtained a scholarship to study a master's degree in Genetics and Animal Breeding at Stellenbosch University, South Africa. The topic of research during his masters was selective breeding of the South African Abalone (*Haliotis Midae*), during which he developed a great appreciation for phenotyping, quantitative genetics and statistical methods. During the final year of his master's degree, he obtained an Erasmus Mundus Scholarship for a research exchange with the Katholieke Universiteit Leuven, Belgium during which he worked on genetic parameters for growth and carcass quality traits in the Flemish pig breeding program. This master's thesis resulted in the first genetic parameters for growth traits in South African abalone and was awarded cum laude in 2013. During University vacations and after his M.S.c, he worked as an assistant manager at a land-based recirculation facility in East London on Dusky Kob (*Argyromomus Japonicus*), which resulted in a collaborative research project and the first breeding program for the species. In 2014 Gareth Difford was awarded an Erasmus Mundus joint PhD scholarship between Aarhus University, Denmark and Wageningen Univeristy, The Netherlands. The PhD topic is 'Genetic control of Methane emissions, Feed efficiency and Metagenomics in Dairy Cattle' (This Thesis) defended 28 September 2018. After his PhD, he continues as a Researcher at Nofima, Norway working on phenomics and quantitative genetics in aquaculture species.

Peer review publications

1. Vlok, A. C., **Difford, G. F.**, Rhode, C., & Brink, D. (2016). An Assessment of Hatchery Cohort Growth Rates of South African Abalone, *Haliotis midae*, Across Four Commercial Environments. *Journal of the World Aquaculture Society*, 47(5), 658-666.
2. **Difford, G. F.**, Lassen, J., & Løvendahl, P. (2016). Interchangeability between methane measurements in dairy cows assessed by comparing precision and agreement of two non-invasive infrared methods. *Computers and Electronics in Agriculture*, 124, 220-226.
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4. Zhu, Z., Noel, S. J., **Difford, G. F.**, Al-Soud, W. A., Brejnrod, A., Sørensen, S. J., ... & Højberg, O. (2017). Community structure of the metabolically active rumen bacterial and archaeal communities of dairy cows over the transition period. *PLoS one*, 12(11), e0187858.
5. Shetty, N., **Difford, G.**, Lassen, J., Løvendahl, P., & Buitenhuis, A. J. (2017). Predicting methane emissions of lactating Danish Holstein cows using Fourier transform mid-infrared spectroscopy of milk. *Journal of Dairy Science*, 100(11), 9052-9060.
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7. **Difford, G. F.**, Løvendahl, P., Lassen, J., Guldbbrandtsen, B., & Sahana, G. (2018). Microbiability—new insights into (genetic) modelling methane emissions of cattle. In *11th World Congress on Genetics Applied to Livestock Production*.
8. Zhu, Z., Kristensen, L., **Difford, G. F.**, Poulsen, M., Noel, S. J., ... & Højberg, O. (2018). Changes in rumen bacterial and archaeal communities over the transition period in primiparous Holstein dairy cows. *Journal of Dairy Science*, DOI: <https://doi.org/10.3168/jds.2017-14366>.
9. Løvendahl, P., **Difford, G.F.**, Li, B., Chagunda, M.G.G., Huhtanen, P., Lidauer, M.H., Lassen, J., & P. Lund. (2018). Selecting for improved feed efficiency and reduced methane emissions in dairy cattle (a review). Accepted. *Animal*.

10. **Difford, G.F.**, Plichta, D.P., Løvendahl, P., Lassen, J., Noel, S.J., Højberg, O., Wright, A.D.G., Zhu, Z., Kristensen, L., Nielsen, H.B., Guldbbrandtsen, B., & Sahana, G. (2018). Host genetics and the rumen microbiome jointly associate with methane emissions in dairy cows. Accepted. PLOS Genetics.
11. Sorg, D., **Difford, G.F.**, Mühlbach, S., Kuhla, B., Swalve, H.H., Lassen, J., Strabel, T., & Pszczola, M. (2018). Comparison of a laser methane detector with the GreenFeed and two breath analysers for on-farm measurements of methane emissions from dairy cows. *Computers and Electronics in Agriculture*, 153, 285-294.

Conference proceedings, abstracts and presentations

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2. **Difford, G.F.**, Lassen, J., & Løvendahl, P. 2015. Disagreements on current time an overlooked source of noise? Cost Action Methagene Annual Meeting. Wageningen, The Netherlands.
3. Lassen, J., **Difford, G.F.**, & Løvendahl, P. 2015. Genetic approaches for methane mitigation strategies. The 66th Annual Meeting of the European Federation of Animal Science (EAAP). Vol. 21 The Netherlands : Wageningen Academic Publishers, 2015. p. 92-92.
4. **Difford, G.F.**, Lassen, J., & Løvendahl, P. 2016. Comparing equipment for recording enteric methane concentrations, carbon dioxide concentrations and their ratio in the breath of Danish Holsteins during automated milking. The 6th Greenhouse Gas and Animal Agriculture Conference (GGAA), Melbourne, Australia. (Poster).
5. **Difford, G.F.**, Lassen, J., & Løvendahl, P. 2016. Genes and microbes, the next step in dairy cattle breeding. Book of Abstracts of the 67th Annual Meeting of the European Association for Animal Production (EAAP). Wageningen Academic Publishers.
6. de Haas, Y., Pryce, J.E., Wall, E., McParland, S., **Difford, G.F.**, & Lassen, J. 2016. Genetic and genomic analyses of enteric methane emissions of dairy cattle. The 67th Annual Meeting of the European Association for Animal Production (EAAP). Wageningen Academic Publishers.

7. Szalanski, M., **Difford, G.F.**, Løvendahl, P., Lassen, J. 2016. Comparing methods to analyse raw, large-scale methane data. The 67th Annual Meeting of the European Association for Animal Production (EAAP). Wageningen Academic Publishers.
8. **Difford, G.F.**, Løvendahl, P., Lassen, J., & de Haas, Y. 2016. Case studies of sniffer comparisons in Denmark. Cost Action Methagene Annual Meeting. Padova, Italy.
9. **Difford, G.F.**, de Haas, Y., Visker, M.H.P.W., Lassen, J., Bovenhuis, H., Veerkamp, R., & Løvendahl, P. 2017. Do breath gas measurements hold the key to unlocking the genetics of feed efficiency in dairy cows? The 68th Annual Meeting of the European Federation of Animal Science (EAAP), Tallinn, Estonia. Vol. 23 2017. ed. Wageningen Academic Publishers, 2017. p. 184.
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13. **Difford, G.F.**, Løvendahl, P., Lassen, J., Guldbrandtsen, B., & Sahana, G. 2018. Invited speaker: Microbiability – new insights into (genetic) modelling methane emissions of cattle. Proceedings of the World Congress on Genetics Applied to Livestock Production, Auckland, New Zealand: Volume Challenges - Environmental. Vol. 11 2018. 405.
14. **Difford, G.F.**, Gebreyesus, G., Løvendahl, P., Buitenhuis, A. J., Lassen, J. Guldbrandtsen, B., & G. Sahana. 2018. The 69th Annual Meeting of the European Federation of Animal Science (EAAP), Dubrovnik, Croatia. Vol. 23 2017. ed. Wageningen Academic Publishers, 2018.

Individual Training Plan



Mandatory courses (6 ECTS)

| | Year |
|--|------|
| Welcome to EGS-ABG | 2014 |
| EGS-ABG Summer Research School | 2015 |
| EGS-ABG Summer Research School | 2018 |
| Ethics course: Consequences of selection | 2014 |

Advanced Scientific courses (24 ECTS)

| | |
|--|------|
| Introduction to theory and implementation of Genomic Selection | 2014 |
| Methagene Training School – Methane Physiology for Geneticists | 2014 |
| Linear Models in Animal Breeding | 2015 |
| Gene Mapping | 2015 |
| Quantitative Genomics | 2015 |
| Methagene Training School – Rumen Microbial Ecosystem | 2016 |
| Methagene Training School – Breeding for complex traits | 2017 |
| Nova Course – Feed Efficiency in Dairy Cattle | 2016 |

Professional Skills support courses (6 ECTS)

| | |
|-------------------------------------|------|
| QGG research skills | 2014 |
| R programming for animal scientists | 2015 |

Dissemination of knowledge (ECTS)

Teaching

| | |
|--|------|
| Teaching assistant – Methagene Training School: Large scale methane data handling, analysis and interpretation | 2015 |
| Teaching assistant – Linear Models in Animal Breeding | 2016 |
| Teaching assistant – Population Genetics | 2016 |
| MSc supervision | 2016 |

International conferences

| | |
|---|-------------|
| Annual meeting of the European Federation of Animal Science (EAAP) | 2015 - 2018 |
| Cost Action Methagene Annual Meeting | 2015 - 2017 |
| The 6th Greenhouse Gas and Animal Agriculture Conference (GGAA), | 2016 |
| World Congress on Genetics Applied to Livestock Production (WCGALP) | 2018 |

Seminars and workshop

| | |
|---|------------|
| Annual Gensap meeting | 2015 |
| Nordic Feed Efficiency in Nordic cattle (FUNC) annual meeting | 2016 |
| Nordic Cattle Meeting on Genomic Selection | 2015, 2018 |

Colophon

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